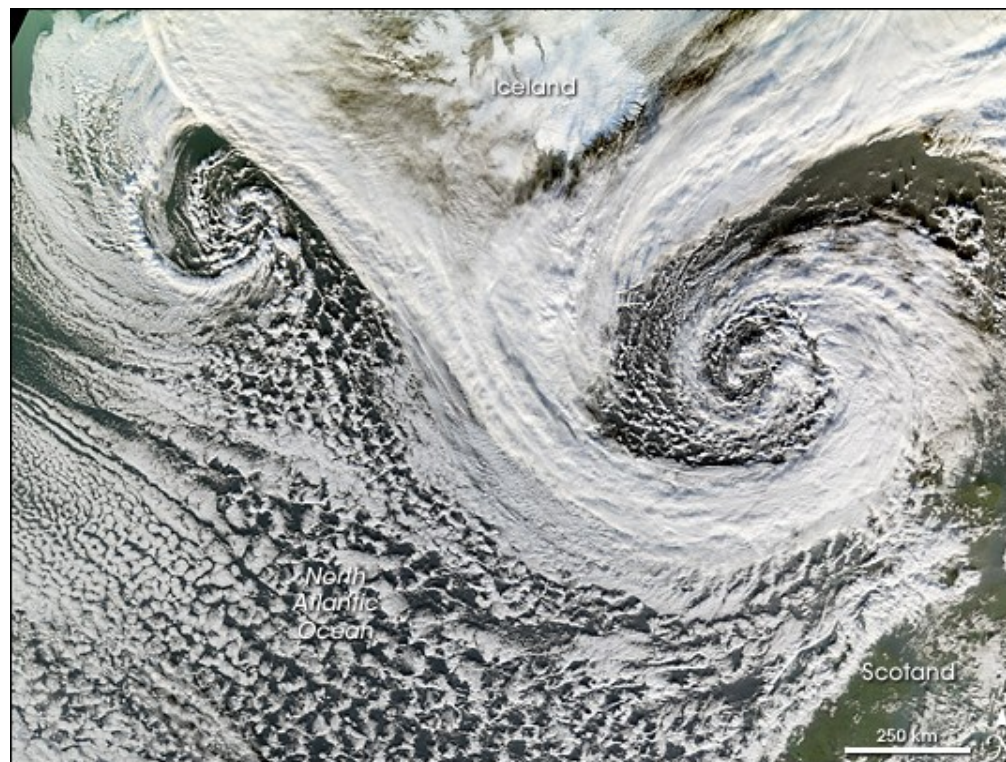


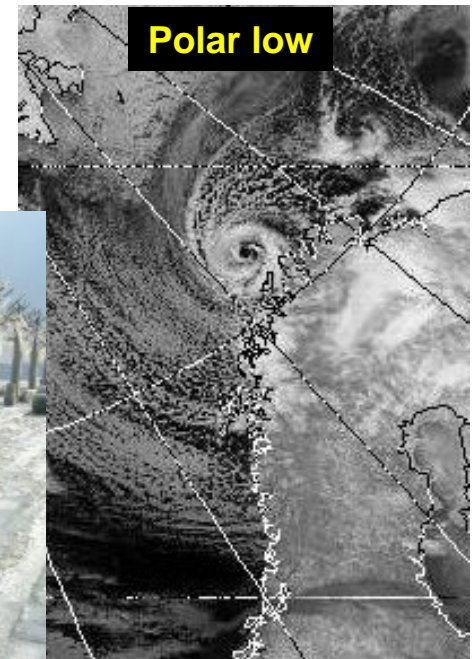
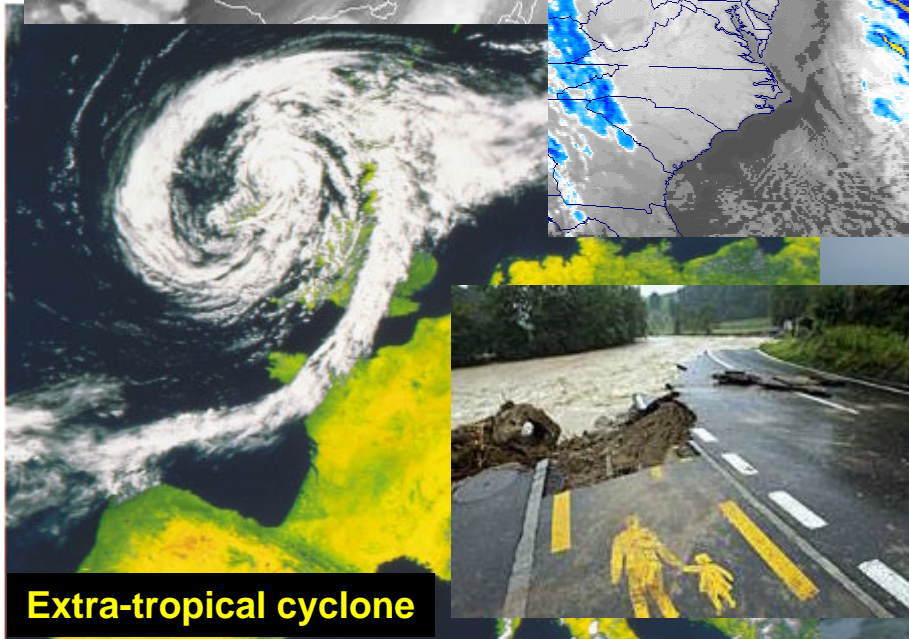
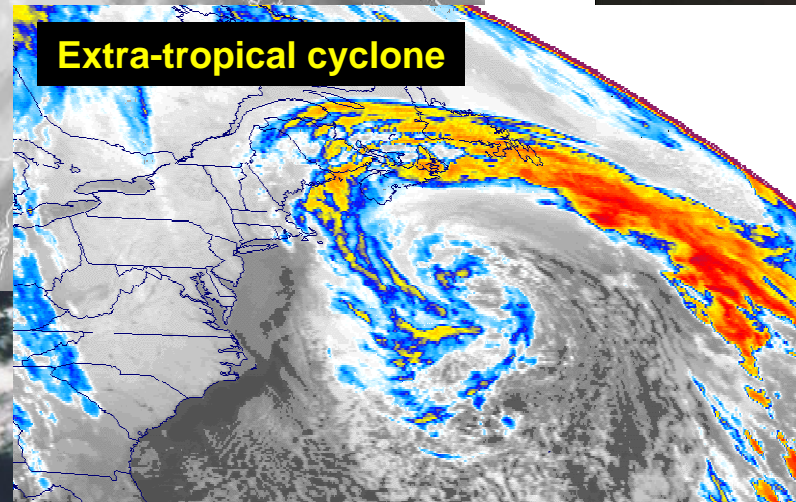
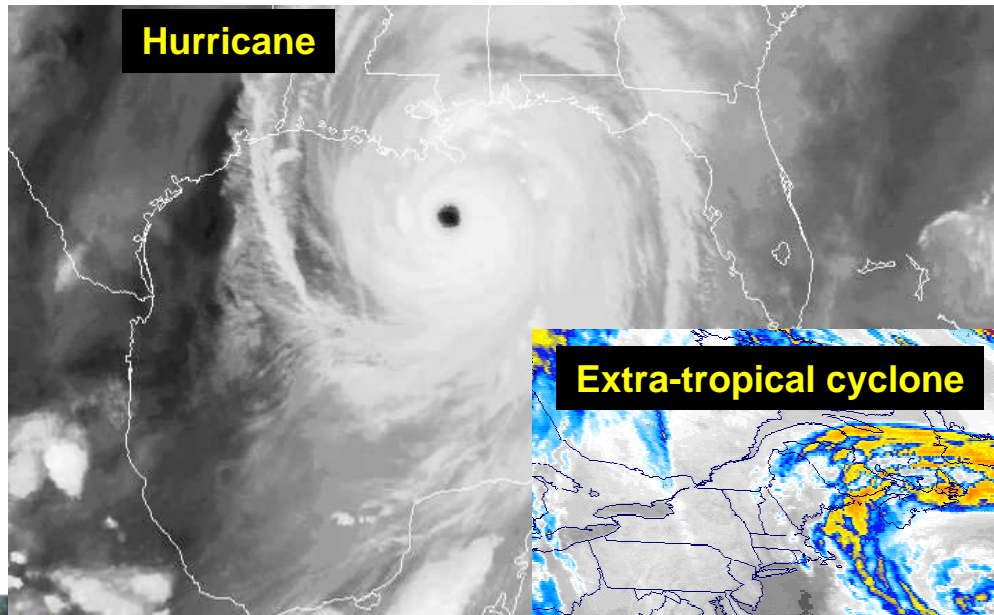
Clustering of Extreme Storms

David Stephenson, Renato Vitolo, Chris Ferro

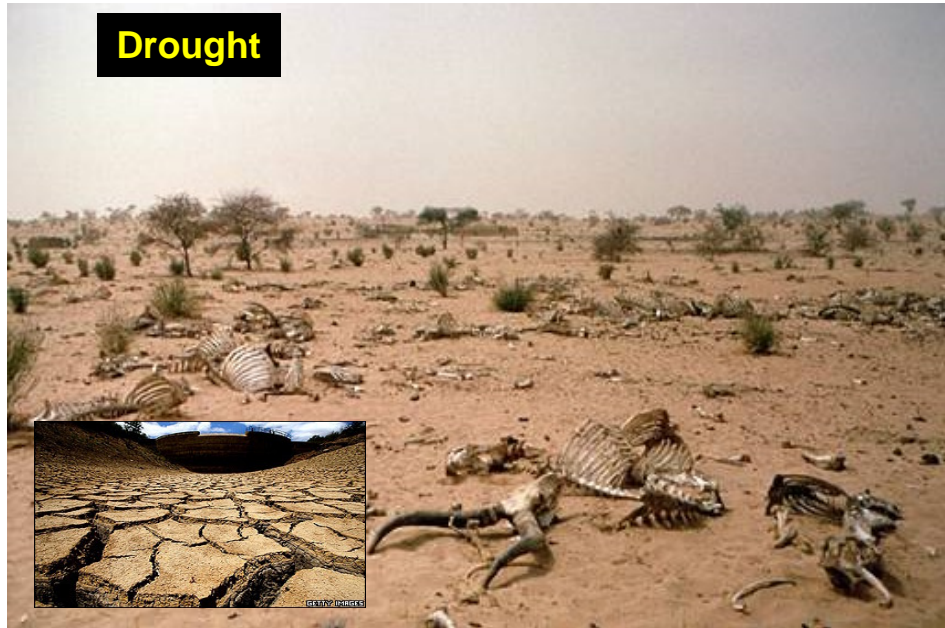


KNMI/EURANDOM workshop "Climate change and EVT", May 11-13 2009, Eindhoven

Examples of wet and windy extremes



Examples of dry and hot extremes



Drought



Wild fire

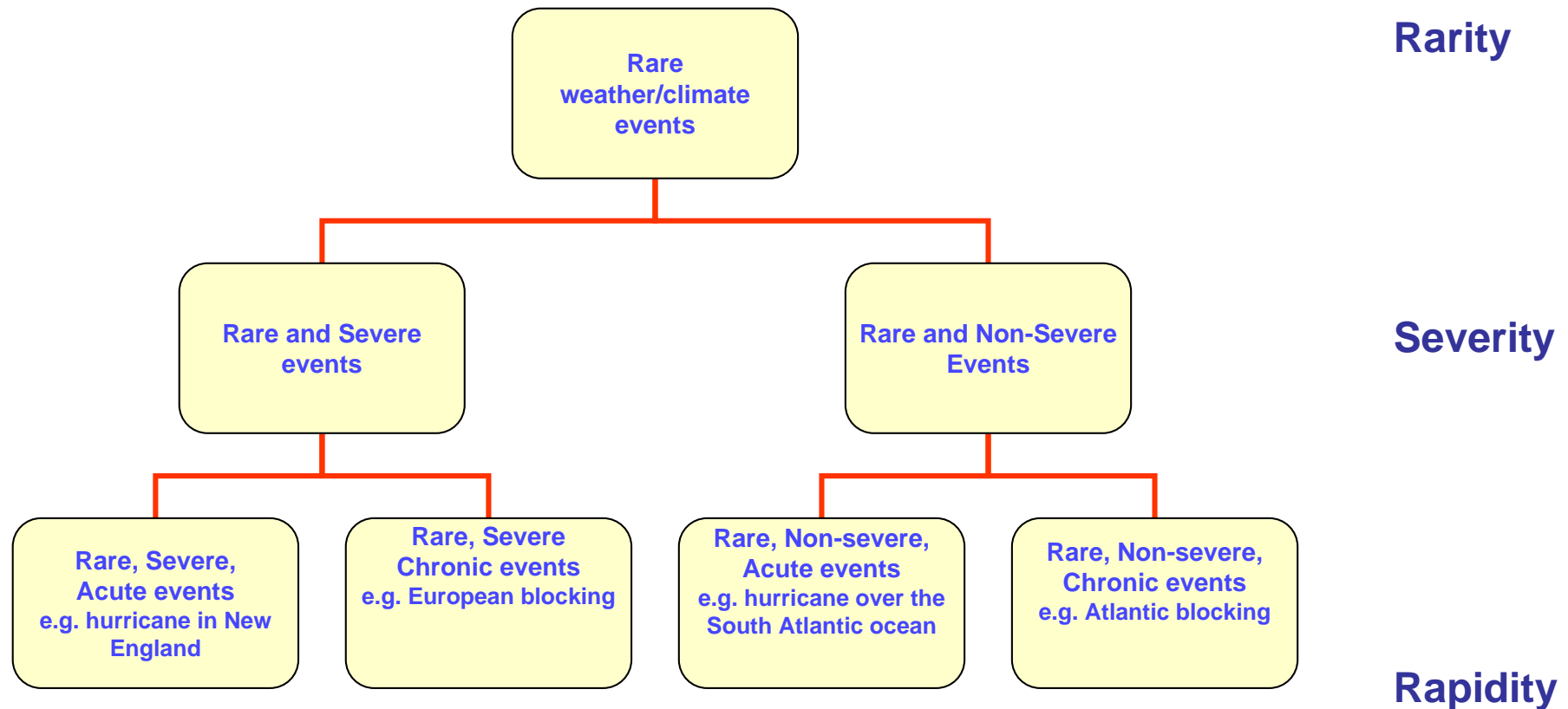


Dust storm



Dust storm

Simple taxonomy



Acute:

Having a rapid onset and following a short but severe course.

Chronic:

Lasting for a long period of time or marked by frequent recurrence

What do we mean by “extreme”?

