

Estimating the jump measure of a Lévy process

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Outline

The model

High versus low frequency

Efficient low-frequency estimation

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Summary

Lévy-Khintchine formula

Lévy-Khintchine characterisation

The characteristic function of a Lévy process L is given by:

$$\varphi_t(u) := \mathbb{E} [e^{iuL_t}] = \exp(t\psi(u))$$

$$\psi(u) := -\frac{\sigma^2}{2}u^2 + ibu + \int_{-\infty}^{\infty} (e^{iux} - 1 - \frac{iux}{1+x^2}) \nu(dx)$$

with volatility $\sigma \geq 0$, drift $b \in \mathbb{R}$ and jump measure ν .

(ν is σ -finite with $\int (1 \wedge x^2) \nu(dx) < \infty$)

Lévy-Itô decomposition if $\nu(\mathbb{R}) < \infty$:

$$L_t = \sigma W_t + \tilde{b}t + \sum_{k=1}^{N_t} Y_k$$

with Brownian motion W_t , $\tilde{b} = b - \int \frac{x}{1+x^2} \nu(dx)$, a Poisson process N_t of intensity $\lambda := \nu(\mathbb{R})$ and jumps $Y_k \sim \nu/\lambda$.

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Statistical problem

Observations

$(L_{t_0}, L_{t_1}, \dots, L_{t_n})$ with $t_{i+1} - t_i = \Delta$ equidistant.

Assumptions

$T = n\Delta \rightarrow \infty$ (long-time asymptotics),

$\Delta = \Delta_n \rightarrow 0$ fast (*high frequency*)

or $\Delta > 0$ fixed resp. $\Delta = \Delta_n \rightarrow 0$ slowly (*low frequency*).

Goal

Estimate the triplet (σ^2, b, ν) with

- $\sigma^2 \geq 0$;
- $b \in \mathbb{R}$;
- measure ν with interest in $\int f d\nu$ for different integrands f .

High-frequency estimation

Idea:

Large increments $X_k := L_{t_k} - L_{t_{k-1}}$ indicate a jump
(larger than diffusive oscillations, i.e. $\gg \sqrt{\Delta \log \log \Delta^{-1}}$)

Figueroa-Lopez, Houdré (2006):

$g_\nu \in L^2$ density of ν , $D \in \mathbb{R} \setminus \{0\}$, (φ_j) ONB in $L^2(D)$:

$$\hat{g}_\nu := \sum_{j=1}^J \hat{\beta}_j \varphi_j \text{ with } \hat{\beta}_j := T^{-1} \sum_{k=1}^n \varphi_j(X_k)$$

If g_ν has Sobolev-regularity s on D and $\Delta_T \rightarrow 0$ sufficiently fast:

$$\begin{aligned} \mathbb{E}[\|\hat{g}_\nu - g_\nu\|_{L^2(D)}^2] &\lesssim T^{-2s/(2s+1)} \\ \mathbb{E}\left[\left(\int_D f \hat{g}_\nu - \int_D f g_\nu\right)^2\right] &\lesssim T^{-1} \text{ for all } f \in L^4 \end{aligned}$$

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Idea for low-frequency estimation

Transformation to density estimation

$X_k := L_{t_k} - L_{t_{k-1}}$, $k = 1, \dots, n$, are i.i.d.

with characteristic function

$$\varphi(u) = \exp\left(\Delta\left(-\frac{\sigma^2}{2}u^2 + ibu + \int_{-\infty}^{\infty} \left(e^{iux} - 1 - \frac{iux}{1+x^2}\right) \nu(dx)\right)\right)$$

Empirical characteristic function

$$\varphi_n(u) := \frac{1}{n} \sum_{k=1}^n e^{iuX_k} \xrightarrow{n \rightarrow \infty} \varphi(\sigma^2, b, \nu; u) \text{ under } \mathbb{P}_{(\sigma^2, b, \nu)}$$

Minimum-distance estimator

$$(\hat{\sigma}_n^2, \hat{b}_n, \hat{\nu}_n) := \operatorname{arginf}_{(\sigma^2, b, \nu)} d(\varphi_n, \varphi(\sigma^2, b, \nu; \bullet))$$

Lévy-Ornstein-Uhlenbeck-type process, option calibration:

Jongbloed/van der Meulen/van der Vaart 2005, Cont/Tankov 2004

Consistency for $\Delta > 0$ fixed

Assumptions on the distance

$$\lim_{n \rightarrow \infty} d(\varphi_n, \varphi(\sigma^2, b, \nu; \bullet)) = 0 \quad \mathbb{P}_{(\sigma^2, b, \nu)} \text{-a.s.}$$

$$\lim_{m \rightarrow \infty} d(\varphi^{(m)}, \varphi) = 0 \Rightarrow \lim_{m \rightarrow \infty} \int_s^t \varphi^{(m)}(u) du = \int_s^t \varphi(u) du \quad \forall s, t \in \mathbb{R}.$$

