

# Analysis of an on-line algorithm for page importance computation

Nelly Litvak University of Twente n.litvak@ewi.utwente.nl

Joint work with Philippe Robert (INRIA, France)

## Outline



- PageRank algorithm
- On-line methods
- Convergence of the algorithm
- The Cat and Mouse process



- A user types a query to find relevant pages.
- Problem: Normally, there are hundreds of relevant pages. In which order should we list the pages for the user??

- A user types a query to find relevant pages.
- Problem: Normally, there are hundreds of relevant pages. In which order should we list the pages for the user??
- The 'best'-text-match-techniques quickly became insufficient (besides, they are not user-friendly)

- A user types a query to find relevant pages.
- Problem: Normally, there are hundreds of relevant pages. In which order should we list the pages for the user??
- The 'best'-text-match-techniques quickly became insufficient (besides, they are not user-friendly)
- S. Brin and L. Page (1998), J.M. Kleinberg (1999)

Idea: List most important and popular pages first. Define the importance through the hyperlink structure

- A user types a query to find relevant pages.
- Problem: Normally, there are hundreds of relevant pages. In which order should we list the pages for the user??
- The 'best'-text-match-techniques quickly became insufficient (besides, they are not user-friendly)
- S. Brin and L. Page (1998), J.M. Kleinberg (1999)

Idea: List most important and popular pages first. Define the importance through the hyperlink structure

S. Brin, L. Page, R. Motwami and T. Winograd (1998) The PageRank citation ranking: bringing order to the web.

- PageRank  $\pi_i$  of page *i* is the long run fraction of time that a random surfer spends on page *i*.
- 'Easily bored surfer' model. With probability *c* (=0.85), a surfer follows a randomly chosen outgoing link. Otherwise, he/she jumps to a random page.



$$\pi_i = \sum_{j \to i} \frac{c}{d_j} \pi_j + \frac{1-c}{N}$$

N – # pages pages  $d_j$  – out-degree of j $c < 1 \Rightarrow$  solution exists

- PageRank  $\pi_i$  of page *i* is the long run fraction of time that a random surfer spends on page *i*.
- 'Easily bored surfer' model. With probability *c* (=0.85), a surfer follows a randomly chosen outgoing link. Otherwise, he/she jumps to a random page.



$$\pi_i = \sum_{j \to i} \frac{c}{d_j} \pi_j + \frac{1-c}{N}$$

N - # pages pages  $d_j -$ out-degree of j $c < 1 \Rightarrow$  solution exists

Page is important if many important pages link to it!

- *n* is the total number of pages
- $P = (p_{ij})$  hyperlink matrix

 $p_{ij} = \begin{cases} 1/d_i & \text{if } j \text{ is one of the } d_i \text{ outgoing links of } i \\ 1/n & \text{if } d_i = 0 \\ 0 & \text{otherwise} \end{cases}$ 

- Modified transition matrix:  $\tilde{P} = cP + (1 c)(1/N)E$ *E* is an  $n \times n$  matrix consisting of one's, c = 0.85
- PageRank vector:  $\pi \tilde{P} = \pi$ ,  $\pi \underline{1} = 1$

- n is the total number of pages
- $P = (p_{ij})$  hyperlink matrix

$$p_{ij} = \begin{cases} 1/d_i & \text{if } j \text{ is one of the } d_i \text{ outgoing links of } i \\ 1/n & \text{if } d_i = 0 \\ 0 & \text{otherwise} \end{cases}$$

• Modified transition matrix:  $\tilde{P} = cP + (1 - c)(1/N)E$ *E* is an  $n \times n$  matrix consisting of one's, c = 0.85

• PageRank vector: 
$$\pi \tilde{P} = \pi$$
,  $\pi \underline{1} = 1$ 



• Power Iterations:  $\pi^{(0)} = (1/N, ..., 1/N); \quad \pi^{(t)} = \pi^{(t-1)}\tilde{P}, \ t > 0$ 

Accuracy of the order  $c^t$  (50–100 iterations with c = 0.85)

Nelly Litvak

• PageRank is a stationary distribution of a huge Markov chain: state space is the set of all Web pages.

- PageRank is a stationary distribution of a huge Markov chain: state space is the set of all Web pages.
- The task of PageRank computation trigged enormous developments in the well established area of numerical solution of large MC.