Large meteorological values

- Maximum value (i.e. a local extremum)
- Exceedance above a high threshold
- Record breaker (threshold=max of past values)

Rare event

(e.g. less than 1 in 100 years – $p=0.01$)

Large losses (*severe or high-impact*)

(e.g. \$200 billion if hurricane hits Miami)

RISK = Expected loss due to the event

= $\text{Pr}(\text{event}) \times \text{loss}(\text{event})$

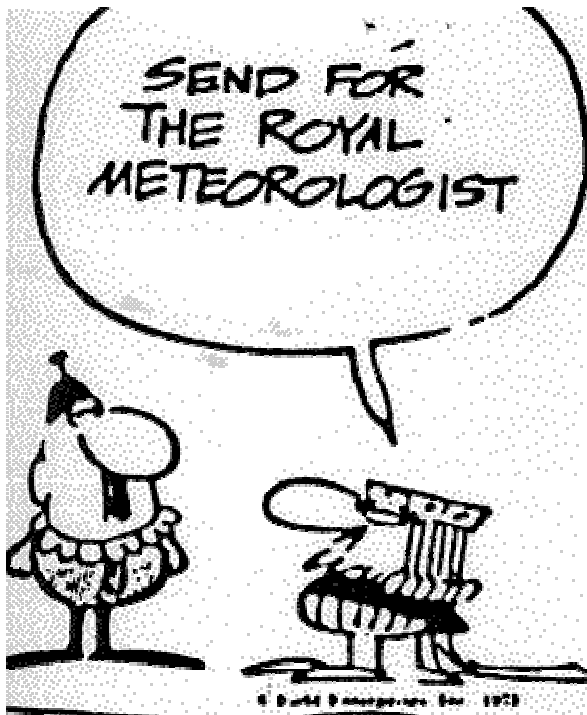
where loss = vulnerability x exposure

Definition of “extreme” event is:

- context-dependent;
- not a binary property (it’s a relative concept!)

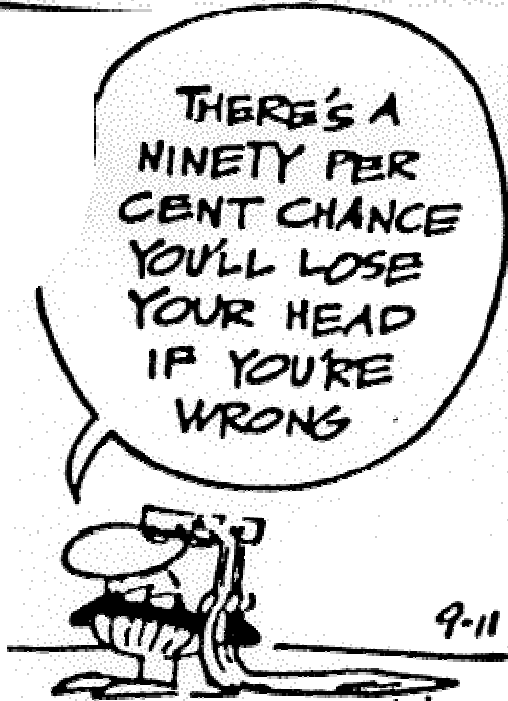
Gare Montparnasse, 22 Oct 1895





© The Wizard of Id
by Brant Parker and Johnny Hart
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1234



Societal relevance of storm clustering

Clustering of European winter storms leads to cumulative insurance losses comparable to those from a catastrophic hurricane:

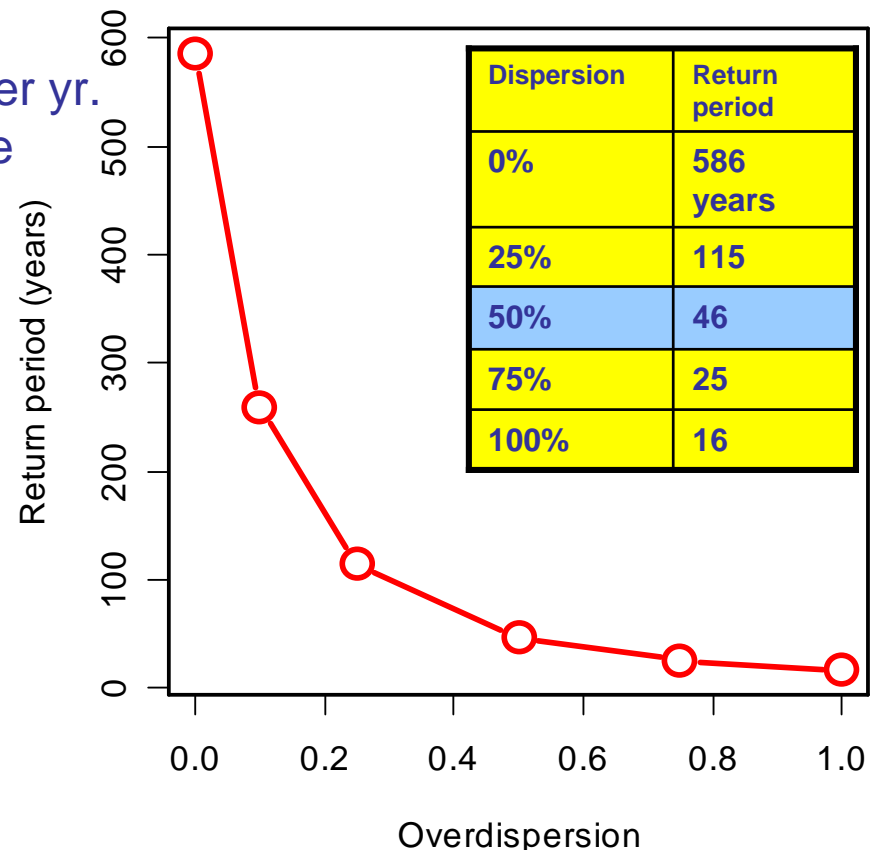
- Dec 1999: 3 consecutive storms (loss \$7.5 bn)
- Dec 1989/Jan 1990: 8 consecutive storms (loss \$10.5 bn)

Example 1:

Reinsurance typically bought for 2 events per yr. Assume to buy cover from 15% exceedance probability level:

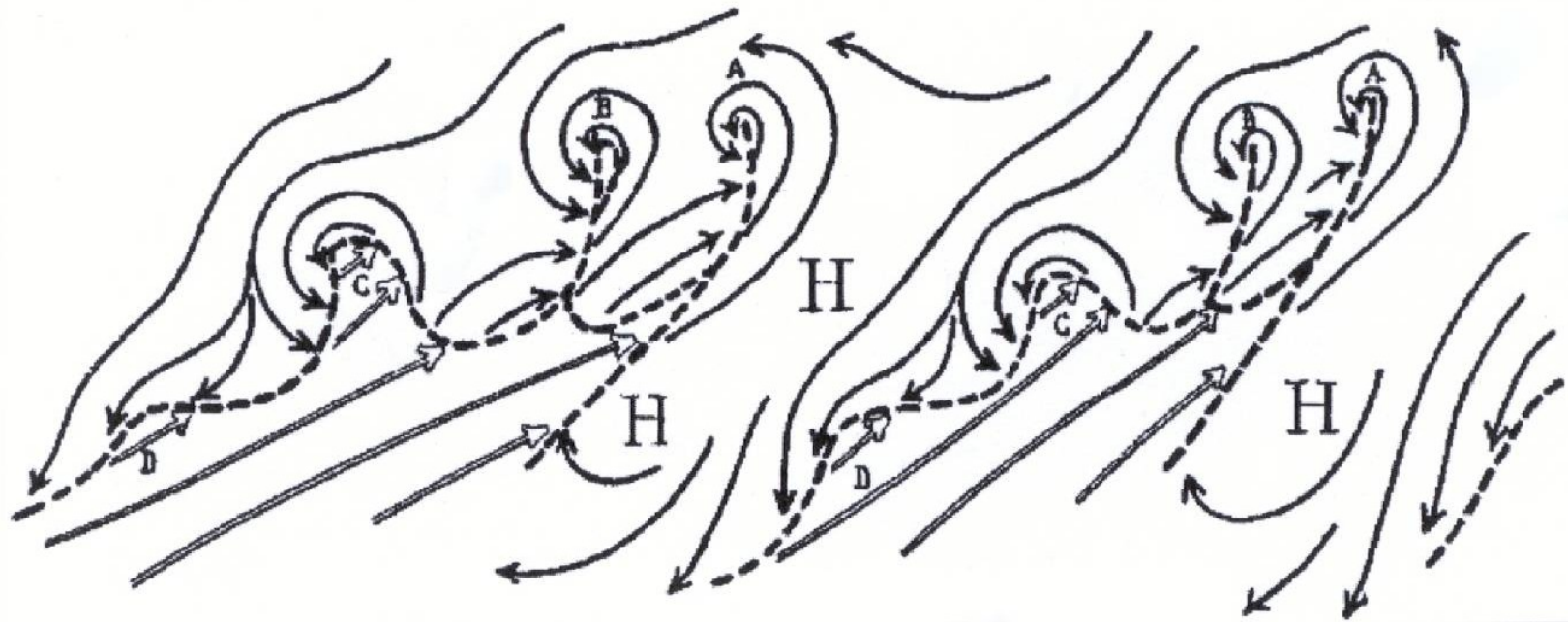
Dispersion =	0	0.5
Prob[2 events/yr]	0.011	0.021
Prob[>2 events/yr]	0.001	0.008

Example 2: Return periods for >15 storms/month estimated using negative binomial with a mean rate of 5.7 cyclones/month



Some key scientific questions

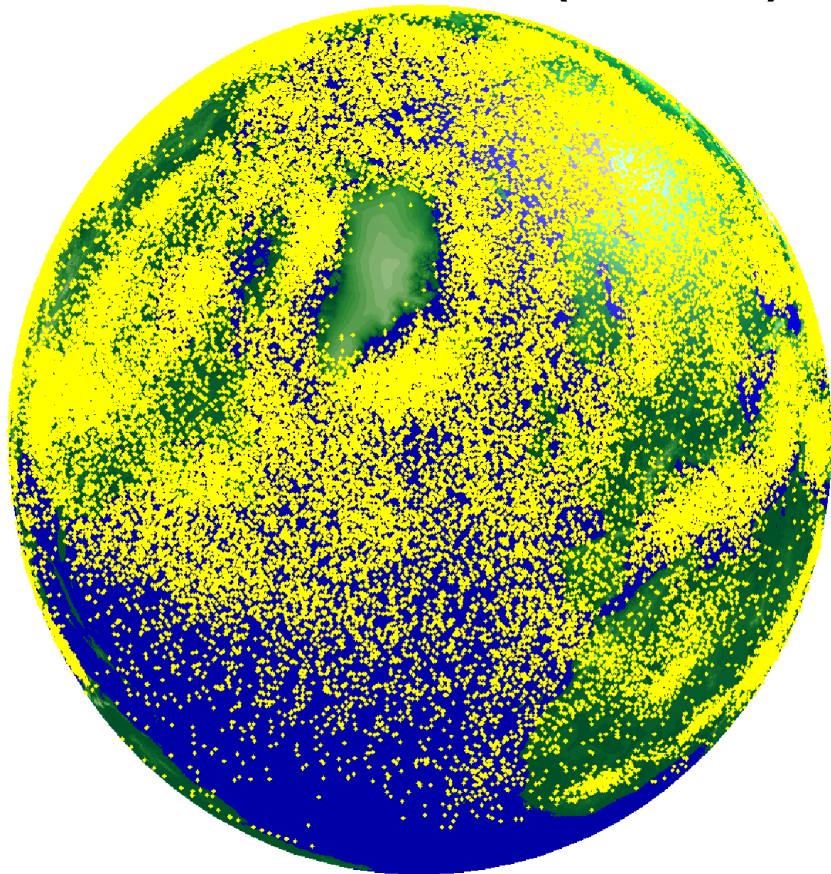
- How much do extratropical wintertime storms cluster?
- How does clustering depend on storm intensity?
- Can large-scale flow be used to explain the clustering?
- What are the implications of this for estimating return periods?



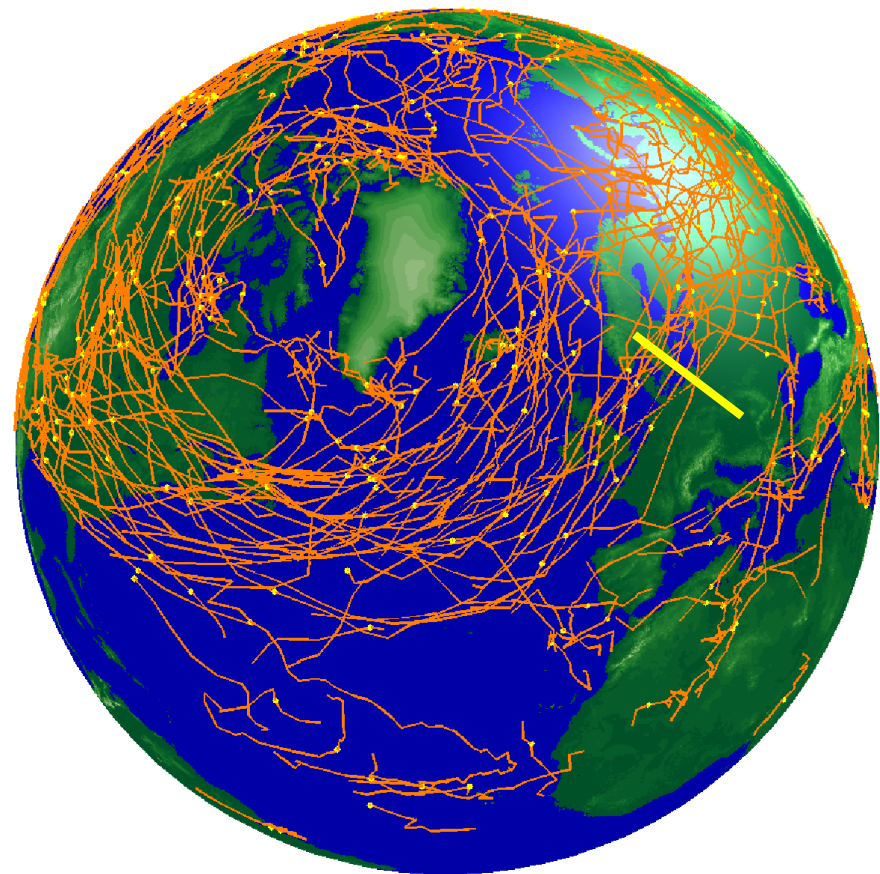
Storm feature tracking 1948-2003

- Eastward cyclone tracks identified objectively using TRACK software
- Extended winters (1 Oct-31 Mar)
- 6 hourly NCAR/NCEP reanalyses from 1948-2003