Example: $d(f, g) = (\int |f - g|^p w)^{1/p}$ with $w \in L^1$, $w > 0$

Then:

$$\begin{aligned} d(\varphi(\hat{\sigma}_n^2, \hat{b}_n, \hat{\nu}_n; \bullet), \varphi(\sigma^2, b, \nu; \bullet)) \\ \leq d(\varphi(\hat{\sigma}_n^2, \hat{b}_n, \hat{\nu}_n; \bullet), \hat{\varphi}_n) + d(\hat{\varphi}_n, \varphi(\sigma^2, b, \nu; \bullet)) \\ \leq 2d(\hat{\varphi}_n, \varphi(\sigma^2, b, \nu; \bullet)) \xrightarrow{\text{a.s.}} 0 \end{aligned}$$

and therefore $\int_s^t \varphi(\hat{\sigma}_n^2, \hat{b}_n, \hat{\nu}_n; u) du \xrightarrow{\text{a.s.}} \int_s^t \varphi(\sigma^2, b, \nu; u) du$.

This implies $\mathbb{P}_{(\hat{\sigma}_n^2, \hat{b}_n, \hat{\nu}_n)} \Rightarrow \mathbb{P}_{(\sigma^2, b, \nu)}$ a.s.

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Consistency II

Does $\mathbb{P}_{(\hat{\sigma}_n^2, \hat{b}_n, \hat{\nu}_n)} \xrightarrow{n \rightarrow \infty} \mathbb{P}_{(\sigma^2, b, \nu)}$ imply $\hat{\sigma}_n^2 \rightarrow \sigma^2$, $\hat{b}_n \rightarrow b$, $\hat{\nu}_n \rightarrow \nu$?

Proposition (Gnedenko/Kolmogorov 1949)

$\mathbb{P}_{(\sigma_n^2, b_n, \nu_n)} \Rightarrow \mathbb{P}_{(\sigma^2, b, \nu)} \iff b_n \rightarrow b$ and $\nu_{\sigma, n} \Rightarrow \nu_\sigma$ weakly

with the finite measure $\nu_\sigma(dx) := \sigma^2 \delta_0(dx) + \frac{x^2}{1+x^2} \nu(dx)$.

Corollary

The minimum-distance estimator is consistent for (b, ν_σ) .

Proposition

The volatility σ^2 cannot be estimated uniformly consistently:
no estimator $\tilde{\sigma}_n^2$ can satisfy for any $\sigma^2 \geq 0$

$$\forall \varepsilon > 0 : \limsup_{n \rightarrow \infty} \sup_{b, \nu} \mathbb{P}_{(\sigma^2, b, \nu)} (|\tilde{\sigma}_n^2 - \sigma^2| > \varepsilon) = 0.$$

The same is true, e.g., for the Blumenthal-Gettoor index.

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Kolmogorov representation

Assumption: $\mathbb{E}[L_t^2] < \infty$ or equivalently $\int x^2 \nu(dx) < \infty$.

Canonical representation (Kolmogorov 1932):

$$\begin{aligned} \varphi(u) &= \exp \left(\Delta \left(-\frac{\sigma^2}{2} u^2 + i u b + \int_{-\infty}^{\infty} (e^{i u x} - 1 - i u x) \nu(dx) \right) \right) \\ &= \exp \left(\Delta \left(i u b + \int_{-\infty}^{\infty} \frac{e^{i u x} - 1 - i u x}{x^2} \nu_{\sigma}(dx) \right) \right) \end{aligned}$$

with $\nu_{\sigma}(dx) = \sigma^2 \delta_0(dx) + x^2 \nu(dx)$.

Note: here $\mathbb{E}[L_1] = b$ and $\text{Var}(L_1) = \nu_{\sigma}(\mathbb{R})$ hold.

Important observation:

$$\Delta^{-1} \frac{d^2}{du^2} \log(\varphi(u)) = \int_{-\infty}^{\infty} \frac{d^2}{du^2} \frac{e^{i u x} - 1 - i u x}{x^2} \nu_{\sigma}(dx) = -\mathcal{F} \nu_{\sigma}(u)$$

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Two constructions of estimators

$$\Delta^{-1} \frac{d^2}{du^2} \log(\varphi(u)) = \int_{-\infty}^{\infty} \frac{d^2}{du^2} \frac{e^{iux} - 1 - iux}{x^2} \nu_{\sigma}(dx) = -\mathcal{F}\nu_{\sigma}(u)$$