- PageRank is a stationary distribution of a huge Markov chain: state space is the set of all Web pages.
- The task of PageRank computation trigged enormous developments in the well established area of numerical solution of large MC.
  - Advanced linear algebra methods to speed up power iterations (off-line)

- PageRank is a stationary distribution of a huge Markov chain: state space is the set of all Web pages.
- The task of PageRank computation trigged enormous developments in the well established area of numerical solution of large MC.
  - Advanced linear algebra methods to speed up power iterations (off-line)
  - Monte Carlo methods (off-line or on-line)

- PageRank is a stationary distribution of a huge Markov chain: state space is the set of all Web pages.
- The task of PageRank computation trigged enormous developments in the well established area of numerical solution of large MC.
  - Advanced linear algebra methods to speed up power iterations (off-line)
  - Monte Carlo methods (off-line or on-line)
  - Other non-trivial on-line methods. One such method by Abiteboul, Preda and Cobena (1999) we will discuss today.

## **On-line algorithm by Abiteboul, Preda, Cobena (1999)**

University of Twente Enschede - The Netherlands

• Consider a crawler that performs a Markov random walk  $\{C(t), t = 0, 1, ...\}$  on directed graph of N nodes.



- Consider a crawler that performs a Markov random walk  $\{C(t), t = 0, 1, ...\}$  on directed graph of N nodes.
- Transition matrix P, where  $p_{ij} > 0$  iff there is a link (edge) from i to j. The goal is to define the stationary distribution  $\pi$ .



- Consider a crawler that performs a Markov random walk  $\{C(t), t = 0, 1, ...\}$  on directed graph of N nodes.
- Transition matrix P, where  $p_{ij} > 0$  iff there is a link (edge) from i to j. The goal is to define the stationary distribution  $\pi$ .
- Algorithm:
  - At t = 0, each node receives an equal amount 1/N of cash.

- Consider a crawler that performs a Markov random walk  $\{C(t), t = 0, 1, ...\}$  on directed graph of N nodes.
- Transition matrix P, where  $p_{ij} > 0$  iff there is a link (edge) from i to j. The goal is to define the stationary distribution  $\pi$ .
- Algorithm:
  - At t = 0, each node receives an equal amount 1/N of cash.
  - Each time the crawler visits a node i, the node distributes all its cash among its outgoing links proportional to  $p_{ij}$ .

- Consider a crawler that performs a Markov random walk  $\{C(t), t = 0, 1, ...\}$  on directed graph of N nodes.
- Transition matrix P, where  $p_{ij} > 0$  iff there is a link (edge) from i to j. The goal is to define the stationary distribution  $\pi$ .
- Algorithm:
  - At t = 0, each node receives an equal amount 1/N of cash.
  - Each time the crawler visits a node i, the node distributes all its cash among its outgoing links proportional to  $p_{ij}$ .
  - $X_i(t)$ : the amount of cash at node *i* at time *t*

- Consider a crawler that performs a Markov random walk  $\{C(t), t = 0, 1, ...\}$  on directed graph of N nodes.
- Transition matrix P, where  $p_{ij} > 0$  iff there is a link (edge) from i to j. The goal is to define the stationary distribution  $\pi$ .
- Algorithm:
  - At t = 0, each node receives an equal amount 1/N of cash.
  - Each time the crawler visits a node i, the node distributes all its cash among its outgoing links proportional to  $p_{ij}$ .
  - $X_i(t)$ : the amount of cash at node *i* at time *t*
  - History of a node:  $H_i(t) = \sum_{s=0}^t X_i(s) \mathbf{1}_{[C(s)=i]}$ , amount of cash distributed by node i on [0, t]Total history:  $H(t) = \sum_{i=1}^n H_i(t)$

- Consider a crawler that performs a Markov random walk  $\{C(t), t = 0, 1, ...\}$  on directed graph of N nodes.
- Transition matrix P, where  $p_{ij} > 0$  iff there is a link (edge) from i to j. The goal is to define the stationary distribution  $\pi$ .
- Algorithm:
  - At t = 0, each node receives an equal amount 1/N of cash.
  - Each time the crawler visits a node i, the node distributes all its cash among its outgoing links proportional to  $p_{ij}$ .
  - $X_i(t)$ : the amount of cash at node *i* at time *t*
  - History of a node:  $H_i(t) = \sum_{s=0}^t X_i(s) \mathbf{1}_{[C(s)=i]}$ , amount of cash distributed by node i on [0, t]Total history:  $H(t) = \sum_{i=1}^n H_i(t)$
  - The estimator of  $\pi_i$  at time t is

$$\pi_i(t) = \frac{H_i(t) + X_i(t)}{H(t) + 1}$$

YEP March 11 2008 / 7

#### Nelly Litvak

- $X_i(t)$ : the amount of cash at node *i* at time *t*
- History of a node:  $H_i(t) = \sum_{k=0}^t X_i(k) \mathbf{1}_{[C(k)=i]}$ , amount of cash distributed by node *i* on [0, t]. Total history:  $H(t) = \sum_{i=1}^n H(t)$
- The estimator of  $\pi_i$  at time t is  $\pi_i(t) = (H_i(t) + X_i(t))/(H(t) + 1)$ .

