



# Moments of the maximum of the Gaussian random walk

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## Model definition

Consider the partial sums

$$S_n = X_1 + \dots + X_n$$

with

$$X_1, X_2, \dots \text{i.i.d.}, \quad X_i \sim N(-\beta, 1), \quad \beta > 0$$

The **Gaussian random walk** is then defined as

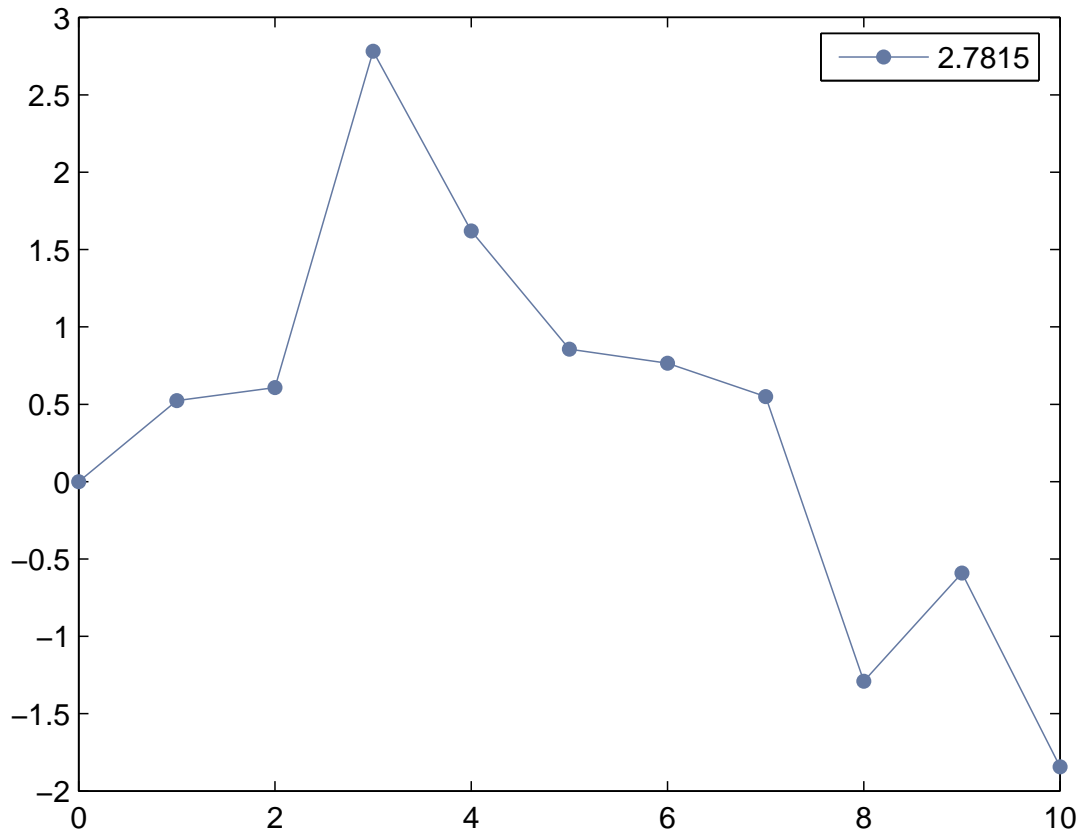
$$\{S_n : n \geq 0\} \quad ; \quad S_0 = 0$$

We are interested in the all-time maximum

$$M_\beta = \max\{S_n : n \geq 0\}$$

We consider the moments  $\mathbb{E}M_\beta^k$  and the role of the drift  $\beta$

Gaussian random walk with drift = -0.1



## Motivation for studying $\mathbb{E}M_\beta^k$

### Queues in conventional heavy traffic

Scaled version of the queue behaves **approximately** as  $M_\beta$

- *G/G/1 queue*: Kingman (1962,1965)

### Queues in Halfin-Whitt scaling

In the limit, the queue behaves **exactly** as  $M_\beta$

- *G/M/N queue*: Halfin-Whitt (1981)
- *G/D/N queue*: Jelenković-Mandelbaum-Momčilović (2004)
- *Call centers*: Borst-Mandelbaum-Reiman (2004)

### Equidistant sampling of Brownian motion

- *Testing for drift*: Chernoff (1965)
- *Corrected diffusion approximations*: Siegmund (1979,1985)
- *Option pricing*: Broadie-Glasserman-Kou (1997)

