# Bayesian ODE-penalized B-spline model with Gaussian mixture as error distribution

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# Motivating example

#### Theophylline dataset

Data for the kinetics of the anti-asthmatic drug theophylline with 12 subjects, concentration were measured at 11 time point over 25 hours for each subject.

#### **Differential equation**

$$\begin{cases} \frac{dx_1}{dt}(t) = -k_a x_1(t), \\ \frac{dx_2}{dt}(t) = \frac{k_a}{V} x_1(t) - k_e x_2(t), \\ x_1(0) = D \\ x_2(0) = 0. \end{cases}$$

-  $x_1(t)$  quantity of drug in stomach, -  $x_2(t)$  concentration of drug in blood.



Figure : Plasma concentration of theophylline over time

### Introduce the

# Bayesian ODE-penalized B-spline approach

in the case of systems of affine differential equations

when the conditional error distribution is a

penalized mixture of Gaussian distributions

# Theoretical part

# Differential equation and observations

#### **Differential equation**

$$\begin{cases} \frac{d\mathbf{x}}{dt}(t) = \mathbf{f}(\mathbf{x}(t),\theta) \\ \mathbf{x}(0) = \mathbf{x}_0 \end{cases}$$

with:

-  $(\mathbf{x}(t))^{T} = (x_{1}(t), \dots, x_{d}(t))$  the set of *d* state functions and  $\mathbf{x}_{0}$  the set of initial conditions,

-  $\theta \in \mathbb{R}^q$  the vector of unknown parameters,

-  $\mathbf{f}$  a known affine function of  $\mathbf{x}$ .

### Differential equation and observations

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- **f** a known affine function of **x**.

#### Measurement

A subset  $\mathcal J$  of the *d* state functions are observed with additive measurement errors:

$$y_{jk} = x_j \left( t_{jk} \right) + \tau_j^{-1/2} \epsilon_{jk}$$

with  $\mathbb{E}(\epsilon_j) = 0$  and  $\mathbb{V}(\epsilon_j) = 1$ .

# B-spline approximation

Each state function  $x_j(t)$  is approximated using a B-spline basis function expansion:

$$\widetilde{x}_{j}(t) = \sum_{k=1}^{K_{j}} B_{jk}(t) c_{jk} = (\mathbf{B}_{j}(t))^{T} \mathbf{c}_{j}$$

where:

- $\mathbf{B}_{\mathbf{i}}(t)$  is the vector of B-spline basis function evaluated at time t,
- $\mathbf{c}_{\mathbf{j}}$  is the vector of spline coefficients.

# Influence of the number of knots and the spline coefficients

# ODE-penalty (Ramsay et al., 2007)

The penalty for the *j*th equation quantifies the proximity of the approximation  $\tilde{x}_j(t)$  to the *j*th component of the solution  $\mathbf{x}(t)$ :

$$PEN_{j} = \int \left(\frac{d\widetilde{x}_{j}}{dt}(t) - f_{j}(\widetilde{\mathbf{x}}(t), \theta)\right)^{2} dt$$
$$PEN = \sum_{j=1}^{d} \gamma_{j} PEN_{j}$$
$$= \mathbf{c}^{T} \mathbf{R}(\theta, \gamma) \mathbf{c} + 2\mathbf{c}^{T} \mathbf{r}(\theta, \gamma) + I(\theta, \gamma)$$

where: -  $\mathbf{c}^{T} = (\mathbf{c_1}^{T}, \dots, \mathbf{c_d}^{T})$  is the vector of spline coefficients, -  $\gamma^{T} = (\gamma_1, \dots, \gamma_d)$  is the vector of ODE-adhesion parameters.

# Influence of the ODE-adhesion parameter

Prior distribution for the spline coefficients (Jaeger and Lambert, 2011)

$$ho\left(\mathbf{c}| heta,\gamma
ight) \propto \exp\left(-rac{1}{2}\sum_{j=1}^{d}\gamma_{j}PEN_{j}
ight)$$

Prior distribution for the spline coefficients (Jaeger and Lambert, 2011)

$$p(\mathbf{c}| heta,\gamma) = \propto \exp\left(-rac{1}{2}\sum_{j=1}^{d}\gamma_{j}PEN_{j} - rac{1}{2}\left(\mathbf{c}-\mu_{\mathbf{c}}
ight)^{T}\mathbf{P}_{\mathbf{c}}\left(\mathbf{c}-\mu_{\mathbf{c}}
ight)
ight)$$

with:

-  $\mu_{c}$  and  $\mathbf{P}_{c}$  to express uncertainty w.r.t initial conditions of the states,

Prior distribution for the spline coefficients (Jaeger and Lambert, 2011)

$$p(\mathbf{c}|\theta,\gamma) \propto \exp\left(-\frac{1}{2}\left\{\mathbf{c}^{T}\left(\mathbf{R}\left(\theta,\gamma\right)+\mathbf{P_{c}}\right)\mathbf{c}-2\mathbf{c}^{T}\left(-\mathbf{r}\left(\theta,\gamma\right)+\mathbf{P_{c}}\mu_{c}\right)\right\}\right)$$

with:

