

DYNAMIC PROGRAMMING OPTIMIZATION OVER RANDOM DATA: THE SCALING EXPONENT FOR NEAR-OPTIMAL SOLUTIONS

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YEP - EURANDOM

"REGULAR" NEAR-MINIMAL PROBLEM

- Cost function $F : \mathbb{R}^n \rightarrow \mathbb{R}$ reaching its minimum at x^* .
- Relation between the distance $\delta = |x - x^*|$ and the difference $\varepsilon = F(x) - F(x^*)$:
if F is "smooth",

$$\varepsilon(\delta) = \inf\{F(x) - F(x^*), |x - x^*| \geq \delta\} \sim C\delta^2.$$

\implies the **scaling exponent** corresponding to this problem is 2.

NEAR-MINIMAL SOLUTIONS IN COMBINATORIAL OPTIMIZATION

[Aldous-Percus (2003)]

Consider a combinatorial optimization problem with n constituents under a suitable probability model.

As n goes to infinity, how cost increases when the optimal solution undergoes a small perturbation δ ?

Statistical physics suggests that there exists a **scaling exponent** such that the cost increases as

$$\delta^\alpha.$$

New classification of algorithms, intuition suggests that

- low value of $\alpha \Rightarrow$ near-optimal solutions obtained via 'local changes' \Rightarrow algorithmically easy.
- high value of $\alpha \Rightarrow$ near-optimal solutions obtained via long range changes (?) \Rightarrow algorithmically hard.

A TRIVIAL EXAMPLE

Let $(\xi_i, i \geq 1)$ be i.i.d. copies of a strictly positive random variable ξ , and

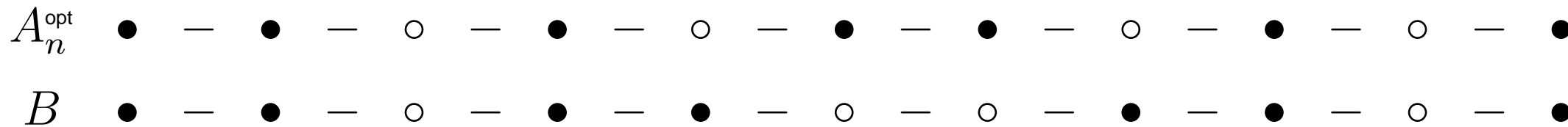
$$M_n = \max_{A \subseteq \{1,2,\dots,n\}} \sum_{i \in A} (\xi_i - 1) = \max_{A \subseteq \{1,2,\dots,n\}} f(A).$$

with $f(A) = \sum_i 1(i \in A)(\xi_i - 1)$.

The maximum is attained by choosing

$$A_n^{\text{opt}} = \{i \in [n] : \xi_i > 1\}.$$

PERTURBATION OF THE OPTIMAL SET



$\implies A_n^{\text{opt}} \triangle B = 4$. Note that a.s.

$$f(A_n^{\text{opt}}) - f(B) > 0.$$

SCALING EXPONENT

We define the random variable:

$$\varepsilon_n(\delta) = \min_{B \subseteq [n]} \left\{ \frac{f(B) - f(A_n^{\text{opt}})}{n} : |B \Delta A_n^{\text{opt}}| \geq \delta n \right\}$$

Do we have, a.s.

$$\varepsilon(\delta) := \lim_n \varepsilon_n(\delta) \sim C\delta^\alpha?$$

What is the value of α ?

Universality paradigm: α should not depend on the details of the model.

A TRIVIAL EXAMPLE

Now, fix $0 < \delta < 1$ and

$$M'_n = \max_{A' \subseteq [n]} \sum_{i \in A'} (\xi_i - 1)$$

$$\text{subject to } |A' \triangle A_n^{\text{opt}}| \geq \delta n$$

The maximum is attained by $B_n^{\text{opt}} = A_n^{\text{opt}} \triangle D$ where D is the set of indices of the $\lceil \delta n \rceil$ smallest values of $|\xi_i - 1|$.