## Queues and Halfin-Whitt scaling

### Example

$$W_{\lambda,n+1} = (W_{\lambda,n} + A_{\lambda,n} - s)^+$$

where  $A_{\lambda,n}$  is Poisson distributed with mean  $\lambda < s$ . Let

$$W_\lambda = \lim_{n \rightarrow \infty} W_{\lambda,n}$$

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### *Square-root staffing*

$$s = \lambda + \beta\sqrt{\lambda}$$

and

$$W = \lim_{\lambda \rightarrow \infty} \frac{1}{\sqrt{\lambda}} W_\lambda$$

gives

$$W \stackrel{d}{=} (W + N(-\beta, 1))^+ \stackrel{d}{=} \max\{S_n : n \geq 0\} =: M_\beta$$

## Outline

1. Exact expression for  $\mathbb{E}M_\beta$
2. Exact expressions for all moments  $\mathbb{E}M_\beta^k$
3. Equidistant sampling of Brownian motion



"Despite the apparent simplicity of the problem, there does not seem to be an explicit expression even for  $\mathbb{E}M_\beta\dots$ , but it is possible to give quite sharp inequalities and asymptotic results for small  $\beta$ ."

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Kingman showed that for  $\beta \downarrow 0$

$$\mathbb{E}M_\beta = \frac{1}{2\beta} - c + \mathcal{O}(\beta)$$

where

$$c = \frac{1}{\sqrt{2\pi}} \sum_{n=1}^{\infty} \frac{1}{\sqrt{\sqrt{n}(\sqrt{n} + \sqrt{n-1})^2}} \approx 0.58$$

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Kingman showed that for  $\beta \downarrow 0$

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where

$$c = \frac{1}{\sqrt{2\pi}} \sum_{n=1}^{\infty} \frac{1}{\sqrt{\sqrt{n}(\sqrt{n} + \sqrt{n-1})^2}} = -\frac{\zeta(\frac{1}{2})}{\sqrt{2\pi}}$$

## Riemann zeta function

The Riemann zeta function  $\zeta$  is defined as

$$\zeta(s) = \sum_{n=1}^{\infty} n^{-s}, \quad \operatorname{Re} s > 1$$

This definition is extended by analytic continuation to the entire complex plane except  $s = 1$ , where  $\zeta$  has a simple pole.

Calculation of  $\zeta$  is routine.

## Theorem

For  $0 < \beta < 2\sqrt{\pi}$  we have

$$\mathbb{E}M_\beta = \frac{1}{2\beta} + \frac{\zeta\left(\frac{1}{2}\right)}{\sqrt{2\pi}} + \frac{1}{4}\beta + \frac{\beta^2}{\sqrt{2\pi}} \sum_{r=0}^{\infty} \frac{\zeta\left(-\frac{1}{2} - r\right)}{r!(2r+1)(2r+2)} \left(\frac{-\beta^2}{2}\right)^r$$

## Proof

For  $M_\beta = \max\{S_n : n \geq 0\}$  we have from **Spitzer's identity**

$$J_1(\beta) := \mathbb{E}M_\beta = \sum_{n=1}^{\infty} \frac{1}{n} \mathbb{E}S_n^+$$

From  $S_n \sim N(-\beta n, n)$  we get

$$J_1(\beta) = \sum_{n=1}^{\infty} \frac{1}{n\sqrt{2\pi}} \int_{\beta\sqrt{n}}^{\infty} (\sqrt{n}x - \beta n) e^{-x^2/2} dx$$

Upon changing variables according to  $y = x/\sqrt{n}$  ...

## Proof (cont'd)

... we can write  $J_1(\beta)$  as

$$J_1(\beta) = \sum_{n=1}^{\infty} \frac{n^{1/2}}{\sqrt{2\pi}} \int_{\beta}^{\infty} (y - \beta) e^{-\frac{1}{2}ny^2} dy$$

Differentiating twice w.r.t.  $\beta$  yields \*

$$J_1^{(2)}(\beta) = \sum_{n=1}^{\infty} \frac{n^{1/2}}{\sqrt{2\pi}} e^{-\frac{1}{2}n\beta^2}$$

using

$$\frac{d}{d\beta} \left[ \int_{\beta}^{\infty} f(y, \beta) dy \right] = -f(\beta, \beta) + \int_{\beta}^{\infty} \frac{\partial f}{\partial \beta}(y, \beta) dy$$

\* Idea of **Chang-Peres (1997)** in their analysis of  $\mathbb{P}(M_{\beta} = 0)$

## Lerch's transcendent

Lerch's transcendent  $\Phi$  is defined as the analytic continuation of the series

$$\Phi(z, s, v) = \sum_{n=0}^{\infty} (v + n)^{-s} z^n$$

Note that  $\zeta(s) = \Phi(1, s, 1)$ .