355,460 VOR zeniths (maxima)

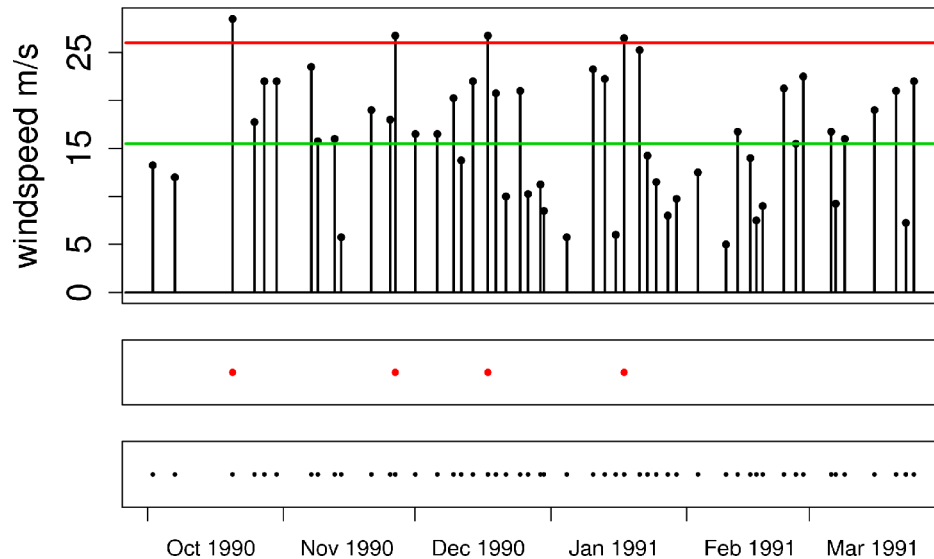


The Storms of Dec 1989-Feb 1990

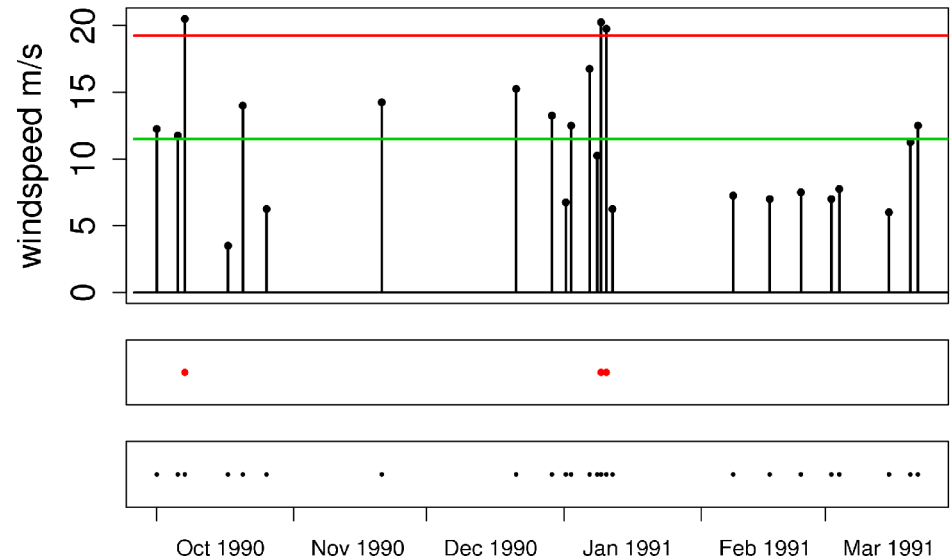


Clustering of storms

Transits $\pm 10^\circ$ of Nova Scotia (45°N, 60°W)



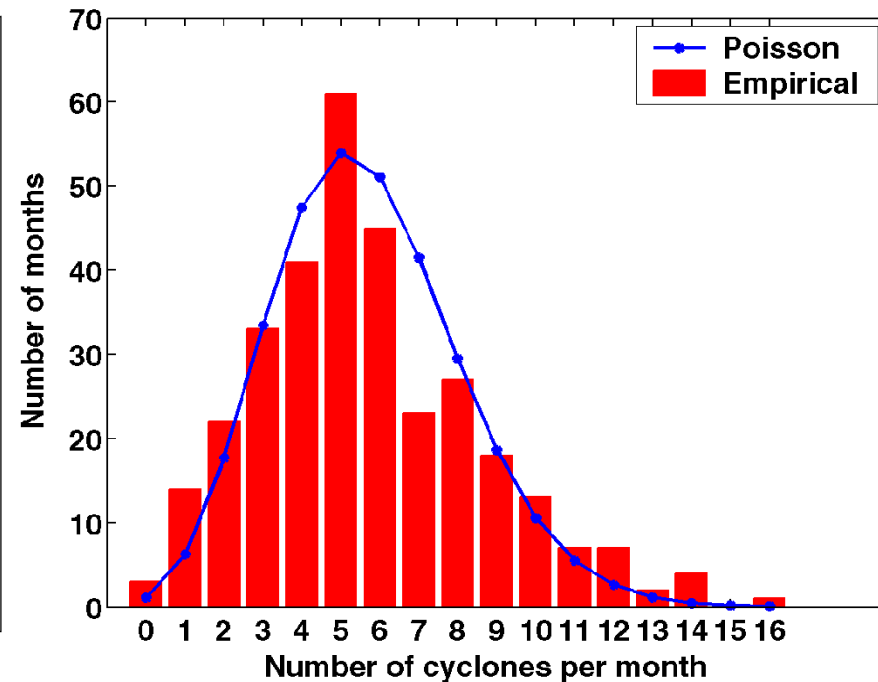
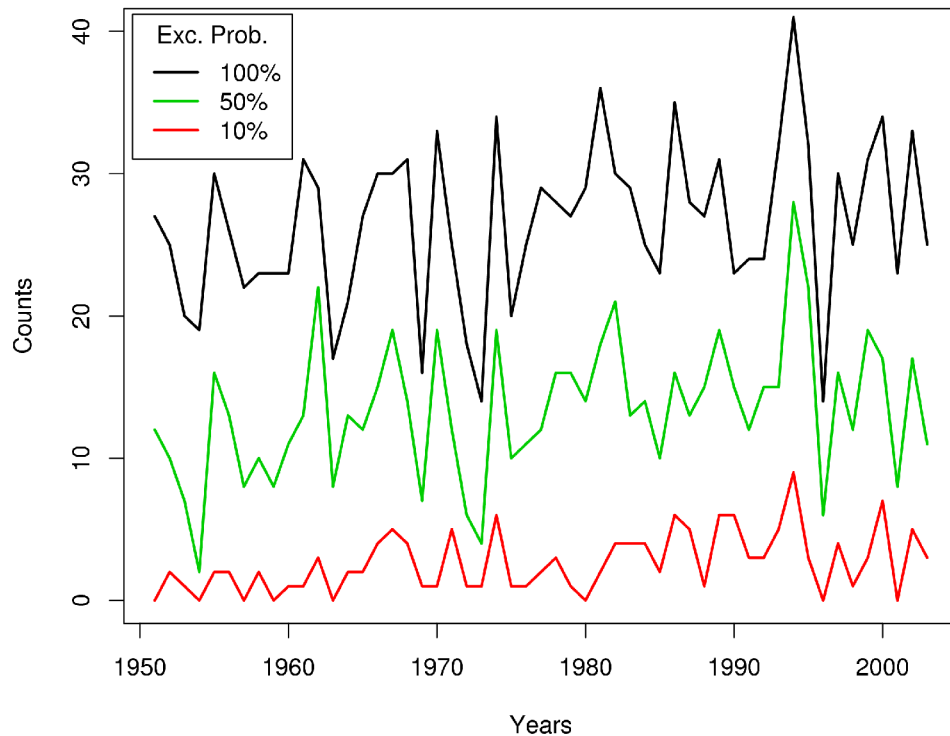
Transits $\pm 10^\circ$ of Berlin (52°N, 12.5°E)



→ Clustered over Europe but not over western Atlantic

Dispersion of storm counts

Oct-Mar counts near Berlin



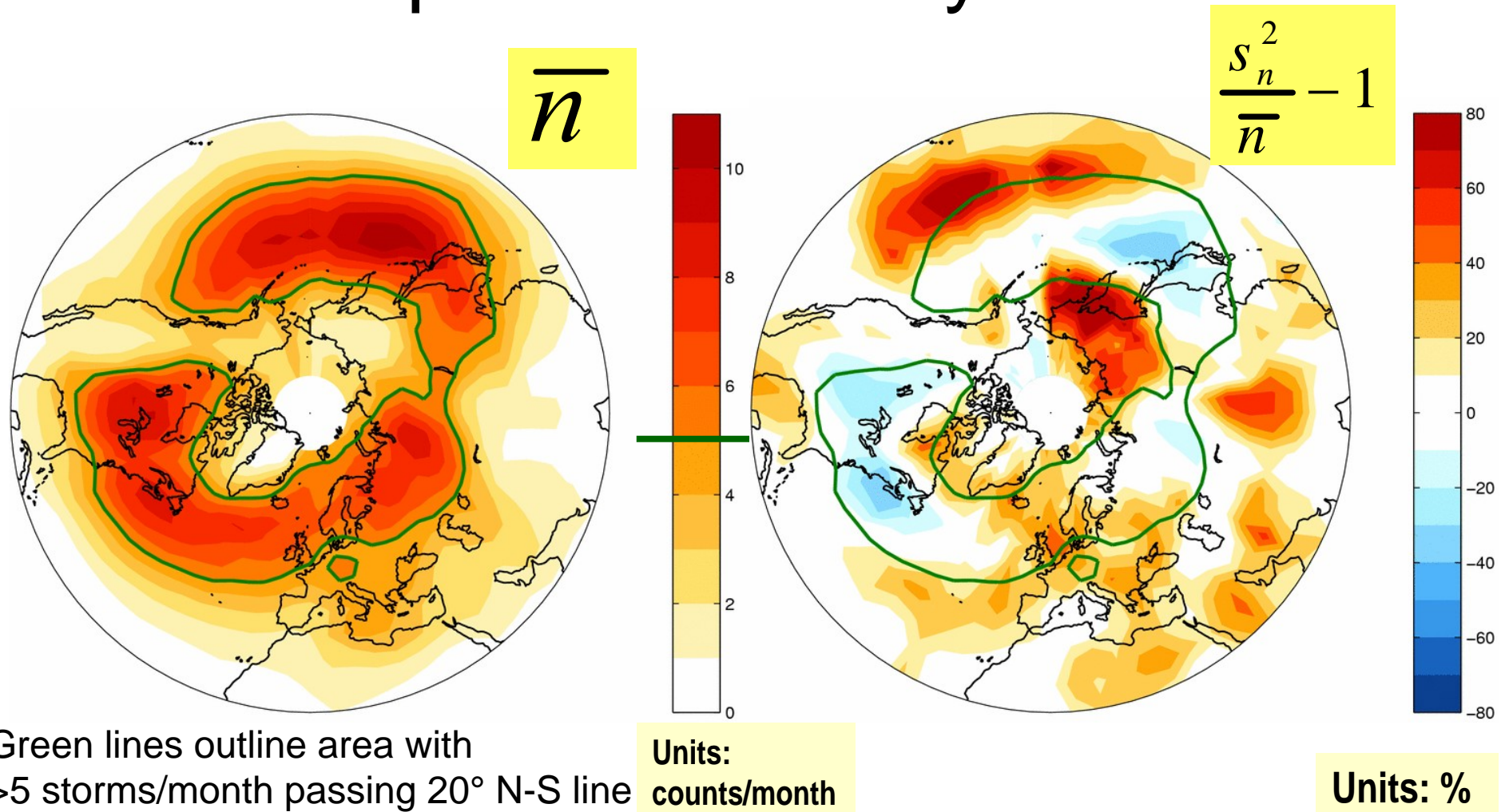
Dispersion statistic:

$$\psi = \frac{\text{Var}(n)}{\text{mean}(n)} - 1$$

=0 when $\text{Var}(n) = \text{Mean}(n)$ e.g. Poisson distributed counts

>0 when $\text{Var}(n) > \text{Mean}(n)$

Mean and dispersion of monthly counts



Green lines outline area with >5 storms/month passing 20° N-S line

→ Regions of overdispersion (reds) and underdispersion (blues)