- $X_i(t)$ : the amount of cash at node *i* at time *t*
- History of a node:  $H_i(t) = \sum_{k=0}^t X_i(k) \mathbf{1}_{[C(k)=i]}$ , amount of cash distributed by node *i* on [0, t]. Total history:  $H(t) = \sum_{i=1}^n H(t)$
- The estimator of  $\pi_i$  at time t is  $\pi_i(t) = (H_i(t) + X_i(t))/(H(t) + 1)$ .

- $X_i(t)$ : the amount of cash at node *i* at time *t*
- History of a node:  $H_i(t) = \sum_{k=0}^t X_i(k) \mathbf{1}_{[C(k)=i]}$ , amount of cash distributed by node *i* on [0, t]. Total history:  $H(t) = \sum_{i=1}^n H(t)$
- The estimator of  $\pi_i$  at time t is  $\pi_i(t) = (H_i(t) + X_i(t))/(H(t) + 1)$ .

Pf. Cover time: a time needed to visit all states and come back to the original state. The cover time is finite w.p. 1. During one cover time cycle, H(t) grows by at least 1 cash unit.

- $X_i(t)$ : the amount of cash at node *i* at time *t*
- History of a node:  $H_i(t) = \sum_{k=0}^t X_i(k) \mathbf{1}_{[C(k)=i]}$ , amount of cash distributed by node *i* on [0, t]. Total history:  $H(t) = \sum_{i=1}^n H(t)$
- The estimator of  $\pi_i$  at time t is  $\pi_i(t) = (H_i(t) + X_i(t))/(H(t) + 1)$ .

Pf. Cover time: a time needed to visit all states and come back to the original state. The cover time is finite w.p. 1. During one cover time cycle, H(t) grows by at least 1 cash unit.

Theorem.  $\pi_i(t) \to \pi$  a.s. as  $t \to \infty$ .

- $X_i(t)$ : the amount of cash at node *i* at time *t*
- History of a node:  $H_i(t) = \sum_{k=0}^t X_i(k) \mathbf{1}_{[C(k)=i]}$ , amount of cash distributed by node *i* on [0, t]. Total history:  $H(t) = \sum_{i=1}^n H(t)$
- The estimator of  $\pi_i$  at time t is  $\pi_i(t) = (H_i(t) + X_i(t))/(H(t) + 1)$ .

Pf. Cover time: a time needed to visit all states and come back to the original state. The cover time is finite w.p. 1. During one cover time cycle, H(t) grows by at least 1 cash unit.

Theorem.  $\pi_i(t) \to \pi$  a.s. as  $t \to \infty$ . Pf. From the APC paper

$$H_i(t) + C_i(t) = 1/n + \sum_j p_{ji} H_j(t), \quad i = 1, \dots, n$$

Divide by H(t):  $\pi_i(t) = [1/n - C_i(t)]/H(t) + \sum_j p_{ji}\pi_j(t), \quad i = 1, ..., n$ Solution:  $\pi_i(t) = \pi_i + \left[ \frac{(1/n - \mathbf{C}(t))}{H(t)} \left( \sum_{k=0}^{\infty} P^k - \mathbf{1}^T \pi \right) \right]_i$ (here 1 is a row-vector of ones)

Nelly Litvak

• The speed of convergence is determined by the term [(1/n - C(t))/H(t)], with H(t) in denominator. Thus, the algorithm converges as 1/H(t) when  $t \to \infty$ .

- The speed of convergence is determined by the term [(1/n C(t))/H(t)], with H(t) in denominator. Thus, the algorithm converges as 1/H(t) when  $t \to \infty$ .
- $H(t) = \sum_{s=0}^{t} \sum_{i} X_i(s) \mathbf{1}_{[C(s)=i]}$ . The cash process  $\{X_i(t)\}$  is determining for the speed of convergence of the algorithm.