### Lemma

For  $|\ln z| < 2\pi$ ,  $s \neq 1, 2, 3, \dots$ , and  $v \neq 0, -1, -2, \dots$  we have

$$\Phi(z, s, v) = \frac{\Gamma(1-s)}{z^v} (\ln 1/z)^{s-1} + z^{-v} \sum_{r=0}^{\infty} \zeta(s-r, v) \frac{(\ln z)^r}{r!}$$

with  $\zeta(s, v) := \Phi(1, s, v)$  the Hurwitz zeta function.

**Proof** Bateman §1.11(8)



## Proof (cont'd)

$$\begin{aligned}
 J_1^{(2)}(\beta) &= \sum_{n=1}^{\infty} \frac{n^{1/2}}{\sqrt{2\pi}} e^{-\frac{1}{2}n\beta^2} \\
 &= \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}\beta^2} \Phi(z = e^{-\frac{1}{2}\beta^2}, s = -\frac{1}{2}, v = 1)
 \end{aligned}$$

Using Bateman's result we get for  $0 < \beta < 2\sqrt{\pi}$

$$J_1^{(2)}(\beta) - \frac{\Gamma(\frac{3}{2})}{\sqrt{2\pi}2^{3/2}} \frac{1}{\beta^3} = \frac{1}{\sqrt{2\pi}} \sum_{r=0}^{\infty} \zeta(-r - \frac{1}{2}) \frac{(-\frac{1}{2}\beta^2)^r}{r!}$$

The rhs is a well-behaved function of  $\beta$ . Integrating twice yields

$$J_1(\beta) - \frac{1}{2\beta} = L_0 + L_1\beta + \frac{1}{\sqrt{2\pi}} \sum_{r=0}^{\infty} \frac{\zeta(-r - \frac{1}{2})(-\frac{1}{2})^r \beta^{2r+2}}{r!(2r+1)(2r+2)}$$

where  $L_1$  and  $L_0$  are integration constants

## Proof (cont'd)

These can be found using **Euler-Maclaurin summation**, among other things.

$$L_0 = \frac{\zeta(1/2)}{\sqrt{2\pi}} \quad , \quad L_1 = \frac{1}{4}$$

This in total gives

$$\mathbb{E}M_\beta = \frac{1}{2\beta} + \frac{\zeta(\frac{1}{2})}{\sqrt{2\pi}} + \frac{1}{4}\beta + \frac{\beta^2}{\sqrt{2\pi}} \sum_{r=0}^{\infty} \frac{\zeta(-\frac{1}{2} - r)}{r!(2r+1)(2r+2)} \left(\frac{-\beta^2}{2}\right)^r$$

## Cumulants

The  $k$ -th cumulant of a random variable  $A$  is defined as the  $k$ -th derivative of  $\log \mathbb{E}e^{sA}$  evaluated at  $s = 0$ .

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The  $k$ -th cumulant of a random variable  $A$  is defined as the  $k$ -th derivative of  $\log \mathbb{E}e^{sA}$  evaluated at  $s = 0$ .

Spitzer's identity leads to

$$\mathbb{E}(e^{sM_\beta}) = \exp \left\{ \sum_{n=1}^{\infty} \frac{1}{n} \mathbb{E}(e^{sS_n^+} - 1) \right\}, \quad \operatorname{Re} s \leq 0$$

Thus

$$\log \mathbb{E}(e^{sM_\beta}) = \sum_{n=1}^{\infty} \frac{1}{n} \int_0^{\infty} (sx + \frac{1}{2}s^2x^2 + \dots) f_{S_n^+}(x) dx$$

with  $f_{S_n^+}$  the density function of  $S_n^+$ , and

$$\frac{d^k}{(ds)^k} \log \mathbb{E}(e^{sM_\beta}) \Big|_{s=0} = \sum_{n=1}^{\infty} \frac{1}{n} \mathbb{E}((S_n^+)^k) =: J_k(\beta), \quad k = 1, 2, \dots$$

## Recall

$$J_1(\beta) = \mathbb{E}M_\beta \quad ; \quad J_2(\beta) = \text{Var}M_\beta \quad ; \quad J_3(\beta) = \mathbb{E}(M_\beta - \mathbb{E}M_\beta)^3$$

and all moments of  $M_\beta$  follow from the cumulants and vice versa.