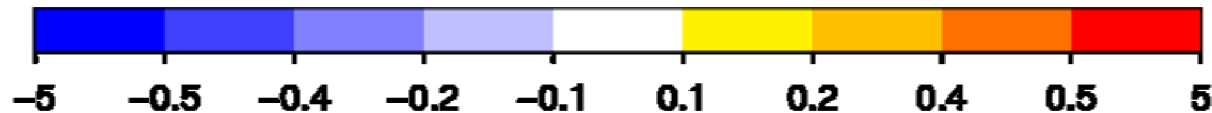
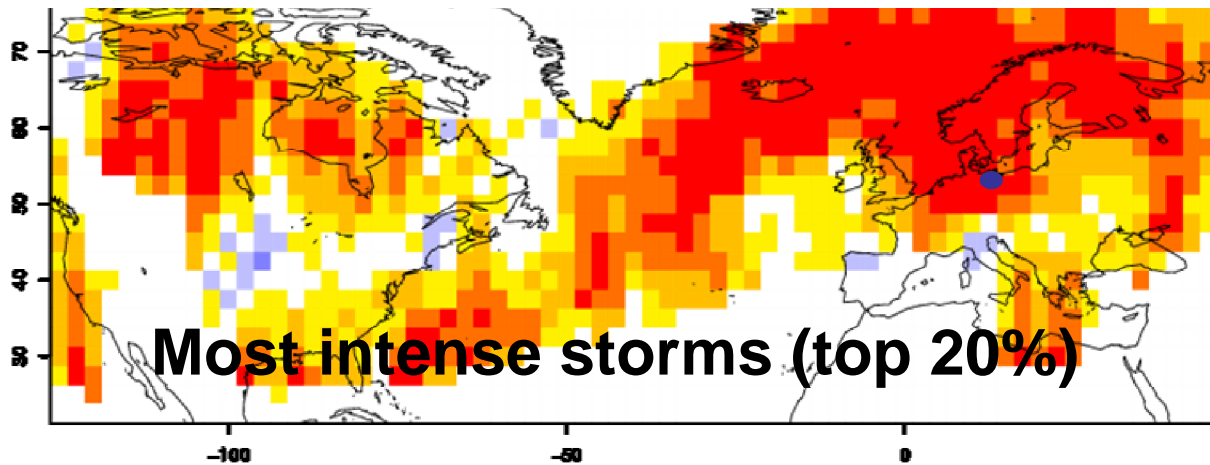
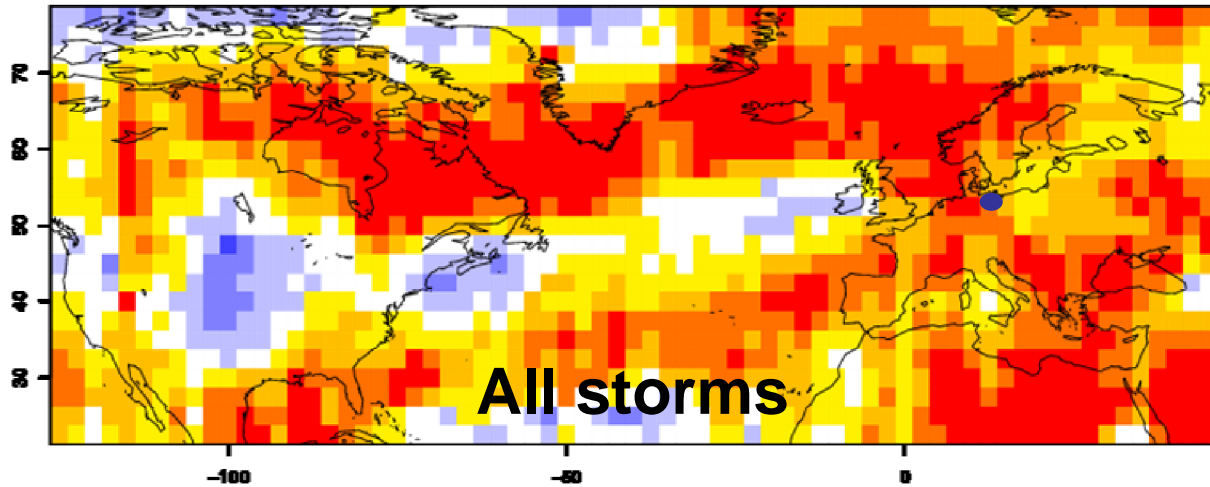
Mailier, P.J., Stephenson, D.B., Ferro, C.A.T. and Hodges, K.I. (2006):

Serial clustering of extratropical cyclones, Monthly Weather Review, 134, pp 2224-2240

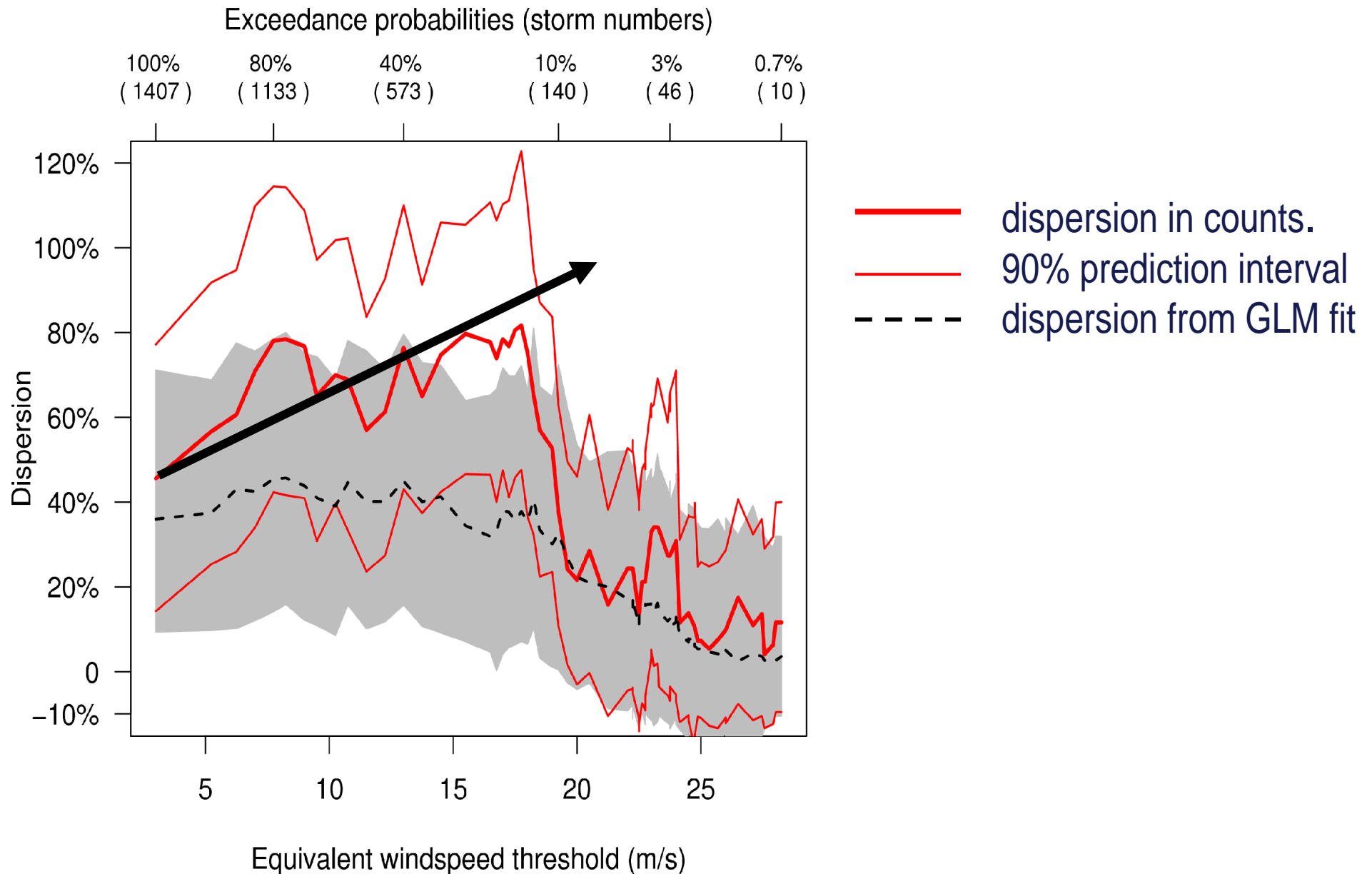
Dispersion increases for intense storms

Dispersion of
3-month counts:

$$\frac{s_n^2}{\bar{n}} - 1$$



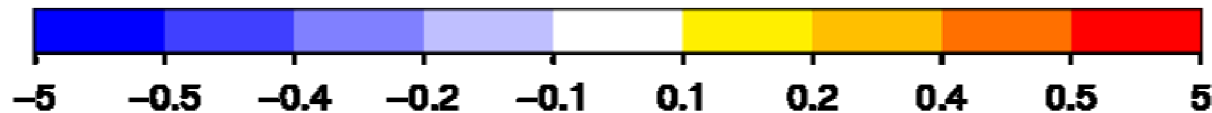
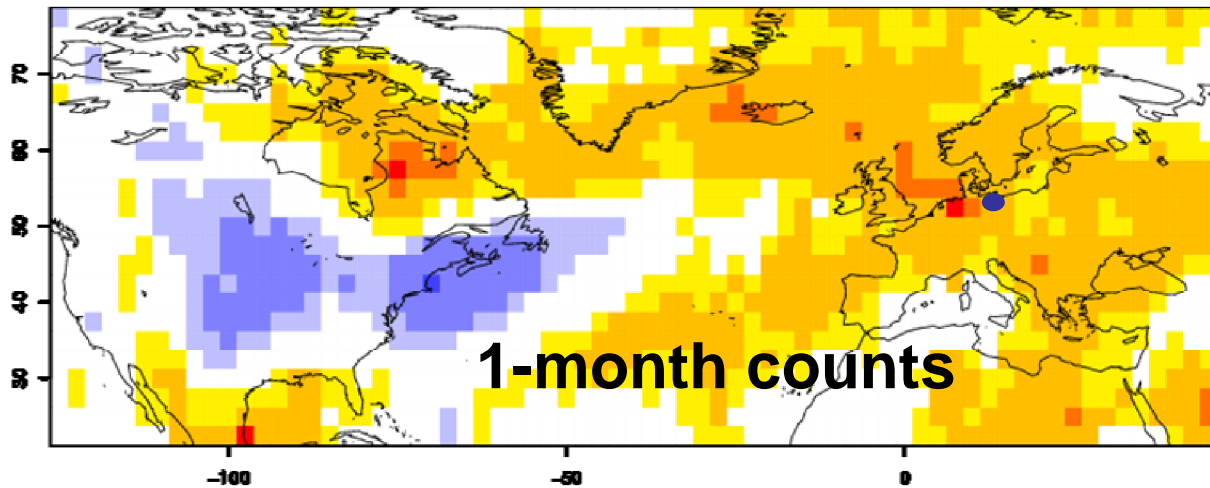
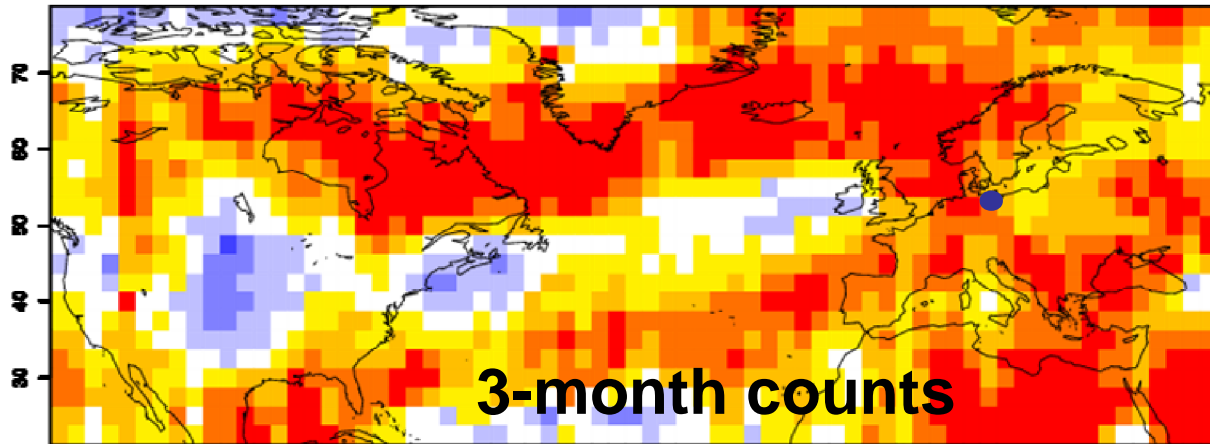
Dispersion versus intensity for Berlin



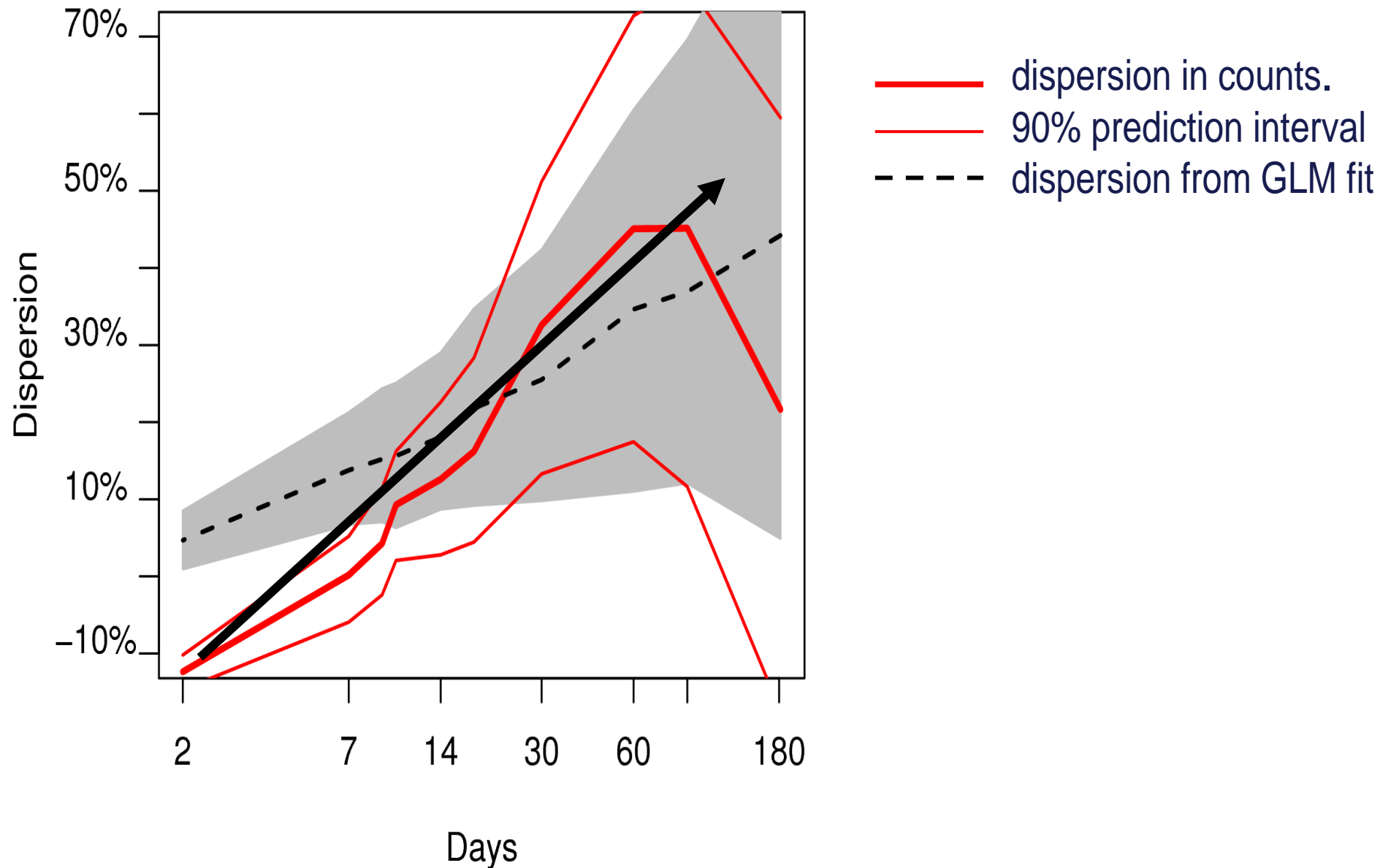
Dispersion increases for longer periods

Dispersion of counts for all storms:

$$\frac{s_n^2}{\bar{n}} - 1$$



Dispersion versus aggregation period for Berlin



Flow-dependent clustering (e.g. UK buses)



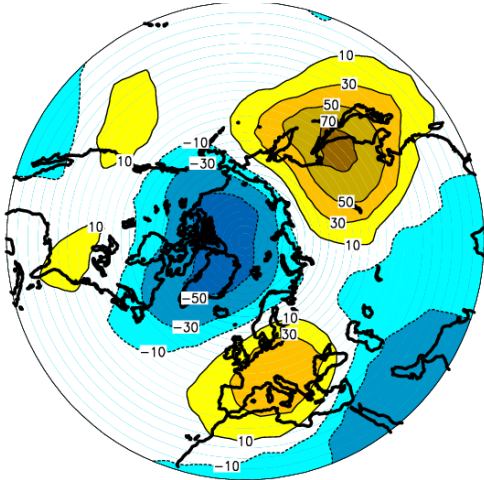
Is this because bus drivers really love each other?