- The speed of convergence is determined by the term [(1/n C(t))/H(t)], with H(t) in denominator. Thus, the algorithm converges as 1/H(t) when  $t \to \infty$ .
- $H(t) = \sum_{s=0}^{t} \sum_{i} X_i(s) \mathbf{1}_{[C(s)=i]}$ . The cash process  $\{X_i(t)\}$  is determining for the speed of convergence of the algorithm.
- But how can we characterize the cash process? For instance, can we prove that  $X_i(t)$  converges, or, at least,  $\mathbb{E}[X_i(t)]$  converges?

- The speed of convergence is determined by the term [(1/n C(t))/H(t)], with H(t) in denominator. Thus, the algorithm converges as 1/H(t) when  $t \to \infty$ .
- $H(t) = \sum_{s=0}^{t} \sum_{i} X_i(s) \mathbf{1}_{[C(s)=i]}$ . The cash process  $\{X_i(t)\}$  is determining for the speed of convergence of the algorithm.
- But how can we characterize the cash process? For instance, can we prove that  $X_i(t)$  converges, or, at least,  $\mathbb{E}[X_i(t)]$  converges?
- The cash process is inconvenient for analysis. Idea: to translate it into an easier process.





Nelly Litvak





• Cat performs the Markov random walk C(t), transition matrix P.



- Cat performs the Markov random walk C(t), transition matrix P.
- At time t, the mouse is at the node M(t).



- Cat performs the Markov random walk C(t), transition matrix P.
- At time t, the mouse is at the node M(t).
- The mouse makes a move only when found by the cat, the event [C(t) = M(t)]. In this case, the mouse makes one step, using the same transition matrix P.



- Cat performs the Markov random walk C(t), transition matrix P.
- At time t, the mouse is at the node M(t).
- The mouse makes a move only when found by the cat, the event [C(t) = M(t)]. In this case, the mouse makes one step, using the same transition matrix P.

### **Relation to the cash process**

- Cat performs the Markov random walk C(t), transition matrix P.
- At time t, the mouse is at the node M(t).
- The mouse makes a move only when found by the cat, the event [C(t) = M(t)]. In this case, the mouse makes one step, using the same transition matrix P.

 $\mathcal{F}_t$ :  $\sigma$ -field generated by C(0), C(1), ..., C(t).

### **Relation to the cash process**

- Cat performs the Markov random walk C(t), transition matrix P.
- At time t, the mouse is at the node M(t).
- The mouse makes a move only when found by the cat, the event [C(t) = M(t)]. In this case, the mouse makes one step, using the same transition matrix P.
- $\mathcal{F}_t$ :  $\sigma$ -field generated by C(0), C(1), ..., C(t).

Theorem. For  $t \ge 0$ ,

$$(X_i(t), 1 \leq i \leq N) \stackrel{\text{dist.}}{=} (\mathbb{P}[M(t) = i \mid \mathcal{F}_{t-1}], 1 \leq i \leq N).$$

In particular, for  $1 \le i \le N$ ,

$$\mathbb{E}(X_i(t)) = \mathbb{P}(M(t) = i).$$

Nelly Litvak

C(t) – cat position; M(t) – mouse position;  $X_i(t)$  – cash on page *i*. We need to prove that:  $\mathbb{E}(X_i(t)) = \mathbb{P}(M(t) = i)$ .

$$\mathbb{P}(M(t+1) = i \mid \mathcal{F}_t) = \sum_{j \neq i} \mathbb{1}_{\{C(t) = j\}} p_{j,i} \times \mathbb{P}(M(t) = j \mid \mathcal{F}_{t-1}) + \mathbb{1}_{\{C(t) \neq i\}} \mathbb{P}(M(t) = i \mid \mathcal{F}_{t-1}).$$

On the other hand, for the cash

$$\mathbb{E}[X_i(t + 1)] = \sum_{j \neq i} \mathbb{1}_{\{C(t) = j\}} p_{j,i} \times \mathbb{E}[X_j(t)] + \mathbb{1}_{\{C(t) \neq i\}} \mathbb{E}[X_i(t)].$$

Same equation! Statement follows by induction in t.