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Prior distribution for the spline coefficients (Jaeger and Lambert, 2011)

$$p(\mathbf{c}|\theta,\gamma) \propto \exp\left(-\frac{1}{2}\left\{\mathbf{c}^{\mathsf{T}}\mathbf{V}_{1}\mathbf{c}-2\mathbf{c}^{\mathsf{T}}\mathbf{v}_{1}\right\}\right)$$

with:

-  $\mu_c$  and  $P_c$  to express uncertainty w.r.t initial conditions of the states, -  $v_1 = -r(\theta, \gamma) + P_c \mu_c$  and  $V_1 = R(\theta, \gamma) + P_c$ ,

Prior distribution for the spline coefficients (Jaeger and Lambert, 2011)

$$\rho(\mathbf{c}|\theta,\gamma) \propto \left(\det(\mathbf{V}_1)\right)^{1/2} \exp\left(-\frac{1}{2}\left\{\mathbf{c}^{\mathsf{T}}\mathbf{V}_1\mathbf{c} - 2\mathbf{c}^{\mathsf{T}}\mathbf{v}_1 + \mathbf{v}_1^{\mathsf{T}}\mathbf{V}_1^{-1}\mathbf{v}_1\right\}\right)$$

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#### Prior distribution for the ODE-adhesion parameters

$$\gamma_{j} \sim \mathcal{E}$$
xp  $\left(10^{-8}
ight)$ 

It indicates the prior confidence on the proposed ODE to model the dynamic of the state  $\mathbf{x}(t)$ .

#### Specification of the conditional distribution

Conditional error distribution (Komárek and Lesaffre, 2008)

$$\epsilon_{jk} | \mathbf{d_j} \sim \sum_{l=-L_j}^{L_j} \pi_{jl} \left( \mathbf{d_j} 
ight) \mathcal{N} \left( \mu_{jl}; \sigma_j^2 
ight)$$

where  $(\mu_{jl})_{l=-L_j}^{L_j}$  is a fixed large number of pre-selected equidistant means and  $\sigma_j^2$  is a fixed variance, with constraints  $\pi_{jl} > 0$  and  $\sum_l \pi_{jl} = 1$ .

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#### Generalized logit transformation of the mixture weights

$$\pi_{jl} = \frac{\exp\left(a_{j,l}\right)}{\sum_{k=-L_j}^{L_j} \exp\left(a_{j,k}\right)}$$

with identifiability constraints  $a_{j,L_j} = 0$  for all  $j \in \mathcal{J}$ . Only  $\mathbf{d}_j = \text{vec}(a_{j,l}; l = -L_j, \dots, L_j - 1), j \in \mathcal{J}$  have to be estimated to estimate the error distributions.

### Fitted error densities



Figure : Fitted error densities (Student, extreme value and mixture of 2 Gaussians)

Finite difference penalty on the transformed weights

$$egin{aligned} Q\left(\mathbf{a_j}|\lambda_j
ight) &= -rac{\lambda_j}{2}\sum_l \left(\mathbf{a}_{j,l} - 3\mathbf{a}_{j,l-1} + 3\mathbf{a}_{j,l-2} - \mathbf{a}_{j,l-3}
ight)^2 \ &= -rac{\lambda_j}{2}\left(\mathbf{Da_j}
ight)^T\mathbf{Da_j} \end{aligned}$$

Finite difference penalty on the transformed weights

$$\begin{split} Q\left(\mathbf{a_j}|\lambda_j\right) &= -\frac{\lambda_j}{2}\sum_{l}\left(a_{j,l} - 3a_{j,l-1} + 3a_{j,l-2} - a_{j,l-3}\right)^2 \\ &= -\frac{\lambda_j}{2}\left(\mathbf{Da_j}\right)^T\mathbf{Da_j} \\ &\Rightarrow Q\left(\mathbf{d_j}|\lambda_j\right) = -\frac{\lambda_j}{2}\mathbf{d_j}^T\mathbf{Pd_j} \end{split}$$

where:

-  $\lambda_j$ ,  $j \in \mathcal{J}$  is the roughness penalty parameter,

- **P** is the submatrix of  $\mathbf{D}^T \mathbf{D}$  without the last line and last column.

Prior distribution for the transformed weights d<sub>j</sub>

$$p\left(\mathbf{d_j}|\lambda_j
ight) \propto \exp\left(-rac{\lambda_j}{2}{\mathbf{d_j}}^{\mathsf{T}}\mathbf{P}\mathbf{d_j}
ight)$$

Prior distribution for the transformed weights d<sub>j</sub>

$$p\left(\mathbf{d}_{\mathbf{j}}|\lambda_{j}\right) \propto \boldsymbol{\lambda}_{j}^{L_{j}-1} \exp\left(-\frac{\lambda_{j}}{2}\mathbf{d}_{\mathbf{j}}^{T}\mathbf{P}\mathbf{d}_{\mathbf{j}}\right)$$

Prior distribution for the transformed weights d<sub>j</sub>

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Prior distribution for the roughness penalty parameters

$$\lambda_{j} \sim \mathcal{E} x p \left( 10^{-8} 
ight)$$

It indicates the prior confidence on the Gaussian distribution for the error distribution.