Using  $S_n \sim N(-\beta n, n)$ , it follows that

$$J_k(\beta) = \sum_{n=1}^{\infty} \frac{1}{n\sqrt{2\pi}} \int_{\beta\sqrt{n}}^{\infty} (\sqrt{n}x - \beta n)^k e^{-x^2/2} dx$$

It obviously holds that

$$J_0(\beta) = \sum_{n=1}^{\infty} \frac{1}{n} \mathbb{P}(S_n > 0)$$

From Spitzer's identity we then know that

$$J_0(\beta) = -\ln \mathbb{P}(M_\beta = 0)$$

## Theorem

Assume  $0 < \beta < 2\sqrt{\pi}$ . There holds

$$J_0(\beta) \stackrel{*}{=} -\ln \beta - \frac{\ln 2}{2} - \frac{\zeta(\frac{1}{2})}{\sqrt{2\pi}}\beta - \frac{1}{\sqrt{2\pi}} \sum_{r=1}^{\infty} \frac{\zeta(-r + \frac{1}{2})(-\frac{1}{2})^r \beta^{2r+1}}{r!(2r+1)}$$

and for  $k = 1, 2, \dots$

$$J_k(\beta) = \frac{(k-1)!}{2^k} \beta^{-k} + \sum_{j=0}^k \binom{k}{j} \frac{(-1)^j \Gamma(\frac{k-j+1}{2})}{\sqrt{2\pi}} \zeta(-\frac{1}{2}k - \frac{1}{2}j + 1) 2^{\frac{k-j-1}{2}} \beta^j$$

$$+ \frac{(-1)^{k+1} k!}{\sqrt{2\pi}} \sum_{r=0}^{\infty} \frac{\zeta(-k - r + \frac{1}{2})(-\frac{1}{2})^r \beta^{2r+k+1}}{r!(2r+1) \cdots (2r+k+1)}$$

\* Result of **Chang-Peres (1997)**

## Proof

1. Use Spitzer's identity and normality of  $S_n$  to obtain

$$J_k(\beta) = \sum_{n=1}^{\infty} \frac{n^{k-1/2}}{\sqrt{2\pi}} \int_{\beta}^{\infty} (y - \beta)^k e^{-\frac{1}{2}ny^2} dy$$

2. Differentiate  $k + 1$  times under the integral sign

$$J_k^{(k+1)}(\beta) = (-1)^{k+1} k! \sum_{n=1}^{\infty} \frac{n^{k-1/2}}{\sqrt{2\pi}} e^{-\frac{1}{2}n\beta^2}$$

3. Rewrite the expression in terms of **Lerch's transcendent**

4. Apply **Bateman's result**

5. Integrate  $k + 1$  times and find the  $k + 1$  integration constants using **Euler-Maclaurin summation**

## Equidistant sampling of Brownian motion

Let  $\{B_t : t \geq 0\}$  be a BM with  $B_0 = 0$ , drift  $-\beta$  and variance 1, so that