Nelly Litvak

## The MC (C(t), M(t)) has a transition matrix $q(\cdot, \cdot)$ given by

 $q_{(i,j),(k,j)} = p_{i,k}$  $q_{(j,j),(k,k')} = p_{j,k}p_{j,k'}.$  if  $i \neq j$ ;

if  $i \neq j$ ;

## The MC (C(t), M(t)) has a transition matrix $q(\cdot, \cdot)$ given by

$$q_{(i,j),(k,j)} = p_{i,k}$$
  
 $q_{(j,j),(k,k')} = p_{j,k}p_{j,k'}.$ 

 $\nu = \mathbb{P}(C_{\infty} = \cdot, M_{\infty} = \cdot)$  is the invariant distribution. Then

$$\nu(i,j) = \sum_{k \neq j} \nu(k,j) p_{k,i} + \sum_{k} \nu(k,k) p_{k,i} p_{k,j}$$

Nelly Litvak

if  $i \neq j$ ;

## The MC (C(t), M(t)) has a transition matrix $q(\cdot, \cdot)$ given by

$$q_{(i,j),(k,j)} = p_{i,k}$$
  
 $q_{(j,j),(k,k')} = p_{j,k}p_{j,k'}.$ 

 $\nu = \mathbb{P}(C_\infty = \cdot\,, M_\infty = \cdot)$  is the invariant distribution. Then

$$\nu(i,j) = \sum_{k \neq j} \nu(k,j) p_{k,i} + \sum_{k} \nu(k,k) p_{k,i} p_{k,j}$$

Note:  $\sum_{j} \nu(i, j) = \pi_i$ . Summing over *i* we get  $\sum_{k} \nu(k, k) p_{k,j} = \nu(j, j)$ , and therefore that there exists some constant *c* such that,

$$\nu(j,j) = c\pi_j, \quad j = 1, \dots, N.$$

YEP March 11 2008 / 13

Nelly Litvak

if  $i \neq j$ ;

## The MC (C(t), M(t)) has a transition matrix $q(\cdot, \cdot)$ given by

$$q_{(i,j),(k,j)} = p_{i,k}$$
  
 $q_{(j,j),(k,k')} = p_{j,k}p_{j,k'}.$ 

 $\nu = \mathbb{P}(C_\infty = \cdot\,, M_\infty = \cdot)$  is the invariant distribution. Then

$$\nu(i,j) = \sum_{k \neq j} \nu(k,j) p_{k,i} + \sum_{k} \nu(k,k) p_{k,i} p_{k,j}$$

Note:  $\sum_{j} \nu(i, j) = \pi_i$ . Summing over *i* we get  $\sum_{k} \nu(k, k) p_{k,j} = \nu(j, j)$ , and therefore that there exists some constant *c* such that,

$$\nu(j,j) = c\pi_j, \quad j = 1, \dots, N.$$

We have  $\mathbb{E}(X_j(t)|C_t = j) = \mathbb{P}(M_t = j|C_t = j) \rightarrow c \text{ as } t \rightarrow \infty$ . On average, the 'transaction' is *c* at each step. Thus,  $\mathbb{E}[H(t)] \sim ct$ .

Nelly Litvak



Denote by

$$p_{i,j}^* = \frac{\pi_j}{\pi_i} p_{j,i}$$

the transition matrix of the reversed Markov chain  $(C^{\ast}(t))$  associated to (C(t)). Let

$$T_j = \inf\{t > 0 : C(t) = j\}.$$



#### Denote by

$$p_{i,j}^* = \frac{\pi_j}{\pi_i} p_{j,i}$$

the transition matrix of the reversed Markov chain  $(C^{\ast}(t))$  associated to (C(t)). Let

$$T_j = \inf\{t > 0 : C(t) = j\}.$$

## Proposition. The stationary distribution of the mouse is

$$\mathbb{P}(M(\infty) = j) = c\mathbb{E}_{\pi} \left[ p_{C(0),j} T_j \right]$$

and

$$c = \left[\sum_{k} \mathbb{E}_{\pi} \left[ p_{C(0),k} T_{k} \right] \right]^{-1}$$

Nelly Litvak

It is assumed for the moment that (C(t)) is a reversible Markov chain, i.e.  $\pi_i p_{ij} = \pi_j p_{ji}$ 

$$\mathbb{E}_{\pi} \left[ p_{C(0),j} T_j \right] = \sum_i \pi_i p_{ij} \mathbb{E}_i(T_j) = \sum_i \pi_j p_{ji} \mathbb{E}_i(T_j)$$
$$= \pi_j \mathbb{E}_j(T_j - 1) = 1 - \pi(y),$$

Consequently,

$$c = \frac{1}{N-1} \text{ and } \mathbb{P}(S_{\infty} = y) = \frac{1 - \pi(y)}{N-1}.$$

Tetali (1994) showed that if (C(t)) is a general recurrent Markov chain, then

$$\sum_{k} \mathbb{E}_{\pi} \left[ p_{C(0),k} T_k \right] \le N - 1.$$
(1)

It follows that the value c = 1/(N-1) obtained for reversible chains, is the minimal possible value of c.