# Bayesian model (Jaeger and Lambert, 2012a)

$$\begin{aligned} \mathbf{y}_{jk} &= \widetilde{\mathbf{x}}_{j}\left(t_{jk}\right) + \tau_{j}^{-1/2}\epsilon_{jk} \\ \epsilon_{jk}|\mathbf{d}_{j} &\sim \sum_{l=-L_{j}}^{L_{j}} \pi_{jl}\left(\mathbf{d}_{j}\right) \mathcal{N}\left(\mu_{jl} + \alpha_{j}; \left(\frac{\sigma_{j}}{\beta_{j}}\right)^{2}\right) \\ \mathbf{p}\left(\mathbf{c}|\theta,\gamma\right) &\propto \left(\det\left(\mathbf{V}_{1}\right)\right)^{1/2}\exp\left(-\frac{1}{2}\left\{\mathbf{c}^{T}\mathbf{V}_{1}\mathbf{c}-2\mathbf{c}^{T}\mathbf{v}_{1} + \mathbf{v}_{1}^{T}\mathbf{V}_{1}^{-1}\mathbf{v}_{1}\right\}\right) \\ \gamma_{j} &\sim \mathcal{E}xp\left(10^{-8}\right) \\ \mathbf{p}\left(\mathbf{d}_{j}|\lambda_{j}\right) &\propto \lambda_{j}^{L_{j}-1}\exp\left(-\frac{\lambda_{j}}{2}\mathbf{d}_{j}^{T}\mathbf{P}_{\mathbf{d},j}\mathbf{d}_{j}\right) \\ \lambda_{j} &\sim \mathcal{E}xp\left(10^{-8}\right) \\ \tau_{j} &\sim \mathcal{G}\left(a_{\tau_{j}}; b_{\tau_{j}}\right) \\ \theta &\sim \mathbf{p}\left(\theta\right). \end{aligned}$$

ODE-model validation (Jaeger and Lambert, 2012b)

If the ODE model is adequate, then one can show that:

$$\lim_{\gamma \to \infty} \left( \log p\left(\gamma | \boldsymbol{\theta}, \boldsymbol{\tau}, \mathbf{y}\right) - \log p\left(\gamma\right) \right) \doteq 0$$

when no prior information is used for the state functions and one common ODE-adhesion parameter  $\gamma$  for the complete ODE, where:

- log  $(p(\gamma|\theta, \tau, \mathbf{y}))$  is the log conditional posterior distribution for  $\gamma$  (marginalized w.r.t the spline coefficients),

-  $\log(p(\gamma))$  is the log prior distribution for  $\gamma$ .

# Application

# Theophylinne dataset

#### **Differential equation**

$$\begin{cases} \frac{dx_1}{dt}(t) &= -k_a x_1(t), \\ \frac{dx_2}{dt}(t) &= \frac{k_a}{V} x_1(t) - k_e x_2(t), \\ x_1(0) &= D \\ x_2(0) &= 0. \end{cases}$$

-  $x_1(t)$  quantity of drug in stomach, -  $x_2(t)$  concentration of drug in blood.



 $\label{eq:Figure: Plasma concentration of the ophylline over time \\$ 

# Theophylinne dataset

#### **Prior information**

- The drug cannot be eliminated more quickly that it is absorbed:  $k_a > k_e$ , - Total confidence on the initial condition of the state function



 $\label{eq:Figure: Plasma concentration of the ophylline over time \\$ 

# Fitted error density



Figure : Fitted error densities with pointwise 80% and 95% credibility interval

# Suitability of the ODE-model



Figure : Histograms of the posterior distribution for  $\gamma_1$  and  $\gamma_2$  with prior density (log10 scale)

Posterior credibility intervals for the quantity of drug in stomach

 ${\sf Figure}$  : Pointwise 80% (dark grey) and 95% (light grey) posterior credibility intervals for the quantity of drug in stomach with posterior median

Posterior credibility intervals for the drug concentration in the plasma

 $\mathsf{Figure}:$  Pointwise 80% (dark grey) and 95% (light grey) posterior credibility intervals for the drug concentration with posterior median

# Conclusion

# Conclusion and further work

#### Conclusion

- Powerful tool that overcomes solving the DE using a numerical method,
- Convenient implementation of the Bayesian ODE-penalized B-spline approach,
- Simple method to include prior information about ODE parameters,
- Possibility to express uncertainty with respect to initial conditions,
- Automatic selection of the ODE-adhesion parameters and of the roughness penalty parameters,
- ODE-model validation,
- Flexible distributional assumption.

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#### Current/further work

- How to deal with non-homogeneous (non) Gaussian data distribution?
- How to deal with input functions?
- Generalize this approach for nonlinear differential equations.

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