Don't think so! More to do with rate of arrival depending on time varying background traffic flow.

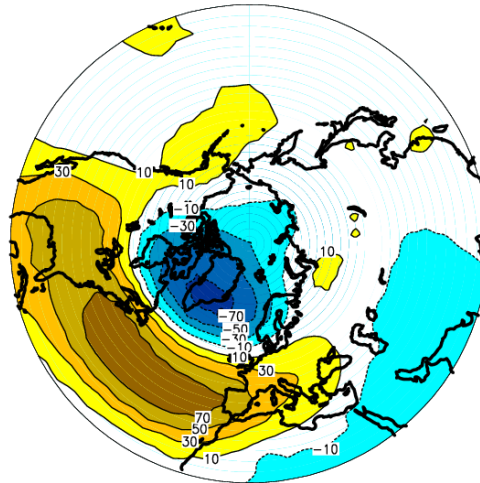
Teleconnection patterns

Leading rotated
EOFs of 700mb
geopotential height

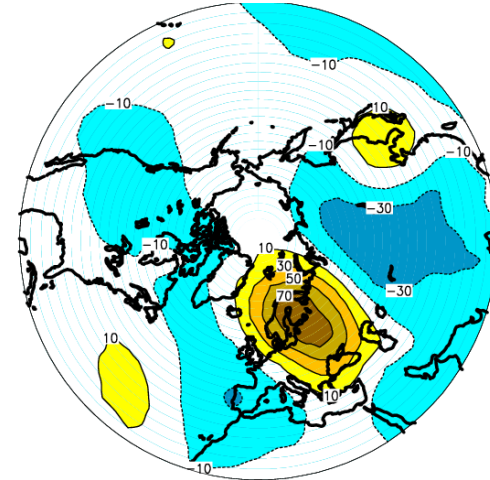
Polar-Eurasian



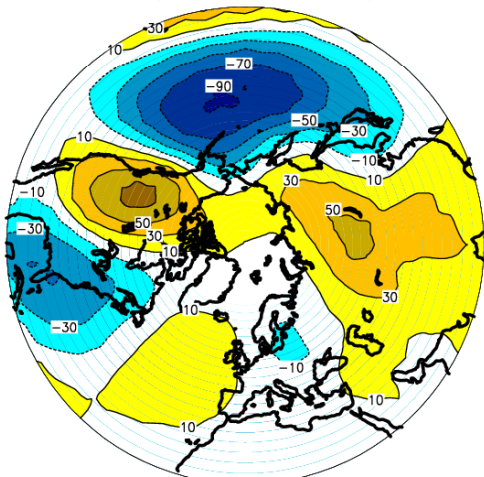
NAO



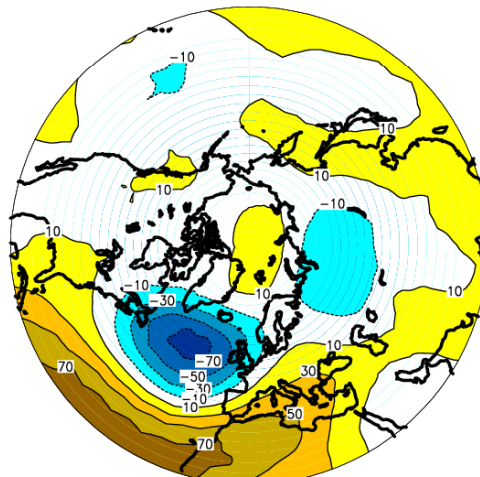
Scandinavian



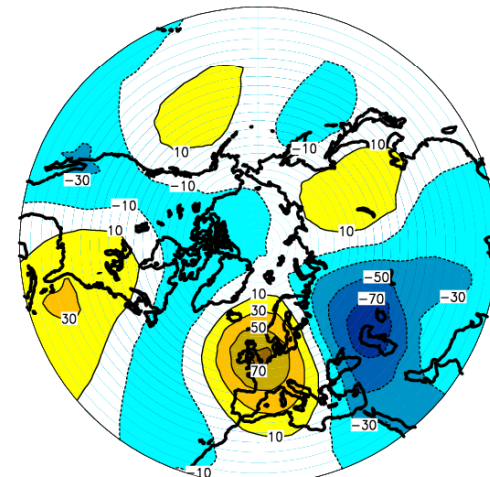
PNA



East Atlantic

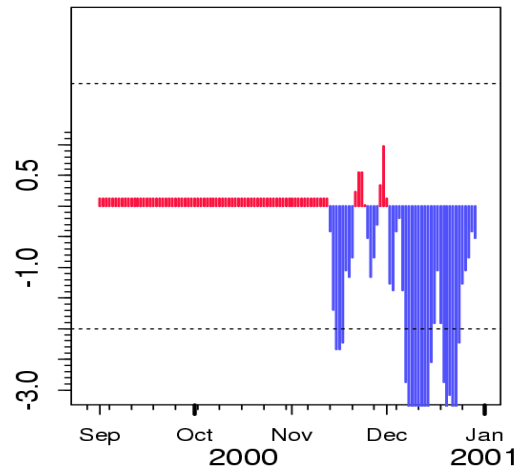


E. Atl/W. Russian

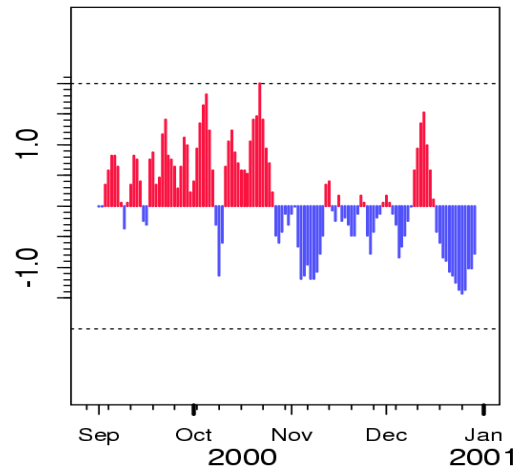


Daily teleconnection indices x_k

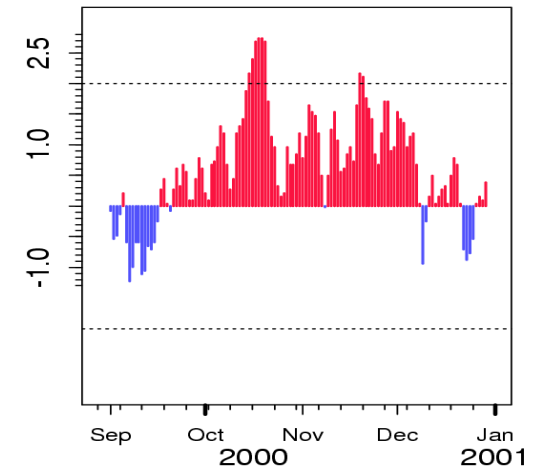
PEU index



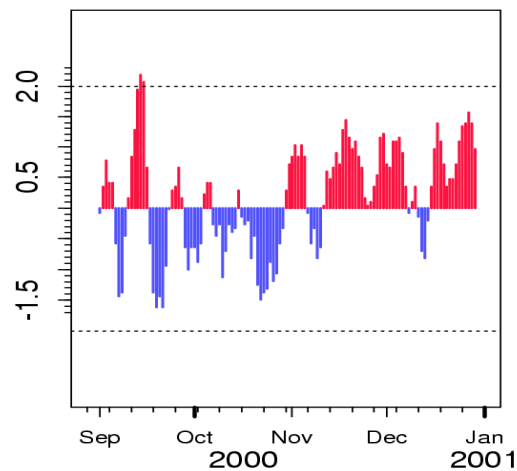
NAO index



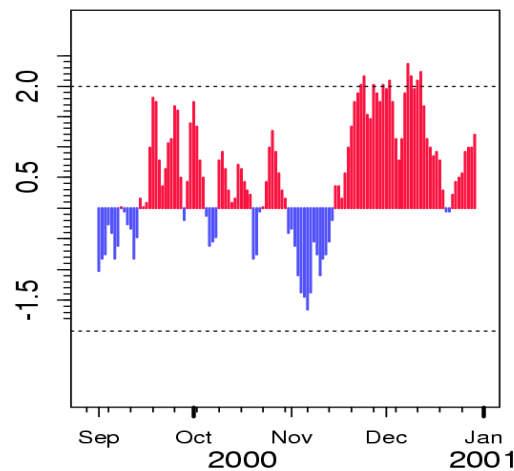
Scandinavian index



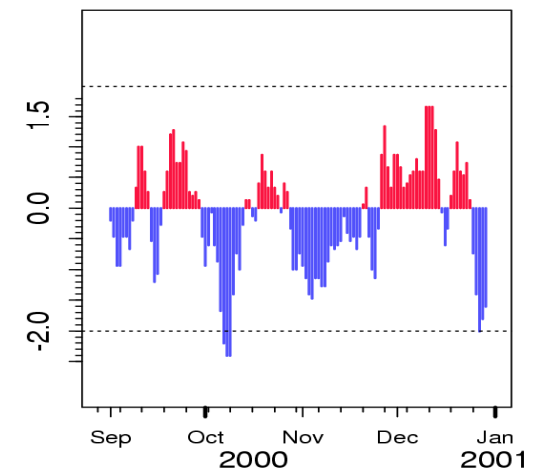
PNA index



East Atlantic index



East Atlantic/West Russia index

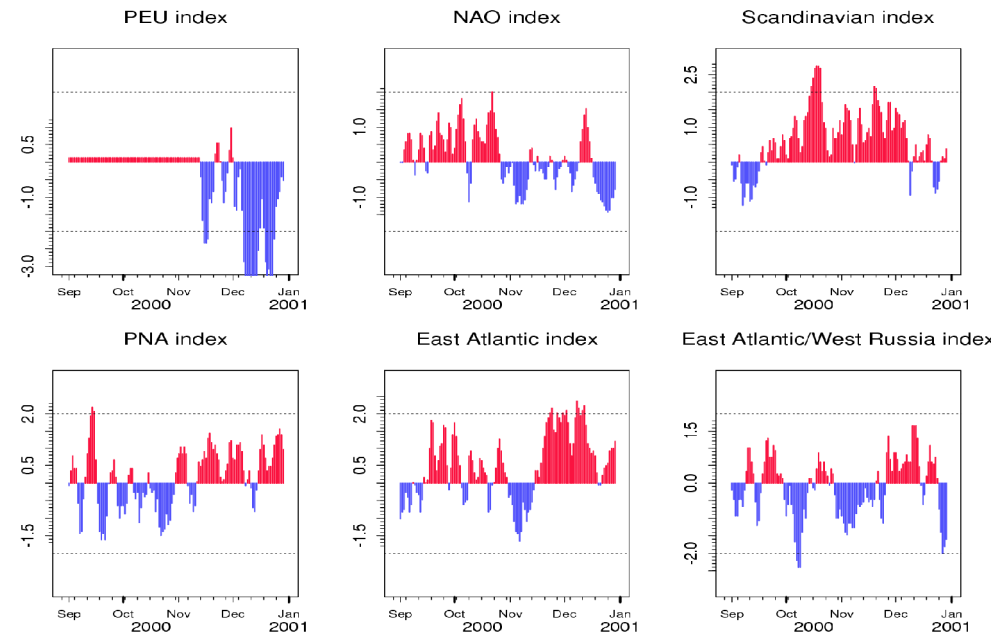


Regression of counts on teleconnections: rate depends on large-scale flow

Poisson regression

$$n/x \sim \text{Poisson}(\mu)$$

$$\log \mu = \beta_0 + \sum_{k=1}^K \beta_k x_k$$



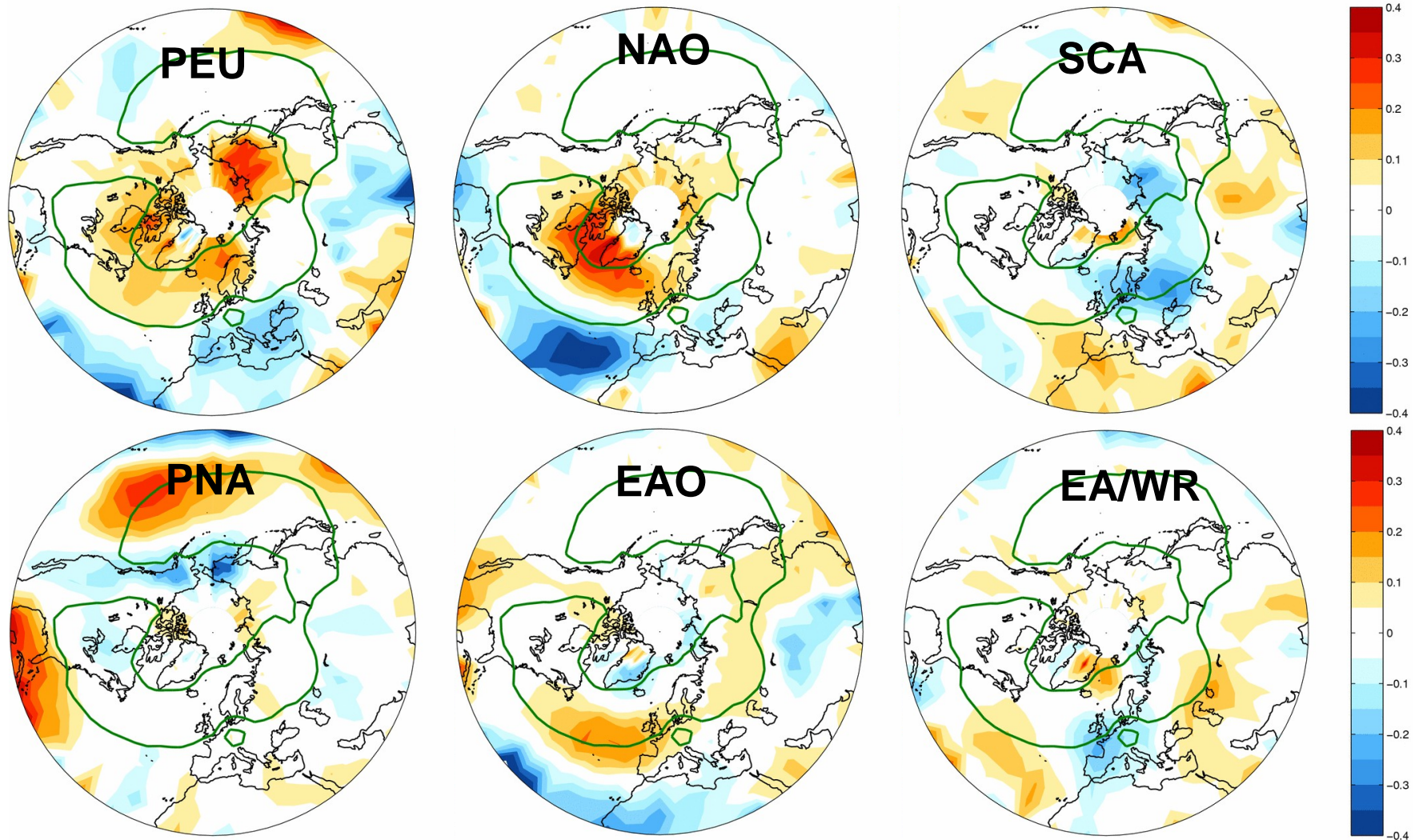
n = number of storms, e.g. monthly counts of windstorms.