$$B_t = -\beta t + W_t,$$

where  $\{W_t : t \geq 0\}$  is a Wiener process. Let

$$\tilde{M} = \max\{B_t : t \geq 0\}$$

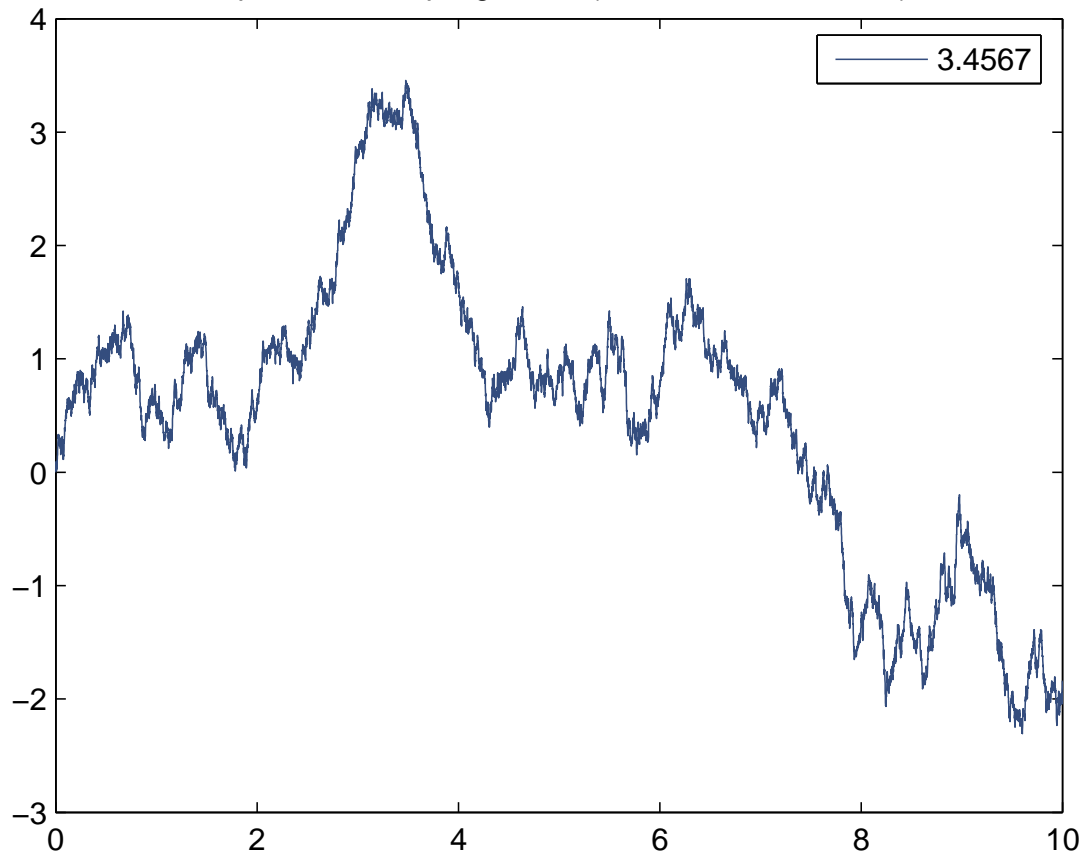
It is well known that  $\mathbb{P}(\tilde{M} \geq x) = e^{-2\beta x}$  and so the  $k$ -th cumulant of  $\tilde{M}$  equals