$$\varepsilon_n(\delta) = n^{-1}(M_n - M'_n).$$

A TRIVIAL EXAMPLE

Assume that ξ has a continuous density h . So as $n \rightarrow \infty$

$$\varepsilon_n(\delta) \rightarrow_{L_1} \varepsilon(\delta) := \int_{1-a(\delta)}^{1+a(\delta)} |x - 1| h(x) dx$$

where $a(\delta)$ is defined by

$$\delta = \int_{1-a(\delta)}^{1+a(\delta)} h(x) dx.$$

So by continuity of $h(x)$, and assuming $0 < h(1) < \infty$, as $\delta \downarrow 0$ we have

$$a(\delta) \sim \frac{\delta}{2h(1)}; \quad \varepsilon(\delta) \sim a^2(\delta)h(1) \sim \frac{\delta^2}{4h(1)}$$

\Rightarrow This is the desired “scaling exponent = 2” result !

NEAR-MINIMAL SOLUTION IN COMBINATORIAL OPTIMIZATION

Under a **suitable probabilistic model**, heuristic / simulation suggest that

- (i) For the **NK-model, near minimal spanning tree**, the scaling exponent is **2**.
- (ii) For the **travelling salesman problem, minimal pair matching**, the scaling exponent is **3**.

⇒ scaling exponent **2** : algorithmically easy problem

⇒ scaling exponent **3** : algorithmically hard problem

For (i) we have rigorous proofs. For (ii) open question.

ILLUSTRATIVE EXAMPLE

Let $(\xi_i, i \geq 1)$ be i.i.d. copies of a strictly positive random variable ξ , and

$$M_n = \max_{A \subseteq [n]} \left(|A| - \sum_{i=1}^{n-1} \xi_i 1(i \in A, i+1 \in A) \right) = \max_{A \subseteq [n]} f(A).$$

$$\implies M_n/n \in [1/2, 1].$$

We may prove that if $\xi_i + \xi_{i+1} < 1$ for some $1 \leq i \leq n-2$ then a.s. A_n^{opt} is unique.

We will assume for simplicity that ξ is an $\exp(\lambda)$ variable.

\implies A.s., for n large enough, we may define A_n^{opt} as the unique optimal set.

SCALING EXPONENT

Now, fix $0 < \delta < 1$ and

$$M'_n = \max_{A' \subseteq [n]} f(A')$$

$$\text{subject to } |A' \triangle A_n^{\text{opt}}| \geq \delta n$$

and then set

$$\varepsilon_n(\delta) = n^{-1}(M_n - M'_n).$$

Do we have:

$$\lim_n \varepsilon_n(\delta) = \varepsilon(\delta) \sim C\delta^2 \quad ?$$

ANALYSIS OF THE OPTIMUM

Recall $M_n = \max_{A \subseteq \{1,2,\dots,n\}} (|A| - \sum_{i=1}^{n-1} \xi_i 1(i \in A, i+1 \in A))$.

We assume for simplicity that ξ is an exponential variable with intensity λ .

Theorem 1. *Almost surely and in L^1 ,*

$$\lim_{n \rightarrow \infty} \frac{M_n}{n} = c(\lambda) = (1 - e^{-\lambda})^{-1} - \lambda^{-1}.$$

$$\longrightarrow \lim_{\lambda \rightarrow \infty} c(\lambda) = 1, \lim_{\lambda \rightarrow 0} c(\lambda) = 1/2.$$

\implies proof based on a simple instance of Aldous' formulation of the cavity method.

IDEA OF PROOF: LEFT SIDED RECURSION

Define $V_1^L = 1$, $W_1^L = 0$, and

$$\begin{aligned} V_i^L &= \max_{i \in A \subseteq \{1, \dots, i-1, i\}} \left(|A| - \sum_{j=1}^{i-1} \xi_j 1(j \in A, j+1 \in A) \right) \\ W_i^L &= \max_{i \notin A \subseteq \{1, \dots, i-1, i\}} \left(|A| - \sum_{j=1}^{i-1} \xi_j 1(j \in A, j+1 \in A) \right) \\ X_i^L &= V_i^L - W_i^L. \end{aligned}$$