μ = time-varying, flow dependent rate.

$x_k(t)$ = large-scale teleconnection indices (covariates).

GLM maximum likelihood estimation of β_0, β_i

Estimated β regression parameters

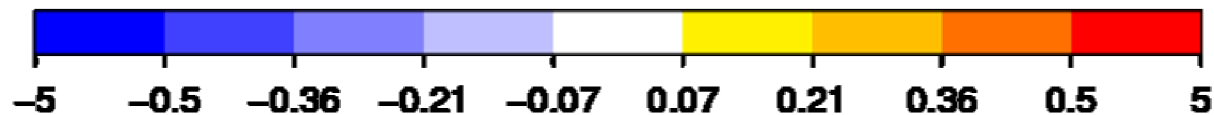
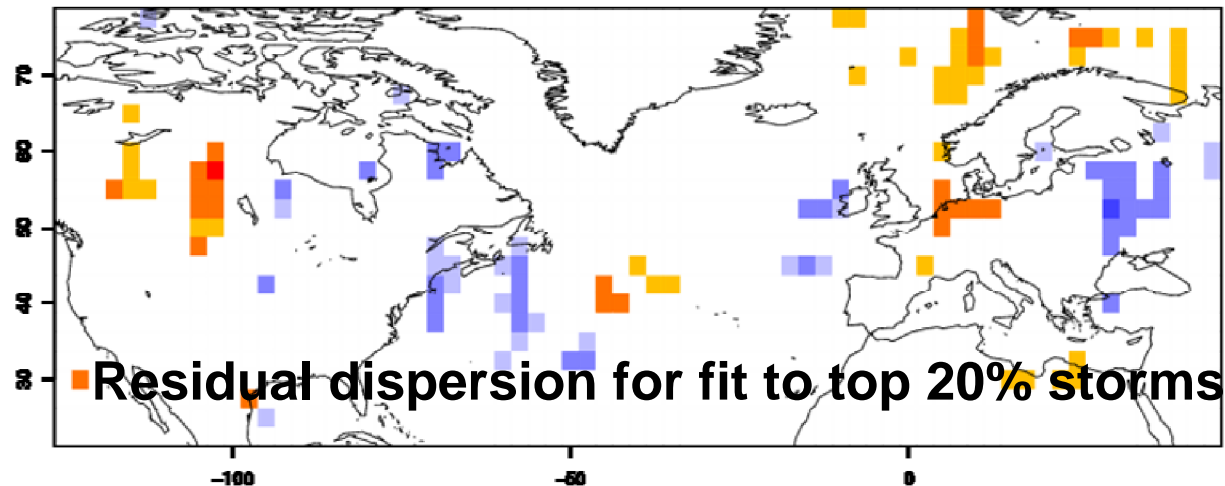
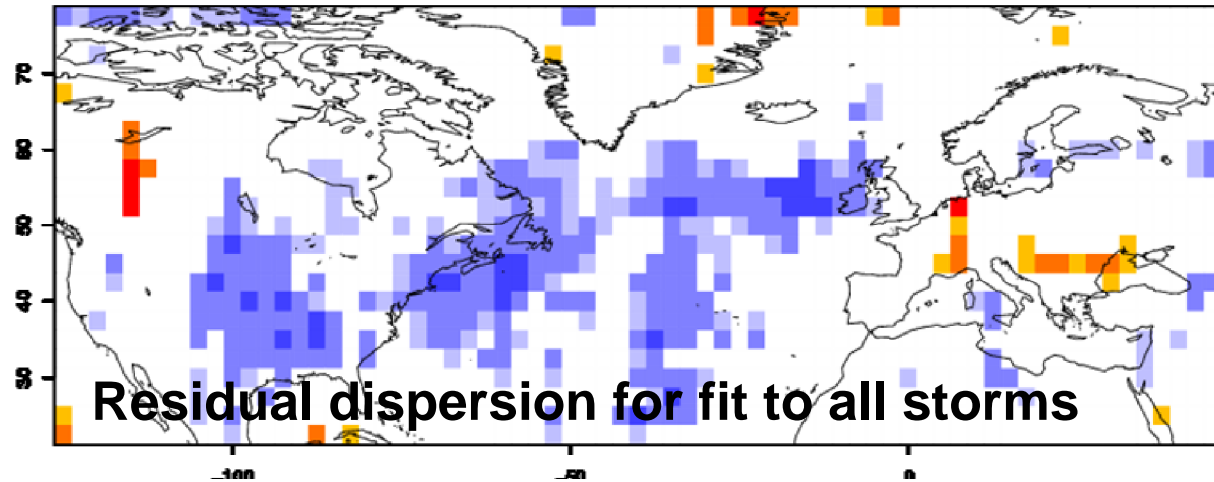


→ all teleconnection patterns are important for European storms

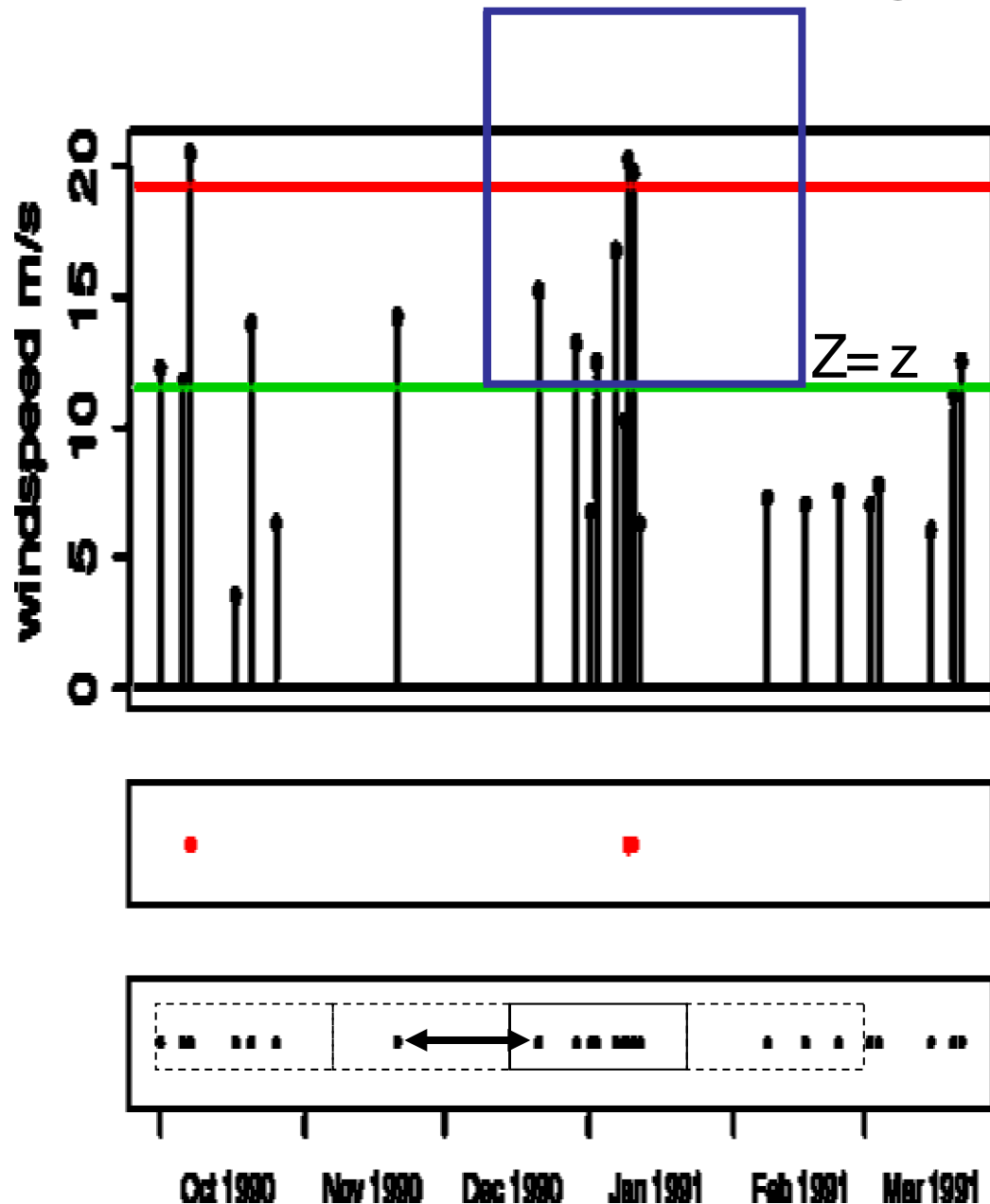
See Seierstad et al. (2007)

Teleconnections account for overdispersion

Residual dispersion
of fits to 3-month
counts



Stochastic modelling of the process



Consider the points $Z_i > z$
 $(i/(n+1), Z_i = (Y_i - b_n)/a_n)$
 $i = 1, 2, \dots, n$

In the limit $n \rightarrow \infty$

$N \sim \text{Poisson}(\Lambda)$

$$\Pr(N = n) = \frac{\Lambda^n e^{-\Lambda}}{n!}$$

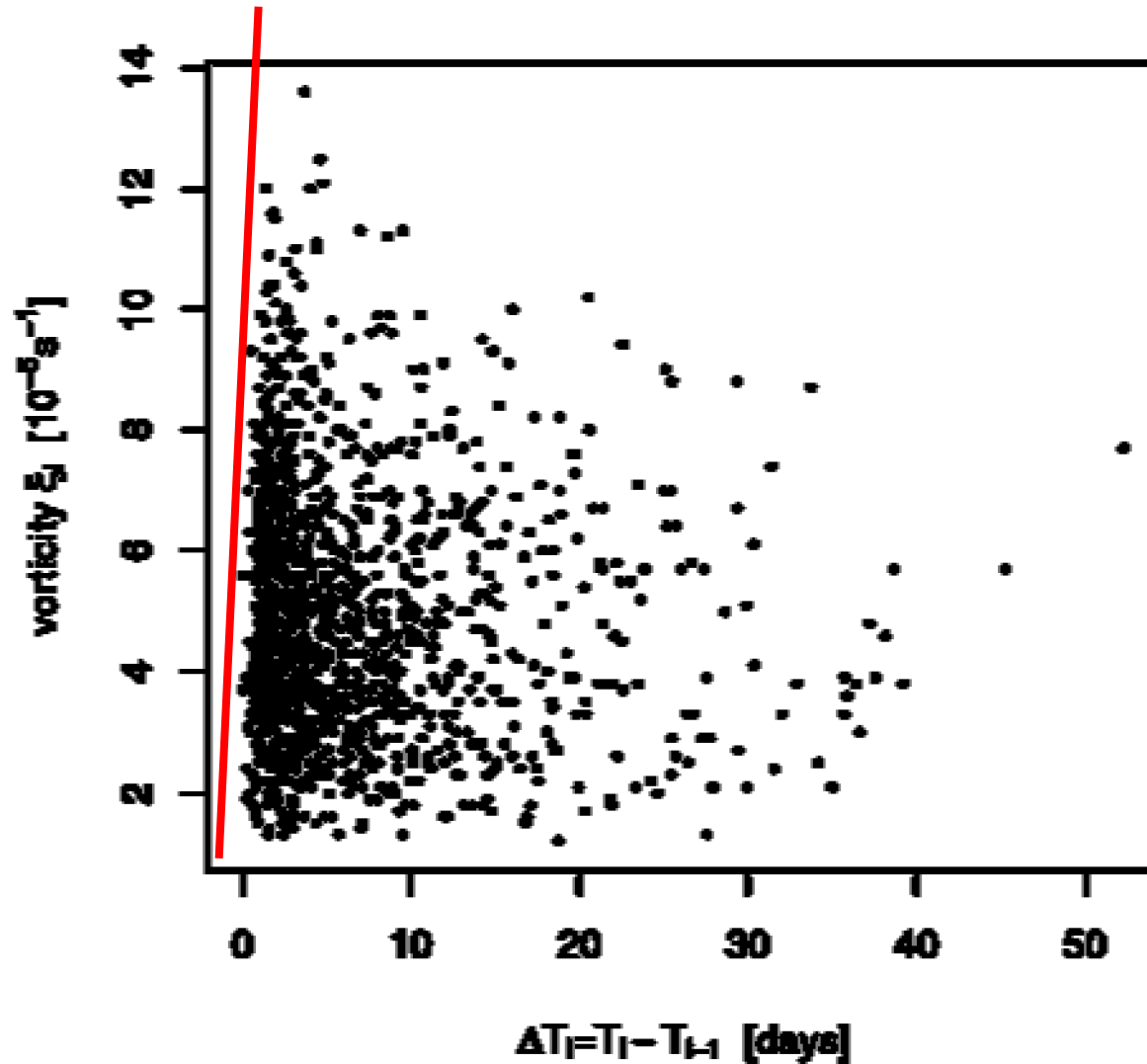
$$\Lambda = (t_2 - t_1) \left\{ 1 + \xi \left(\frac{z - \mu}{\sigma} \right) \right\}^{-1/\xi}$$

$$\rightarrow (t_2 - t_1) \exp \left\{ - \left(\frac{z - \mu}{\sigma} \right) \right\}$$

when $\xi \rightarrow 0$

Distribution of marks & inter-arrival times

Vorticity vs. Interarrival time

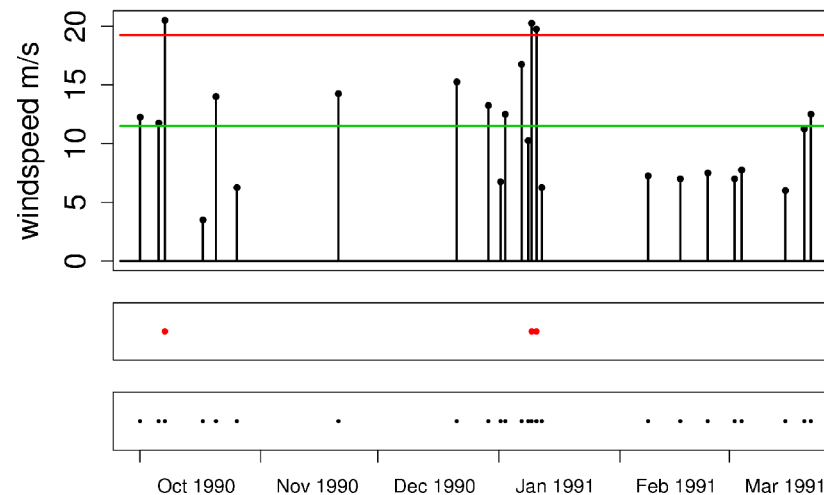


→ More time needed before more intense storms

Modelling times & intensities of extreme storms

Large-scale flow patterns are known to influence the growth rate of storms AND the path of the storms.