Nelly Litvak



• C(t) is a simple random walk on  $\mathbb{Z}_+$  reflected at zero. If the current state is x > 0, then the next state is x + 1 or x - 1 w.p.  $p = \lambda/(\lambda + \mu)$  or  $q = \lambda/(\lambda + \mu)$ . If the current state is 0, then the next state is 1 with probability one.

- C(t) is a simple random walk on  $\mathbb{Z}_+$  reflected at zero. If the current state is x > 0, then the next state is x + 1 or x 1 w.p.  $p = \lambda/(\lambda + \mu)$  or  $q = \lambda/(\lambda + \mu)$ . If the current state is 0, then the next state is 1 with probability one.
- This is a slotted version of the M/M/1 queue with arrival rate  $\lambda$  and service rate  $\mu$ . We assume that the system is stable, that is,  $\rho = \lambda/\mu = p/q < 1$ .

- C(t) is a simple random walk on  $\mathbb{Z}_+$  reflected at zero. If the current state is x > 0, then the next state is x + 1 or x 1 w.p.  $p = \lambda/(\lambda + \mu)$  or  $q = \lambda/(\lambda + \mu)$ . If the current state is 0, then the next state is 1 with probability one.
- This is a slotted version of the M/M/1 queue with arrival rate  $\lambda$  and service rate  $\mu$ . We assume that the system is stable, that is,  $\rho = \lambda/\mu = p/q < 1$ .
- Assume that initially, the mouse is at some remote position x → ∞, while the cat just left the neighborhood of x and went back to zero. The time needed for the cat to reach zero is approximately linear in x. The time needed to come back to x, multiplied by ρ<sup>x</sup>, converges to an exponential random variable with parameter (μ λ)<sup>2</sup>/μ.

- C(t) is a simple random walk on  $\mathbb{Z}_+$  reflected at zero. If the current state is x > 0, then the next state is x + 1 or x 1 w.p.  $p = \lambda/(\lambda + \mu)$  or  $q = \lambda/(\lambda + \mu)$ . If the current state is 0, then the next state is 1 with probability one.
- This is a slotted version of the M/M/1 queue with arrival rate  $\lambda$  and service rate  $\mu$ . We assume that the system is stable, that is,  $\rho = \lambda/\mu = p/q < 1$ .
- Assume that initially, the mouse is at some remote position x → ∞, while the cat just left the neighborhood of x and went back to zero. The time needed for the cat to reach zero is approximately linear in x. The time needed to come back to x, multiplied by ρ<sup>x</sup>, converges to an exponential random variable with parameter (μ λ)<sup>2</sup>/μ.
- Note that for a finite reversible chain, c = 1/(N-1) where N is the number of states. For  $N = \infty$  we obtain that the chain (C(t), M(t)) is null-recurrent. The question is: where does the mouse stay?

• The cat and the mouse meet at some remote state x. Possibilities:

- The cat and the mouse meet at some remote state x. Possibilities:
  - mouse up, cat up  $(p^2)$

- The cat and the mouse meet at some remote state x. Possibilities:
  - mouse up, cat up  $(p^2)$
  - mouse down, cat down  $(q^2)$

- The cat and the mouse meet at some remote state x. Possibilities:
  - mouse up, cat up  $(p^2)$
  - mouse down, cat down  $(q^2)$
  - mouse down, cat up (pq), they will meet again soon

- The cat and the mouse meet at some remote state x. Possibilities:
  - mouse up, cat up  $(p^2)$
  - mouse down, cat down  $(q^2)$
  - mouse down, cat up (pq), they will meet again soon
  - mouse up cat down (pq), the cat leaves 'forever' w.p.  $(1 \rho^2)$ .

- The cat and the mouse meet at some remote state x. Possibilities:
  - mouse up, cat up  $(p^2)$
  - mouse down, cat down  $(q^2)$
  - mouse down, cat up (pq), they will meet again soon
  - mouse up cat down (pq), the cat leaves 'forever' w.p.  $(1 \rho^2)$ .
- Mouse makes a geometric number G of steps, where

$$\mathbb{P}(G=k) = [1 - pq(1 - \rho^2)]^k pq(1 - \rho^2), \quad k = 0, 1, \dots$$

and then the cat leaves for a long time.