1. Plug-in + spectral cut-off at $U > 0$: (cf. Belomestny/Reiß 2006)

$$\mathcal{F}\hat{\nu}_{\sigma}(u) := -\Delta^{-1} \frac{d^2}{du^2} \log(\varphi_n(u)) \mathbf{1}_{[-U, U]}(u), \quad u \in \mathbb{R}$$

2. Minimum-distance estimator with a C^2 -distance to control

$$\left| \frac{\varphi_n''(u)\varphi_n(u) - \varphi_n'(u)^2}{\varphi_n(u)^2} - \frac{\varphi''(u)\varphi(u) - \varphi'(u)^2}{\varphi(u)^2} \right|$$

Two-step estimation procedure

Preliminary estimator:

$$\tilde{b}_n = T^{-1} L_T, \quad \mathcal{F} \tilde{\nu}_{\sigma,n} = -\Delta^{-1} \log(\varphi_n)'' \mathbf{1}(|\varphi_n| \geq \kappa n^{-1/2})$$

Minimum-distance estimator:

optimize locally around $(\tilde{b}_n, \tilde{\nu}_{\sigma,n})$ to find

$$(\hat{b}_n, \hat{\nu}_{\sigma,n}) := \operatorname{arginf}_{(b, \nu_\sigma)} d_{C^2}(\varphi_n, \varphi(b, \nu_\sigma; \bullet))$$

where (weighting $w(u) = \log(e + |u|)^{-1/2-\varepsilon}$)

$$d_{C^2}(g, h) = \max_{j=0,1,2} \Delta^{-j/2} \sup_{u \in \mathbb{R}} (|g^{(j)}(u) - h^{(j)}(u)| w(u))$$

Estimation error

How fast can we estimate $\int f d\nu_\sigma$?

Assume $\mathbb{E}[|L_t|^{4+\varepsilon}] < \infty$. For integrands f of regularity s ($(1+x^2)^{s/2} \mathcal{F}f \in L^1$) we have

$$|\int f d\hat{\nu}_{\sigma,n} - \int f d\nu_\sigma| \lesssim_P v(T, \Delta) \vee T^{-1/2}$$

with

$$\begin{aligned} \sigma > 0 : & & v(T, \Delta) &= (\Delta^{-1} \log(T))^{-s/2} \\ \sigma = 0, |\varphi(u)| \gtrsim e^{-\Delta\alpha|u|} : & & v(T, \Delta) &= (\Delta^{-1} \log(T))^{-s} \\ \sigma = 0, |\varphi(u)| \gtrsim (1+|u|)^{-\beta\Delta} : & & v(T, \Delta) &= T^{-s/(2\beta\Delta)} \\ & & & \text{(times a log-factor)} \end{aligned}$$

These rates are *optimal* in a minimax-sense.

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Estimation vs. discretisation error

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$$|\int f d\hat{\nu}_{\sigma,n} - \int f d\nu_{\sigma}| \lesssim_P T^{-1/2}$$

provided

$$\begin{array}{ll} \sigma > 0 : & \Delta \lesssim T^{-1/s} \\ \sigma = 0, |\varphi(u)| \gtrsim e^{-\Delta\alpha|u|} : & \Delta \lesssim T^{-1/(2s)} \\ \sigma = 0, |\varphi(u)| \gtrsim (1+|u|)^{-\beta\Delta} : & \Delta < s/\beta \end{array}$$

Example:

Gamma-processes ($\varphi(u) = (1 - iu/\lambda)^{-p\Delta}$) require $\Delta < s/p$;
compound Poisson processes do not require Δ to be small.

Main tool in the proof ($\Delta = 1$)

Consider the k -th derivative of the *empirical characteristic process*

$$C_n^{(k)}(u) = \frac{d^k}{du^k} \left(n^{-1/2} \sum_{j=1}^n (e^{iuX_j} - \mathbb{E}[e^{iuX_j}]) \right).$$

Then with the empirical process $\mathbb{G}_n = n^{1/2}(\mathbb{F}_n^{X_1, \dots, X_n} - F^X)$

$$\sup_{u \in \mathbb{R}} |C_n^{(k)}(u)| w(u) = \sup_{h \in H} \int h d\mathbb{G}_n$$

with

$$H = \left\{ x \mapsto w(u) \frac{d^k}{du^k} e^{iux} : u \in \mathbb{R} \right\}.$$

Main tool in the proof (ctd.)

$$H = \left\{ x \mapsto w(u) \frac{d^k}{du^k} e^{iux} : u \in \mathbb{R} \right\}$$

Consequently, empirical process theory yields

$$\sup_n \mathbb{E} \left[\sup_{u \in \mathbb{R}} |C_n^{(k)}(u)| w(u) \right] \lesssim \mathbb{E}[(X_j)^{2k}] + J_{[]} (H)$$

with $J_{[]}$ the *bracketing entropy* of H . For suitable brackets

$$h_j^\pm(x) = (w(u_j) \frac{d^k}{du^k} e^{iu_j x} \pm \varepsilon |x|^k) \mathbf{1}_{[-M, M]}(x) + \|w\|_\infty |x|^k \mathbf{1}_{[-M, M]^c}(x)$$

this $J_{[]}$ is finite for $w(u) \sim (\log(u))^{-1/2-\varepsilon}$ and $\mathbb{E}[|X_j|^{2k+\varepsilon}] < \infty$.