Hence, we try to capture both these processes by modelling extreme storms as a **COMPOUND POISSON PROCESS** - a marked point process where the arrival times are independent of the marks (the vorticities).

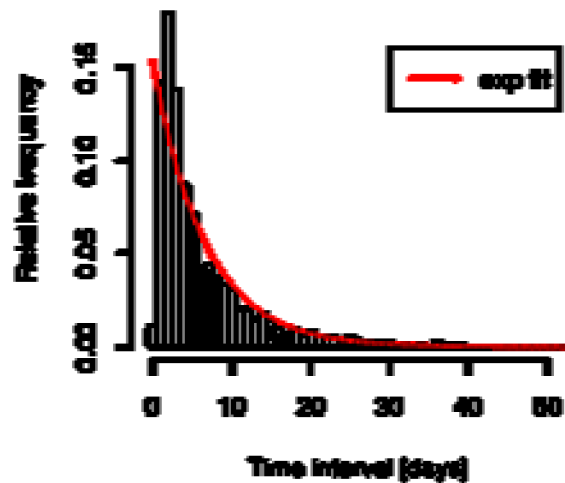


Start by trying a Poisson process with time-varying rate for the arrival times AND a GPD fit for the marks.

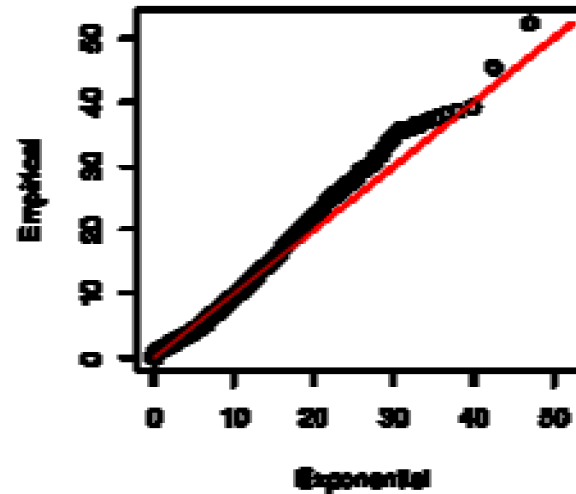
The arrival rate and GPD parameters are allowed to depend on the NAO, EAP, EWP and SCP teleconnection indices (covariates).

Distribution of inter-arrival times

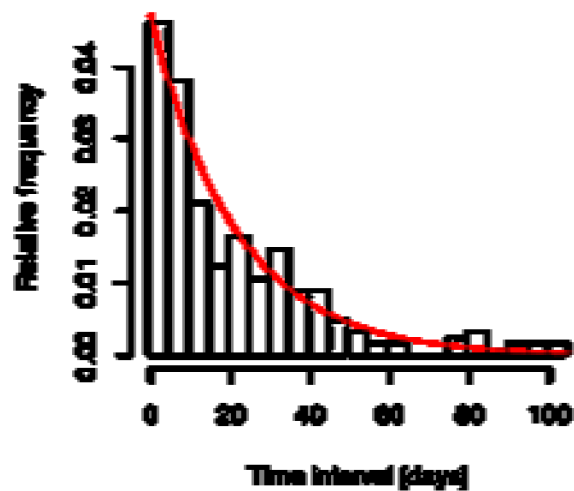
Interarrival times - all storms



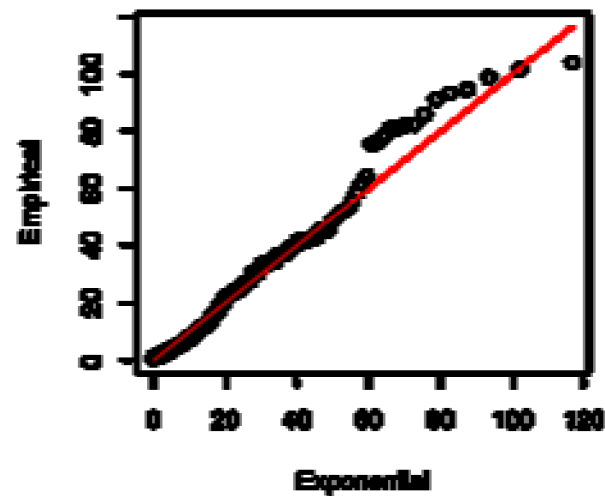
QQ plot, all storms



Interarrival times - top 20% storms



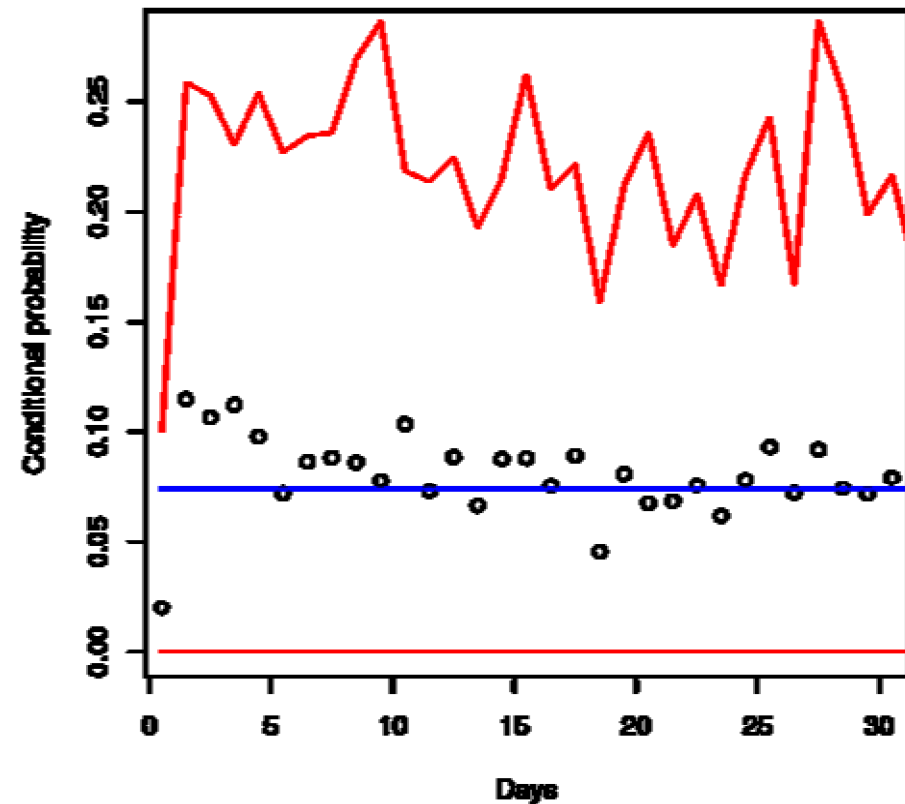
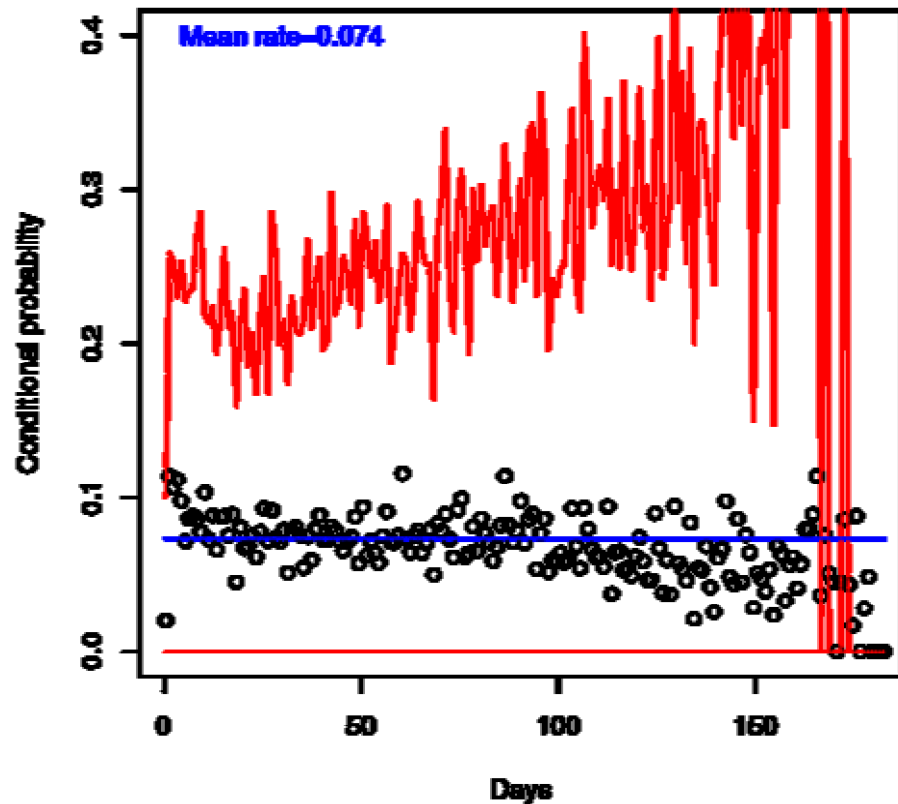
QQ plot, top 20% storms



→ Interarrival times are more dispersed than exponential

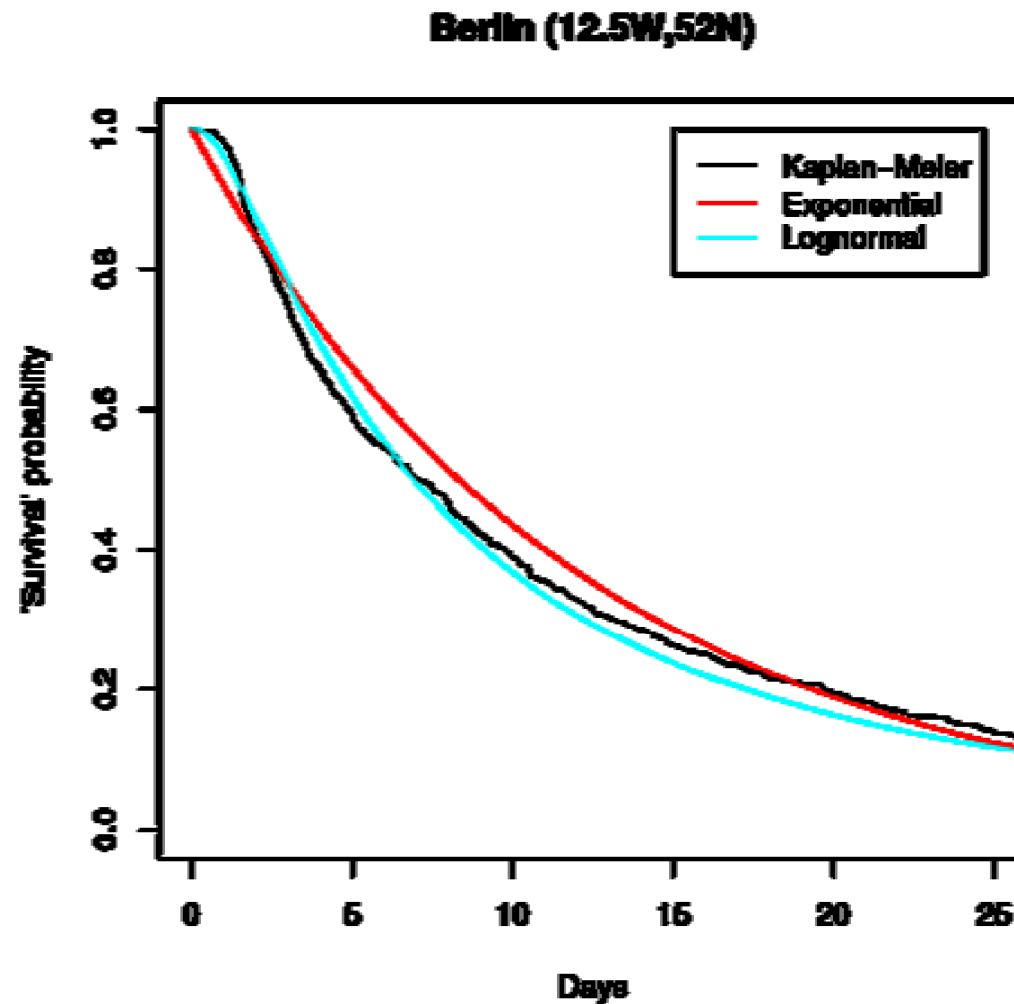
Auto-intensity function for top 50% most intense storms

Conditional probability density that an event arrives at time $T + \text{lag}$, given that an event arrived at time T .



Inhibition at lag=1 day. Storms can't arrive too close (hard-core rejection)
Clustering at lag=2,3,4 days. Higher conditional probability of a 2nd storm.