Then $M_n = \max(V_n^L, W_n^L)$, and by induction $0 \leq X_i^L \leq 1$,

$$X_{i+1}^L = 1 - \min(X_i^L, \xi_i).$$

IDEA OF PROOF: RIGHT SIDED RECURSION

Define similarly, $V_{n,n}^R = 1$, $W_{n,n}^R = 0$ and

$$\begin{aligned} V_{n,i}^R &= \max_{i \in A \subseteq \{i, i+1, \dots, n\}} \left(|A| - \sum_{j=i}^{n-1} \xi_j 1(j \in A, j+1 \in A) \right) \\ W_{n,i}^R &= \max_{i \notin A \subseteq \{i, i+1, \dots, n\}} \left(|A| - \sum_{j=i}^{n-1} \xi_j 1(j \in A, j+1 \in A) \right), \\ X_{n,i}^R &= V_{n,i}^R - W_{n,i}^R. \end{aligned}$$

Recursion as i decreases

$$X_{n,i-1}^R = 1 - \min(X_{n,i}^R, \xi_{i-1}).$$

IDEA OF PROOF: CONSTRUCTION OF A_n^{opt}

$-i - (i + 1) -$	absolute benefit	relative benefit	when used
$- \bullet - - \bullet -$	$V^L + V^R - \xi$	$X^L + X^R - \xi$	if $\xi < \min(X^L, X^R)$
$- \bullet - - \circ -$	$V^L + W^R$	X^L	if $X^R < \min(X^L, \xi)$
$- \circ - - \bullet -$	$W^L + V^R$	X^R	if $X^L < \min(X^R, \xi)$
$- \circ - - \circ -$	$W^L + W^R$	0	never.

IDEA OF PROOF: STATIONARY INFINITE PROCESS

The recursion rule satisfied by X_i^L and $X_{n,i}^R$ does not depend on n . We may define a stationary process $((X_i^L, \xi_i, X_{i+1}^R), -\infty < i < \infty)$ satisfying the same recursion rule:

$$\begin{aligned} X_i^L &= 1 - \min(X_{i-1}^L, \xi_{i-1}) \\ X_i^R &= 1 - \min(X_{i+1}^R, \xi_i). \end{aligned}$$

This type of recursion equation is called a **Recursive Distribution Equation (RDE)**.

\implies For each $i \in \mathbb{Z}$, X_i^L , ξ_i and X_{i+1}^R are independent variables.

\implies It is possible to compute the stationary distribution of this process.

\implies We define an infinite set A^{opt} using the construction table.

IDEA OF PROOF: LOCAL WEAK CONVERGENCE

Lemma 1. *Let U_n be uniform on $\{1, \dots, n\}$. As $n \rightarrow \infty$*

$$((X_{U_n+i}^L, \xi_{U_n+i}, X_{n, U_n+i+1}^R, 1(U_n + i \in A_n^{opt})), -\infty < i < \infty)$$

$$\xrightarrow{d} ((X_i^L, \xi_i, X_{i+1}^R, 1(i \in A^{opt})), -\infty < i < \infty)$$

where the left side is defined arbitrarily for $U_n + i \notin \{1, \dots, n\}$ and where convergence in distribution is with respect to the usual product topology on infinite sequence space.

IDEA OF PROOF: IDENTIFICATION OF THE LIMIT

Recall that $M_n = \max_{A \subseteq \{1,2,\dots,n\}} (|A| - \sum_{i=1}^{n-1} \xi_i 1(i \in A, i+1 \in A))$.

The local weak convergence lemma implies that

$$c(\lambda) = \lim_{n \rightarrow \infty} \mathbb{E} \frac{M_n}{n} = \mathbb{P}(0 \in A^{\text{opt}}) - \mathbb{E} \xi 1(0 \in A^{\text{opt}}, 1 \in A^{\text{opt}}).$$

NEAR MINIMAL SOLUTION

Recall that $f(A) = |A| - \sum_i \xi_i 1(i \in A, i+1 \in A)$ and

$$\varepsilon_n(\delta) = \min_{B \subseteq \{1, \dots, n\}} \left\{ \frac{f(B) - f(A_n^{\text{opt}})}{n} : |B \Delta A_n^{\text{opt}}| \geq \delta n \right\}.$$

Theorem 2. $\varepsilon(\delta) = \lim_n \mathbb{E} \varepsilon_n(\delta)$ exists for all $0 < \delta < 1$, and

$$\limsup_{\delta \downarrow 0} \delta^{-2} \varepsilon(\delta) < \infty,$$

and

$$\liminf_{\delta \downarrow 0} \delta^{-2} \varepsilon(\delta) > 0.$$

\implies Scaling exponent is 2.