- The cat and the mouse meet at some remote state x. Possibilities:
  - mouse up, cat up  $(p^2)$
  - mouse down, cat down  $(q^2)$
  - mouse down, cat up (pq), they will meet again soon
  - mouse up cat down (pq), the cat leaves 'forever' w.p.  $(1 \rho^2)$ .
- Mouse makes a geometric number G of steps, where

$$\mathbb{P}(G=k) = [1 - pq(1 - \rho^2)]^k pq(1 - \rho^2), \quad k = 0, 1, \dots$$

and then the cat leaves for a long time. A step X distributed as follows:

$$X = \begin{cases} +1 & \text{with probability } \frac{p^2 + pq\rho^2}{1 - pq(1 - \rho^2)} = \frac{p^2(1 + \rho)}{1 - pq(1 - \rho^2)}, \\ -1 & \text{with probability } \frac{q^2 + pq}{1 - pq(1 - \rho^2)} = \frac{q}{1 - pq(1 - \rho^2)}. \end{cases}$$

The last step of the mouse is always of size +1. After that the cat and the mouse will meet again after  $\approx$  exponentially distributed time with parameter  $\rho^y(\mu - \lambda)^2/\mu$ , at the mouse's current position y.

Nelly Litvak



• 'Free' continuous-time Markov process  $(\tilde{M}(t))$ : At state x, the transition rate of  $(\tilde{M}(t))$  is  $\rho^x(\mu - \lambda)^2/\mu$ . At each transition, the process  $(\tilde{M}(t))$  makes a jump of a random size

$$\Delta = 1 + \sum_{i=1}^{G} X_i$$

#### 'Free' process

• 'Free' continuous-time Markov process  $(\tilde{M}(t))$ : At state x, the transition rate of  $(\tilde{M}(t))$  is  $\rho^x(\mu - \lambda)^2/\mu$ . At each transition, the process  $(\tilde{M}(t))$  makes a jump of a random size

$$\Delta = 1 + \sum_{i=1}^{G} X_i$$

•  $\mathbb{E}[\Delta] = \mathbb{E}[G]\mathbb{E}[X] + 1 = -\rho^{-1} < 0$ , and  $\mathbb{E}[u^{\Delta}]$  is well-defined if

$$u_1 = \frac{(1 - \sqrt{1 - 4p^2})q}{2p^2} < u < \frac{(1 + \sqrt{1 - 4p^2})q}{2p^2} = u_2.$$

 $\mathbb{E}[u^{\Delta}]$  is well defined on the interval  $[1, 1/\rho]$ .

#### 'Free' process

• 'Free' continuous-time Markov process  $(\tilde{M}(t))$ : At state x, the transition rate of  $(\tilde{M}(t))$  is  $\rho^x(\mu - \lambda)^2/\mu$ . At each transition, the process  $(\tilde{M}(t))$  makes a jump of a random size

$$\Delta = 1 + \sum_{i=1}^{G} X_i$$

•  $\mathbb{E}[\Delta] = \mathbb{E}[G]\mathbb{E}[X] + 1 = -\rho^{-1} < 0$ , and  $\mathbb{E}[u^{\Delta}]$  is well-defined if

$$u_1 = \frac{(1 - \sqrt{1 - 4p^2})q}{2p^2} < u < \frac{(1 + \sqrt{1 - 4p^2})q}{2p^2} = u_2.$$

 $\mathbb{E}[u^{\Delta}]$  is well defined on the interval  $[1, 1/\rho]$ .

• After the first transition of size  $\Delta$ , the transition rate becomes  $\rho^{x+\Delta}(\mu-\lambda)^2/\mu$ . The expected time until the next transition is

$$\mathbb{E}[\rho^{-(x+\Delta)}\mu/(\mu-\lambda)^{2}] = \rho^{-x}\mu\mathbb{E}[\rho^{-\Delta}]/(\mu-\lambda)^{2} = \rho^{-x}\mu/(\mu-\lambda)^{2} \ (!)$$

Nelly Litvak



$$\bar{S}_x(t) = \frac{\tilde{S}([\rho^{-x}\mu/(\mu-\lambda)^2]t)}{x}, \quad t \ge 0.$$



$$\bar{S}_x(t) = \frac{\tilde{S}([\rho^{-x}\mu/(\mu-\lambda)^2]t)}{x}, \quad t \ge 0.$$