$$\frac{(k-1)!}{2^k} \beta^{-k}$$

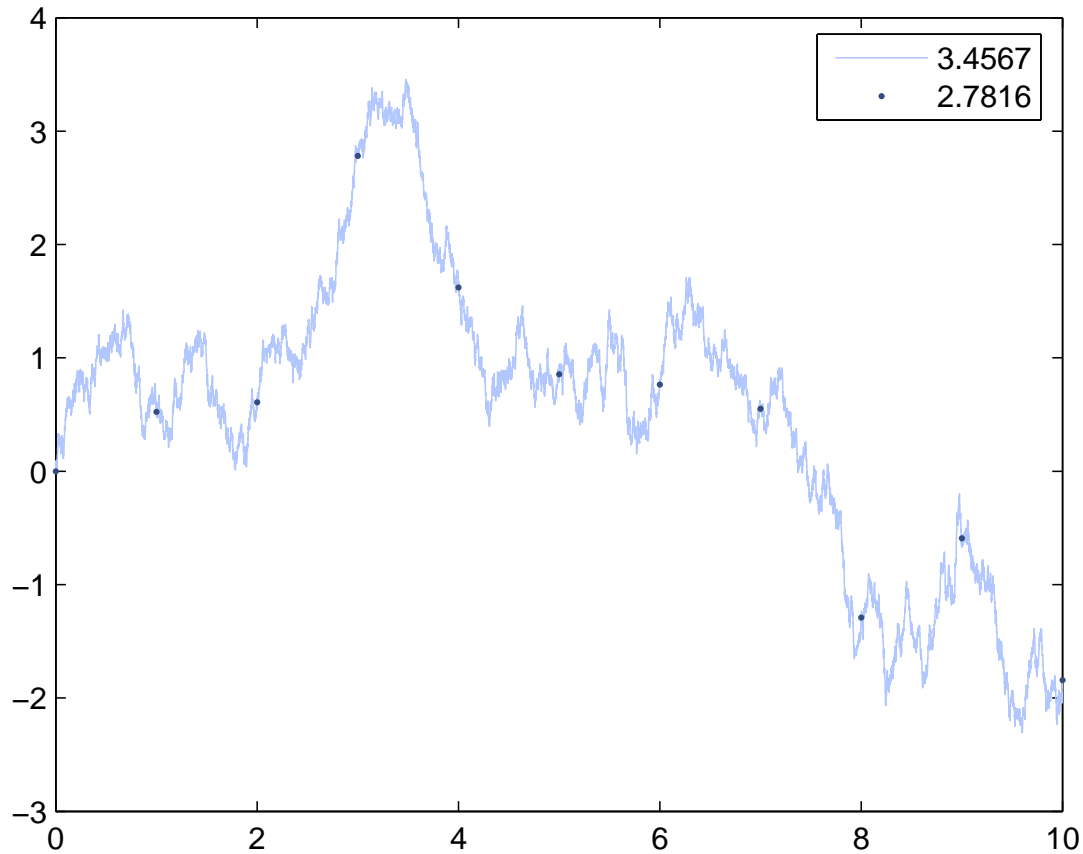


The GRW results from (equidistantly) sampling this BM, and by increasing the sampling frequency the GRW will converge to the BM.

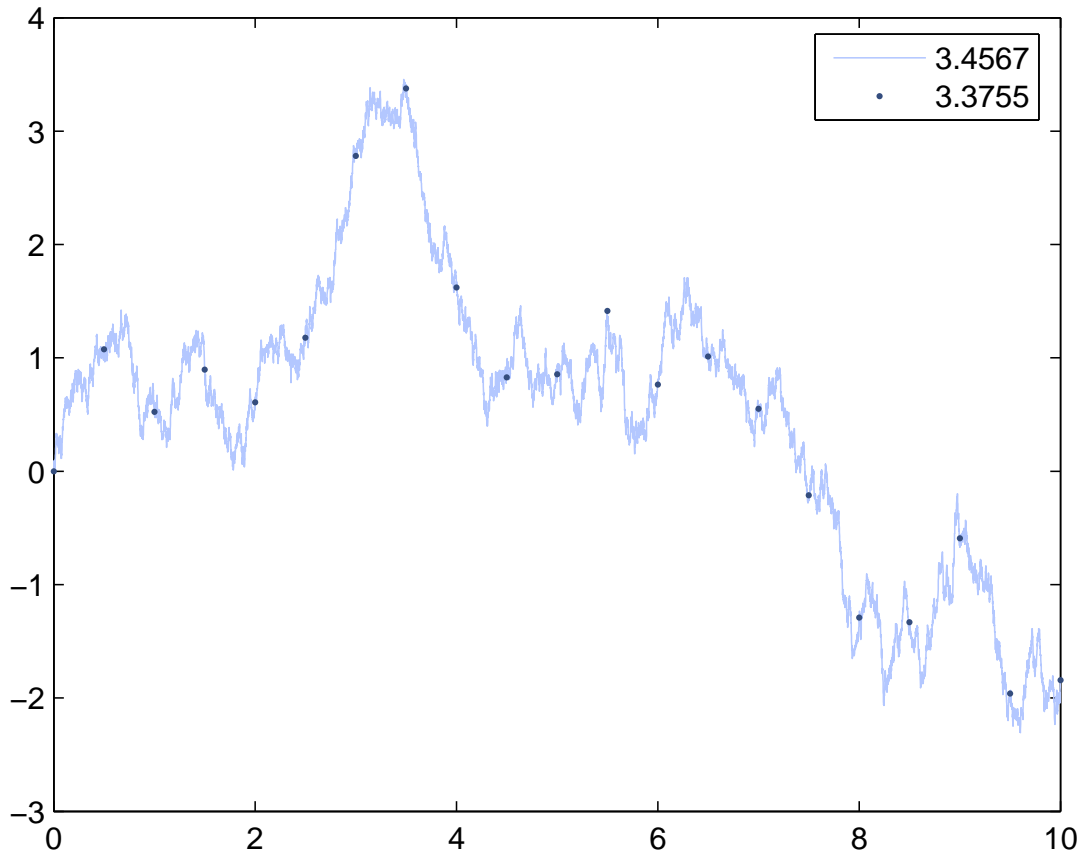
Equidistant sampling of BM (drift=-0.1, variance=1).