Decay of φ versus singularity of ν

Lemma:

Consider a Lévy process with a characteristic function of at most polynomial decay: $|\varphi_t(u)| \gtrsim (1 + u^2)^{-\beta t}$. Then the Lévy measure ν satisfies for any $\varepsilon > 0$

$$\int_{[-1,1]} \log(|x|^{-1})^{-2-\varepsilon} \nu(dx) < \infty.$$

This implies $\sup_{u \in \mathbb{R}} |\varphi'(u)/\varphi(u)| < \infty$ provided $\varphi \in \mathcal{C}^1$.

Example:

The Lévy measure of the Gamma process has a density of order x^{-1} around zero and $\int_{[-1,1]} \log(|x|^{-1})^{-2-\varepsilon} |x|^{-1} dx < \infty$.

Intuition:

The more singular the Lévy measure at zero, the rougher the sample paths, but the smoother the transition density.

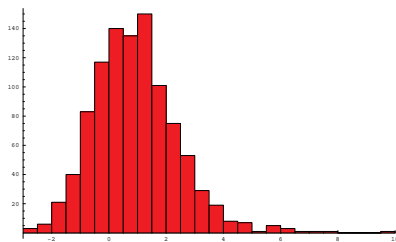
Example

Simulation:

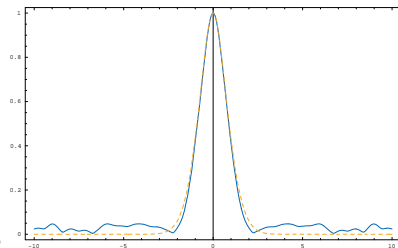
Superposition of Brownian motion and Exp(1)-process

$$\sigma = 1, b = 1, \nu(dx) = x^{-1} e^{-x} \mathbf{1}_{\{x>0\}} dx$$

$$\Delta = 1, n = 1000$$



Histogram of the data.



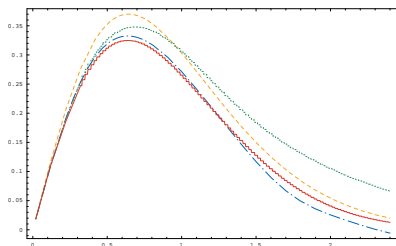
modulus of the empirical (solid blue) and true (dashed orange) characteristic function.

Estimation for the example

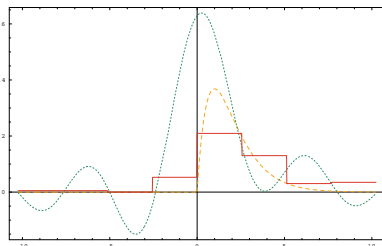
$$\hat{b}_n = 0.922 \text{ (true } b = 1), \hat{\nu}_{\sigma,n}(\{0\})^{1/2} = 1.092 \text{ (true } \sigma = 1)$$

Estimate of $\nu([1, \infty)) = \int_1^\infty x^{-2} \nu_\sigma(dx)$ is 0.16 (true: 0.22).

The HF-estimator $\frac{1}{n} \sum_{i=1}^n \mathbf{1}_{\{L_i - L_{i-1} \geq 1\}}$ yields 0.46.



$\text{Im}(\varphi_n)$ (blue), $\text{Im}(\varphi)$ (orange),
 $\text{Im}(\tilde{\varphi}_n)$ (green) and $\text{Im}(\hat{\varphi}_n)$ (red).



density of $\tilde{\nu}_\sigma$ (green), $\hat{\nu}_\sigma$ (solid)
 and ν_σ (orange);
 $\tilde{\nu}_\sigma$ has no point mass in zero.

Summary

- High. vs. low-frequency observations of a Lévy process
- LF-estimation based on minimum-distance fit of characteristic function.
- LF-estimation of σ^2 cannot be uniformly consistent.
- Adjust to any observation distance Δ .
- LF-rate for $b = \mathbb{E}[L_1]$ is parametric.
- LF-rate for $\int f d\nu_\sigma$ depends on the decay of φ .
- LF-rate $T^{-1/2}$ for $\int f d\nu_\sigma$ if $\Delta < \bar{\Delta}(\varphi, s; T)$ (automatic).

Thank you very much for your attention!

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