Lemma. For any  $t \ge 0$ ,  $\lim_{x \to \infty} \overline{M}_x(t) = \begin{cases} 1, & t < W; \\ -\infty, & t \ge W. \end{cases}$ 

where  $W = \sum_{n=0}^{\infty} [\rho^{-(\Delta_1 + ... + \Delta_n)}] E_n$  with where  $E_1, E_2, ...$  i.i.d. exponential(1) random variables, independent of the  $\Delta_i$ 's



$$\bar{S}_x(t) = \frac{\tilde{S}([\rho^{-x}\mu/(\mu-\lambda)^2]t)}{x}, \quad t \ge 0.$$

Lemma. For any  $t \ge 0$ ,  $\lim_{x \to \infty} \overline{M}_x(t) = \begin{cases} 1, & t < W; \\ -\infty, & t \ge W. \end{cases}$ 

where  $W = \sum_{n=0}^{\infty} [\rho^{-(\Delta_1 + ... + \Delta_n)}] E_n$  with where  $E_1, E_2, ...$  i.i.d. exponential(1) random variables, independent of the  $\Delta_i$ 's





$$\bar{S}_x(t) = \frac{\tilde{S}([\rho^{-x}\mu/(\mu-\lambda)^2]t)}{x}, \quad t \ge 0.$$

Lemma. For any 
$$t \ge 0$$
,  $\lim_{x \to \infty} \overline{M}_x(t) = \begin{cases} 1, & t < W; \\ -\infty, & t \ge W. \end{cases}$ 

where  $W = \sum_{n=0}^{\infty} [\rho^{-(\Delta_1 + ... + \Delta_n)}] E_n$  with where  $E_1, E_2, ...$  i.i.d. exponential(1) random variables, independent of the  $\Delta_i$ 's



Intuition: It takes time to get on some distance from x but then the drop happens in 'no time'

Nelly Litvak



• For the proof, we need to show that: 1) the scaled process converges to 1 before time W and to 0 after time W; 2) the time W is finite.



- For the proof, we need to show that: 1) the scaled process converges to 1 before time W and to 0 after time W; 2) the time W is finite.
- The random variable W satisfies

$$W \stackrel{d}{=} \rho^{-\Delta} W + E$$

This type of random variables is well-known in literature.



- For the proof, we need to show that: 1) the scaled process converges to 1 before time W and to 0 after time W; 2) the time W is finite.
- The random variable W satisfies

$$W \stackrel{d}{=} \rho^{-\Delta} W + E$$

This type of random variables is well-known in literature.

• *W* is extremely heavy-tailed:  $\mathbb{E}[W^s] < \infty$  for

- For the proof, we need to show that: 1) the scaled process converges to 1 before time W and to 0 after time W; 2) the time W is finite.
- The random variable W satisfies

$$W \stackrel{d}{=} \rho^{-\Delta} W + E$$

This type of random variables is well-known in literature.

• *W* is extremely heavy-tailed:  $\mathbb{E}[W^s] < \infty$  for

 $-\log(u_1)/\log(\rho) - 1 < s < 1.$ 

• With other random walk of the cat  $(M/M/\infty)$ , symmetric r.w., etc.) the mouse behavior is entirely different

- For the proof, we need to show that: 1) the scaled process converges to 1 before time W and to 0 after time W; 2) the time W is finite.
- The random variable W satisfies

$$W \stackrel{d}{=} \rho^{-\Delta} W + E$$

This type of random variables is well-known in literature.

• *W* is extremely heavy-tailed:  $\mathbb{E}[W^s] < \infty$  for

- With other random walk of the cat  $(M/M/\infty)$ , symmetric r.w., etc.) the mouse behavior is entirely different
- Work in progress...

- For the proof, we need to show that: 1) the scaled process converges to 1 before time W and to 0 after time W; 2) the time W is finite.
- The random variable W satisfies

$$W \stackrel{d}{=} \rho^{-\Delta} W + E$$

This type of random variables is well-known in literature.

• *W* is extremely heavy-tailed:  $\mathbb{E}[W^s] < \infty$  for

- With other random walk of the cat  $(M/M/\infty)$ , symmetric r.w., etc.) the mouse behavior is entirely different
- Work in progress...
- I hope you liked the cat and mouse...

- For the proof, we need to show that: 1) the scaled process converges to 1 before time W and to 0 after time W; 2) the time W is finite.
- The random variable W satisfies

$$W \stackrel{d}{=} \rho^{-\Delta} W + E$$

This type of random variables is well-known in literature.

• *W* is extremely heavy-tailed:  $\mathbb{E}[W^s] < \infty$  for

- With other random walk of the cat  $(M/M/\infty)$ , symmetric r.w., etc.) the mouse behavior is entirely different
- Work in progress...
- I hope you liked the cat and mouse...
- Thank you for your attention!