Survival function for 50% most intense storms



→ Log-normal gives a much better fit than does Exponential (or Gamma & Weibull)

Rate-varying Poisson process for arrival times (*)

$$\Delta t \sim \text{Exp}(\lambda(t))$$

$$\log \lambda(t) = \beta_0 + \sum_k \beta_k X_k(t)$$

Maximum-likelihood estimates for storms passing Berlin:

	Rate	NAO	EAP	EA/WR	SCA
Beta	2.053	0.092	0.175	-0.25	-0.32
2*SE	0.061	0.067	0.064	0.064	0.062

Statistically significant non-zero effects of EAP, SCA and EA/WR patterns.

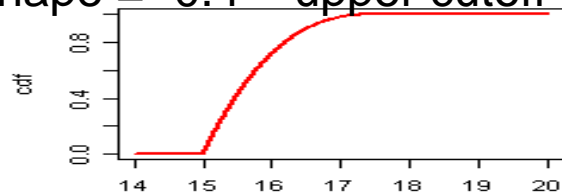
* - of top 50% most intense storms

Generalized Pareto Distribution (GPD)

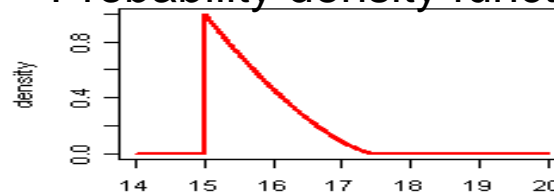
$$F(z) = \Pr\{Z \leq z \mid Z > u\} = 1 - \left[1 + \xi \left(\frac{z - u}{\tilde{\sigma}} \right) \right]^{-1/\xi} \quad \tilde{\sigma} = \sigma + \xi(u - \mu)$$

$$\Rightarrow f(z) = \frac{dF}{dz} = \frac{1}{\tilde{\sigma}} \left[1 + \xi \left(\frac{z - u}{\tilde{\sigma}} \right) \right]^{-1-1/\xi} \rightarrow \begin{cases} 0 & \text{when } z \geq u - \tilde{\sigma} / \xi = \mu - \sigma / \xi \\ \tilde{\sigma}^{-1} \exp(-z / \tilde{\sigma}) & \\ \frac{1}{\tilde{\sigma}} \left(\frac{\xi z}{\tilde{\sigma}} \right)^{-(1+1/\xi)} & \end{cases}$$

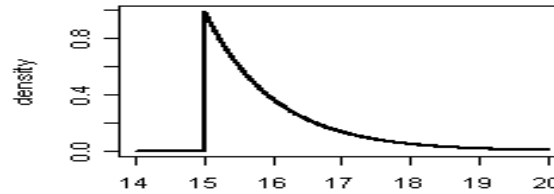
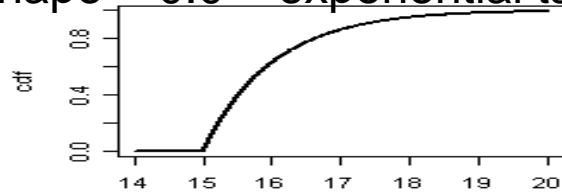
Shape = -0.4 – upper cutoff



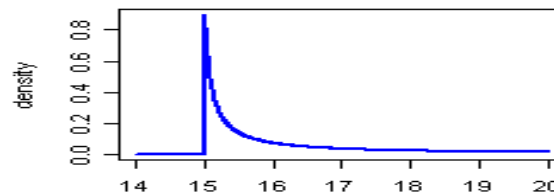
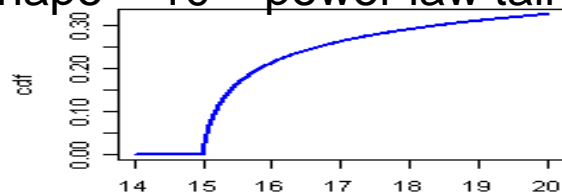
Probability density function



Shape = 0.0 – exponential tail

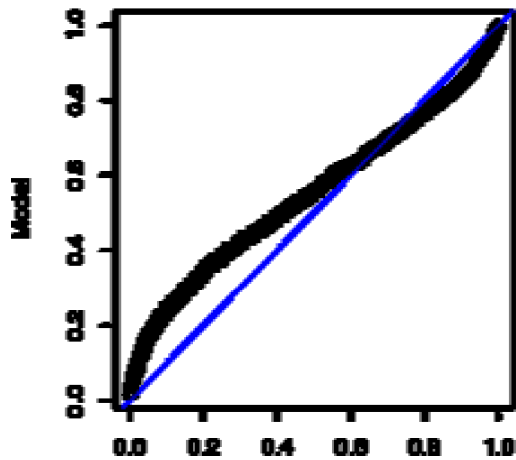


Shape = 10 – power law tail

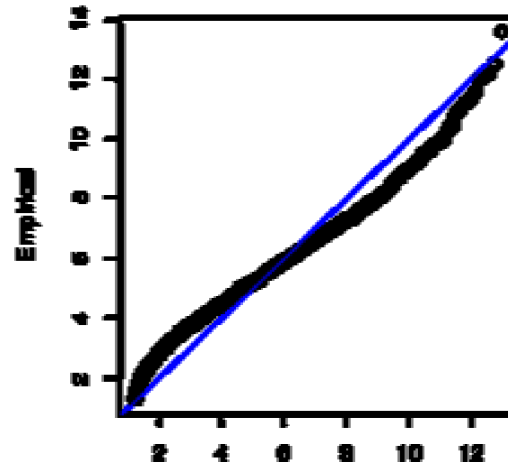


How well does Generalised Pareto Distribution fit the vorticity marks at Berlin?

Probability Plot

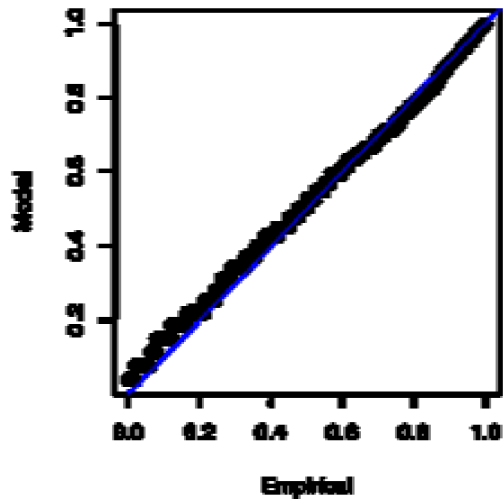


Quantile Plot

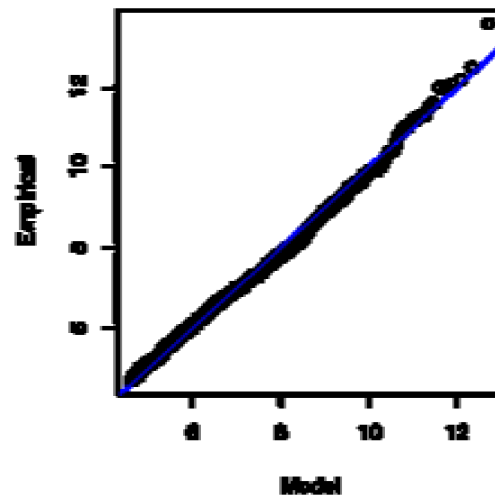


All storms:
Not a good fit

Probability Plot

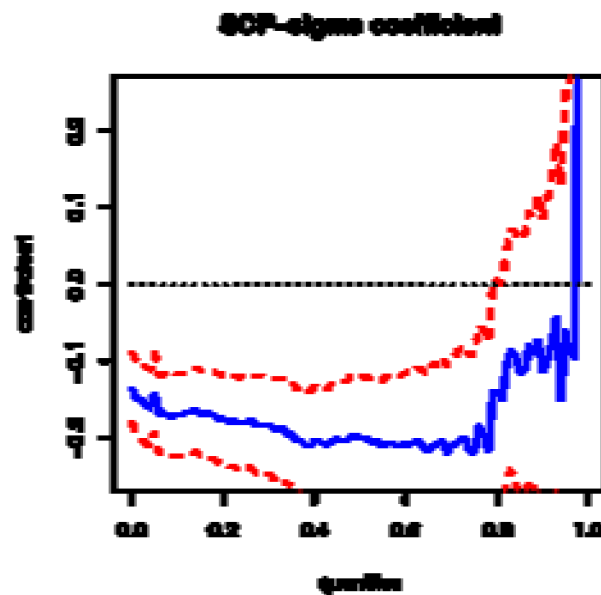
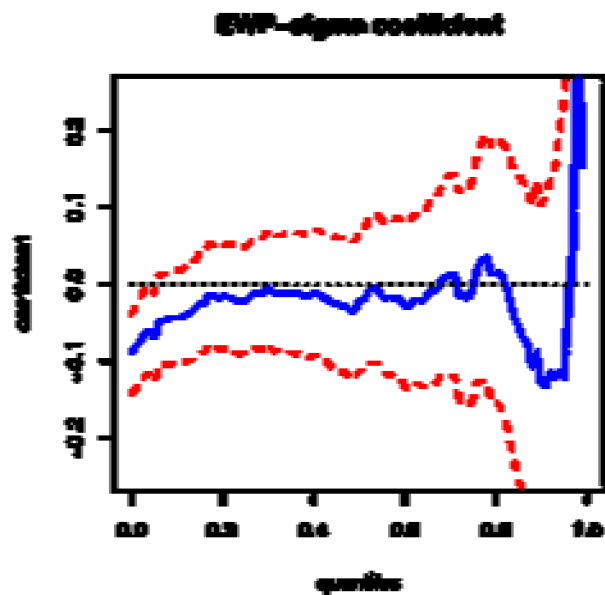
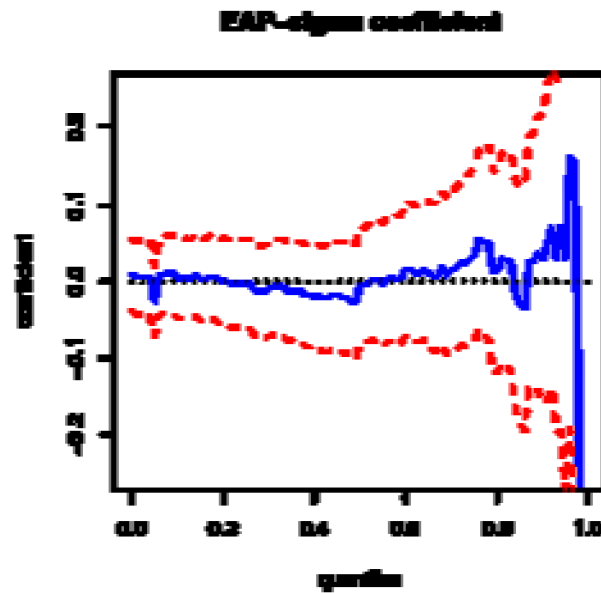
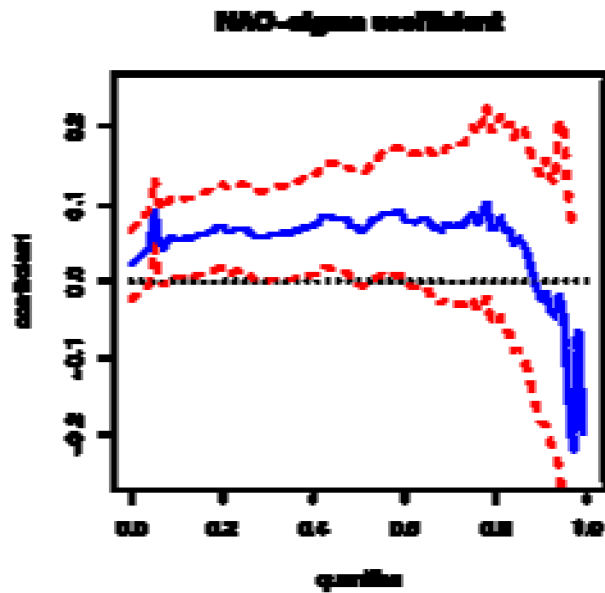


Quantile Plot



Storms with vorticity > 50% quantile:
Much better fit

Including teleconnection covariates in the GPD



$$Y \sim GPD(\sigma, \xi)$$

$$\log \sigma = \alpha_0 + \sum_k \alpha_k X_k(t)$$

blue: estimate of alpha

red: estimate ± 2 s. error

Summary

- Extratropical wintertime storms cluster esp. over NW Europe;
- Dispersion of counts increases for more intense storms;
- Dispersion of counts increases for longer count periods;
- Large-scale teleconnection indices can account for this overdispersion;
- Interarrival times show clustering (and inhibition)
- Have started to use a compound Poisson-GPD process to model the extremes but may need a compound renewal-GPD process

References

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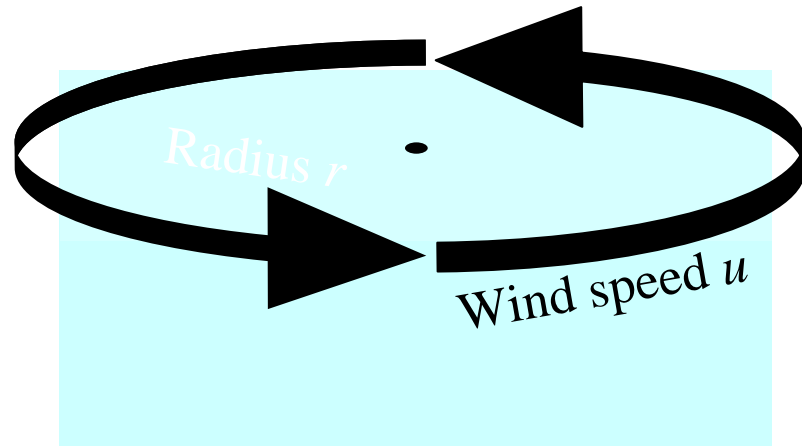
Exeter Climate Systems
www.secam.ex.ac.uk/xcs

Willis Research Network
www.willisresearchnetwork.com



One small
step ...

Relative Vorticity : the “spin”



$$\begin{aligned} \text{VORTICITY} &= \frac{\text{CIRCULATION}}{\text{AREA}} \approx \frac{u \ 2 \pi r}{\pi r^2} \\ &= 2 \frac{u}{r} \\ &= 2 \Omega . \end{aligned}$$

e.g.: wind speed of 15 m/s and radius of 500 km \rightarrow vorticity of 6×10^{-5} /s.

Why use relative vorticity instead of SLP?

- More prominent small-scale features allow earlier detection
- Much less sensitive to the background state
- Directly linked with low-level winds (through circulation) and precipitation (through vertical motion)