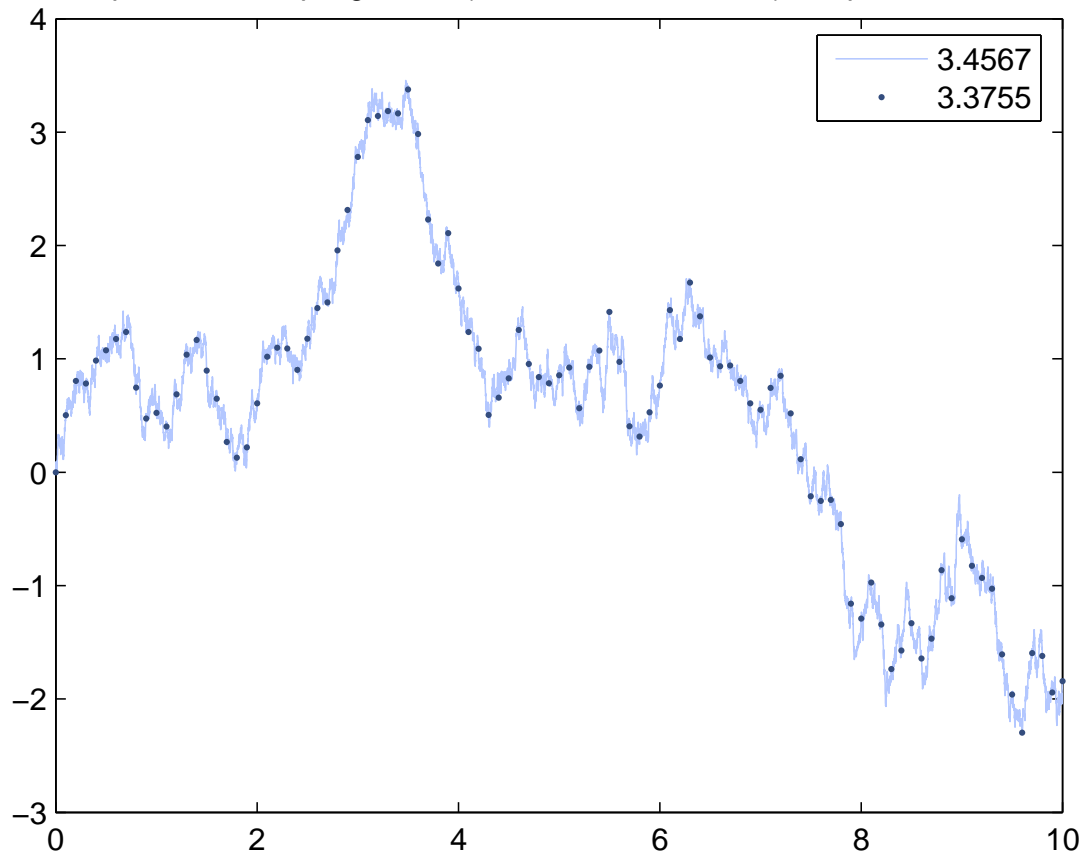
Equidistant sampling of BM (drift=-0.1, variance=1). Steps of size: 1.0.



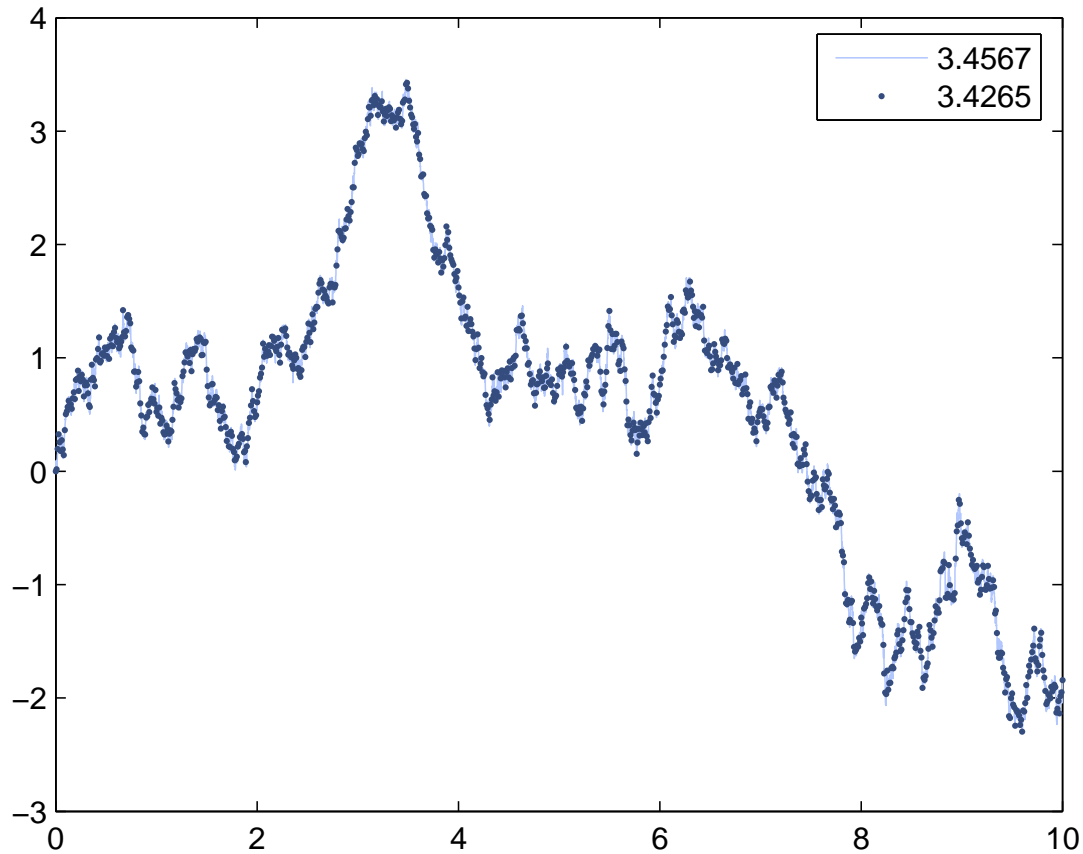
Equidistant sampling of BM (drift=-0.1, variance=1). Steps of size: 0.5.