IDEA OF PROOF: LAGRANGE MULTIPLIER

We introduce the Lagrange multiplier $\theta > 0$, and

$$\max_{B \subseteq \{1, \dots, n\}} \left(|B| - \sum_{i=-\infty}^{\infty} \xi_i 1(i \in B, i+1 \in B) + \theta |B \Delta A_n^{\text{opt}}| \right),$$

and define $B_n^{\text{opt}}(\theta)$ as the corresponding optimizing set.

We extend our previous analysis to this **new optimization problem**.

We define a new pair of random variables $(Z^L(\theta), Z^R(\theta))$ playing the role of (X^L, X^R) .

$\longrightarrow X^{L,R} = Z^{L,R}(0)$ and $A_n^{\text{opt}} = B_n^{\text{opt}}(0)$.

IDEA OF PROOF: QUINTUPLE PROCESS

We define a stationary quintuple process

$$((Z_i^L(\theta), X_i^L, \xi_i, X_{i+1}^R, Z_{i+1}^R(\theta)), -\infty < i < \infty).$$

→ The infinite sets A^{opt} and $B^{\text{opt}}(\theta)$ are built thanks to the quintuple process according to a construction table, and

$$\{i \in B^{\text{opt}}(\theta)\} \text{ is } \sigma((Z_i^L(\theta), X_i^L, \xi_i, X_{i+1}^R, Z_{i+1}^R(\theta))\text{-measurable.}$$

Let $J_i = 1(i \notin A^{\text{opt}}) - 1(i \in A^{\text{opt}})$, we find

$$Z_i^L = 1 - \min(Z_{i-1}^L, \xi_{i-1})1(Z_{i-1}^L \geq 0) + \theta J_i$$

$$Z_i^R = 1 - \min(Z_{i+1}^L, \xi_i)1(Z_{i+1}^L \geq 0) + \theta J_i$$

→ A new RDE for (Z^L, Z^R) that we cannot solve analytically.

IDEA OF PROOF: NEAR-MINIMAL SOLUTIONS

The proportion of items at which A^{opt} and B^{opt} differ is

$$\begin{aligned}\delta(\theta) &= \mathbb{P}(\{0 \in A^{\text{opt}}\} \triangle \{0 \in B^{\text{opt}}\}) \\ &= \mathbb{P}(\{X^L > \min(X^R, \xi)\} \triangle \{Z^L - \theta J_0 > \min((Z^R - \theta J_1)^+, \xi)\})\end{aligned}$$

and the difference in mean benefit per item between A^{opt} and B^{opt} is

$$\begin{aligned}\bar{\epsilon}(\theta) &= \mathbb{E}[1(0 \in A^{\text{opt}}) - \xi 1(0 \in A^{\text{opt}}, 1 \in A^{\text{opt}})] \\ &\quad - \mathbb{E}[1(0 \in B^{\text{opt}}) - \mathbb{E}\xi 1(0 \in B^{\text{opt}}, 1 \in B^{\text{opt}})] \\ &= \mathbb{E}[\text{complicated expression}].\end{aligned}$$

We have

$$\lim_{n \rightarrow \infty} \mathbb{E}_{\varepsilon_n}(\delta(\theta)) = \bar{\epsilon}(\theta).$$

IDEA OF PROOF: VARIATIONAL ANALYSIS

We write

$$Z^{L,R} - \theta J_0 = X^{L,R} + \theta D^{L,R}(\theta).$$

Define $X = (X^L, \xi, X^R)$ and $D(\theta) = (D^L(\theta), 0, D^R(\theta))$, then

$$\begin{aligned} \delta(\theta) = & \mathbb{P}(X \in \Sigma, X + \theta D(\theta) \notin \Sigma) \\ & + \mathbb{P}(X + \theta D(\theta) \in \Sigma, X \notin \Sigma) \end{aligned}$$

with

$$\Sigma = \left\{ (x^L, t, x^R) \in \mathbb{R}^3 : x^L > \min(t, \max(0, x^R)) \right\}.$$

IDEA OF PROOF: VARIATIONAL ANALYSIS

\Rightarrow Let $F(\theta) = \mathbb{P}(X \in \Sigma, X + \theta D(\theta) \notin \Sigma)$, we prove that

$$\delta(\theta) \sim 2F'(\theta)\theta$$

and there is an integral expression for $F'(\theta)$.