Equidistant sampling of BM (drift=-0.1, variance=1). Steps of size: 0.1.



Equidistant sampling of BM (drift=-0.1, variance=1). Steps of size: 0.01.



Redefine the GRW as

$$\{S_n(\beta, \nu) : n = 0, 1, \dots\},$$

where

$$S_0(\beta, \nu) = 0 \quad , \quad S_n(\beta, \nu) = X_{\nu,1} + \dots + X_{\nu,n}$$

with

$$X_{\nu,1}, X_{\nu,2}, \dots \text{ i.i.d. } \quad , \quad X_i \sim N(-\beta/\nu, 1/\nu)$$

Let

$$M_{\nu,\beta} = \max\{S_n(\beta, \nu) : n = 0, 1, \dots\}.$$

Earlier definition corresponds to  $\nu = 1$  with  $M_{1,\beta} =: M_\beta$ .

Since

$$M_{\nu,\beta} \stackrel{d}{=} \nu^{-1/2} M_{\nu^{-1/2}\beta},$$

all characteristics of  $M_{\nu,\beta}$  can be expressed in those of  $M_\beta$ .

Say we sample the BM at points

$$0, \frac{1}{\nu}, \frac{2}{\nu}, \frac{3}{\nu}, \dots$$

with  $\nu$  some positive integer. From our results on  $\mathbb{E}M_\beta$  we find that

$$\mathbb{E}\tilde{M} - \mathbb{E}M_{\nu,\beta} = -\frac{\zeta(\frac{1}{2})}{\sqrt{2\pi\nu}} + \mathcal{O}(1/\nu)$$

Similar result obtained by **Asmussen, Glynn and Pitman (1995)**.



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Similar result obtained by **Asmussen, Glynn and Pitman (1995)**. However, we can easily obtain

$$\mathbb{E}\tilde{M} - \mathbb{E}M_{\nu,\beta} = -\frac{\zeta(\frac{1}{2})}{\sqrt{2\pi\nu}} + \frac{\beta}{4\nu} + \mathcal{O}(\nu^{-3/2})$$

Moreover, our exact analysis of  $M_\beta$  leads to asymptotic expressions up to any order, for all cumulants of the maximum. E.g.

$$\text{Var}\tilde{M} - \text{Var}M_{\nu,\beta} = -\frac{1}{4\nu} - \frac{2\zeta(-\frac{1}{2})}{\sqrt{2\pi}} \frac{\beta}{\nu^{3/2}} + \mathcal{O}(\nu^{-2})$$

## Other, related and future work

1. Exact expressions for  $\mathbb{P}(M = 0)$ ,  $\mathbb{E}M_\beta$ ,  $\text{Var}M_\beta$   
(with A.J.E.M. Janssen)
2. Exact expressions for all cumulants  $J_k(\beta)$ , bounds, analytic continuation for all values of  $\beta > 0$  (not only for  $0 < \beta < 2\sqrt{\pi}$ )  
(with A.J.E.M. Janssen)
3. Discrete queue and Halfin-Whitt scaling  
(with A.J.E.M. Janssen and Bert Zwart)

$$\lim_{\lambda \rightarrow \infty} \frac{1}{\sqrt{\lambda}} W_\lambda \stackrel{d}{=} \max\{S_n : n \geq 0\} = M_\beta$$