\implies Similarly, we could try to check that $\bar{\epsilon}(\theta) \sim G''(\theta)\theta^2$.

However computation is hard...

We using a totally different probabilistic argument, we prove a weaker

$$C_1\theta^2 \leq \bar{\epsilon}(\theta) \leq C_2\theta^2.$$

IDEA OF PROOF: UPPER AND LOWER BOUND

- UPPER BOUND:

$\limsup \delta^{-2} \varepsilon(\delta) < \infty$: we identify a **local configuration** $\bullet \circ \bullet \circ \bullet \circ \bullet$ on $\{i, \dots, i + K\}$ which can be replaced by $\bullet \bullet \circ \bullet \circ \bullet \bullet$ at a low extra cost $1 - \xi_i - \xi_{i+K-1}$.

- LOWER BOUND:

$\liminf \delta^{-2} \varepsilon(\delta) > 0$: we prove that $C_1 \theta^2 \leq \bar{\varepsilon}(\theta)$ by expressing $\bar{\varepsilon}(\theta)$ as a mean cost increase **over a block** of finite random length. We then lower bound the mean increase cost if B^{opt} and A^{opt} differ over a block by a **combinatorial argument**.

SCALING OF THE LAGRANGE PARAMETER

For the *NK-model / minimal spanning tree of a Poisson point process on \mathbb{R}^d* :

$$\delta(\theta) \asymp \theta \quad \text{and} \quad \bar{\epsilon}(\theta) \asymp \theta^2.$$

→ scaling exponent 2.

Under a *suitable probabilistic model*, for the *traveling salesman problem / minimal pair matching*, simulation suggests [Aldous and Percus (2003)]

$$\delta(\theta) \asymp \theta^{1/2} \quad \text{and} \quad \bar{\epsilon}(\theta) \asymp \theta^{3/2}.$$

→ scaling exponent 3.

There exists a related RDE for these optimization problems [Aldous (2001)] but no mathematical explanation of this phenomenon.

THE KAUFFMAN-LEVIN NK-MODEL

The Kauffman-Levin NK model of **random fitness landscape** has attracted extensive literature in statistical physics.

Let $K \geq 2$ and

$$M_n = \max_{A \subseteq [n]} \sum_{i \in A} -W_i(A_{[i, i+K]}),$$

where $A_{[i, i+K]}$ denotes the set A restricted on the interval $\{i, \dots, i+K\}$ and

$$(W_i(B), i \geq 1, B \subseteq [K+1])$$

are independent $\exp(1)$ random variables.

\Rightarrow This is algorithmically easy via **dynamic programming**.

\Rightarrow By Kingman's Subadditive Theorem there is an a.s. limit $n^{-1}M_n \rightarrow c_K$.

THE KAUFFMAN-LEVIN NK-MODEL: NEAR MINIMAL SOLUTION

Now, fix $0 < \delta < 1$ and

$$M'_n = \max_{A' \subseteq [n]} \sum_{i \in A'} -W_i(A_{[i, i+K]})$$

$$\text{subject to } |A' \triangle A_n^{\text{opt}}| \geq \delta n$$

and then set

$$\varepsilon_n(\delta) = n^{-1}(M_n - M'_n)$$

We expect the existence of the a.s. limit

$$\varepsilon(\delta) = \lim_{n \rightarrow \infty} \varepsilon_n(\delta)$$

and **simulation and heuristics** suggest that, as $\delta \downarrow 0$

$$\varepsilon(\delta) \sim C\delta^2.$$

\Rightarrow Again “Scaling exponent = 2” !