

# A trivia question

What is in common between the following locations:

Eindhoven

Leiden

Sydney

Amsterdam

Hadera

Haifa

France

Shfayim

# DROS

Discriminatory random order service:

- Each customer possesses a parameter  $p_i$
- Upon service commencement (no preemption), customer  $i$  enters service with probability  $p_i / \sum_j p_j$ .

Haviv and van der Wal 1997: M/M/1, parameter  $x$ , costs  $x$ .  
What is the equilibrium purchasing strategy?

Answer: pure strategy. Pay

$$\frac{C\rho^2}{\mu(1-\rho)(2-\rho)}$$

# DPS

Similar result for DPS: In DROS lotteries at service commencements, in DPS it is at service completions.

Still open: Equilibrium payment in case of M/G/1? (for both DROS and DPS)

# M/G/1 with relative priority

Class  $i$ :  $\lambda_i, \bar{x}_i, \overline{x^2}_i, p_i$ .

Mean value analysis: Haviv and Van der Wal (2008).

The same if HOL is assumed among classes.

Higher moments: A paper by .... (under review).

Higher moments in case of HOL?

# **When to arrive at a queue with tardiness costs?**

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Eindhoven, 30.09.10

# Concert hall with early birds

- gate opens at zero and closes at  $T$
- FCFS, inclusive of early birds
- all arrivals prior to  $T$  are served
- single server
- service  $\exp(\mu)$
- $N$  = no. of arrivals  $\text{Poisson}(\lambda)$
- $\alpha$  = cost per unit of queueing
- $\beta > 0$  = cost per unit of tardiness (from zero)

# Equilibrium

Symmetric (Nash) equilibrium: an arrival strategy (mixing is possible), if used by all, nobody has an incentive to do otherwise

Hassin and Glazer (1983):  $\beta = 0$  and  $T$  finite

Jain, Juneja and Shimkin, (2010): Fluid approximation,  
 $T = \infty$

Juneja and Shimkin (2010+):  $\beta > 0$ ,  $T = \infty$ , any distribution for  $N$

# Equilibrium, $\beta > 0, T = \infty$

- not a pure strategy
- mixed strategy but without atoms
- mixed strategy with a positive density along an interval
- the arrival interval  $[-w, T_e]$
- uniform density along  $[-w, 0)$
- continuous density but not at zero (downwards)

Assume  $\beta > 0$ : Otherwise,  $T_e = \infty$  and zero waiting costs



# Equilibrium conditions, $T = \infty$

$f(t)$ : density of the arrival strategy

$$\int_{-w}^{T_e} f(t) dt = 1$$

$f(t)$  determines the queueing process

$w(t)$  = mean queueing time if arrive at  $t$

Equilibrium conditions:

$$(\alpha + \beta)w(t) + \beta t = \text{Constant}, \quad -w \leq t \leq T_e$$

$$(\alpha + \beta)w(t) + \beta t \geq \text{Constant}, \quad t < -w, t > T_e$$

Reverse engineering: Find  $w$ ,  $T_e$  and  $f(t)$  such that the equilibrium conditions hold

# Equilibrium

Equilibrium:

$$f(t) = \frac{\mu}{\lambda} \frac{\alpha}{\alpha + \beta}, \quad -w \leq t < 0$$

$f(t)$  is discontinuous at  $t = 0-$

$$\int_0^{T_e} f(t) dt = 1 - w \frac{\mu}{\lambda} \frac{\alpha}{\alpha + \beta}$$

Initial conditions:

$$P_k(0) = e^{-w\mu \frac{\alpha}{\alpha+\beta}} \frac{(w\mu \frac{\alpha}{\alpha+\beta})^k}{k!}, \quad k \geq 0$$

Equilibrium:

$$f(t) = \frac{(1 - P_0(t))\mu}{\lambda} - \frac{\beta\mu}{(\alpha + \beta)\lambda}, \quad 0 \leq t \leq T_e$$

Dynamics:

$$P'_0(t) = P_1(t)\mu - P_0(t)\lambda f(t), \quad 0 < t < T_e$$

$$P'_k(t) = P_{k-1}(t)\lambda f(t) + P_{k+1}(t)\mu - P_k(t)(\lambda f(t) + \mu), \quad 0 < t < T_e, \quad k \geq 1$$

Equilibrium:

$$\alpha(1 - P_0(T_e)) = \beta P_0(T_e) \quad (\text{or } f(T_e) = 0)$$

# Equilibrium, $T < \infty$

- If  $T > T_e$ , as  $T = \infty$
- If  $T < T_e$ , replace  $T_e$  with  $T$  and ignore the last condition

$$\alpha(1 - P_0(T_e)) = \beta P_0(T_e)$$

In fact,

$$\alpha(1 - P_0(T)) > \beta P_0(T)$$

Social cost:  $\lambda \alpha w$

# Concert hall w/o early birds

- gate opens at zero and closes at  $T$
- FCFS, exclusive of early birds
- early birds enter at random
- all arrivals prior to  $T$  are served
- single server
- service  $\exp(\mu)$
- $N$  = no. of arrivals  $\text{Poisson}(\lambda)$
- $\alpha$  = cost per unit of queueing
- $\beta > 0$  = cost per unit of tardiness (from zero)

Hassin and Kleiner (2010):  $\beta = 0$ ,  $T$  finite

# Equilibrium

1. if  $T \leq T_1$ , pure: arrive at zero  
 $T_1 = \infty$  is possible
2. if  $T_1 < T \leq T_e$ ,
  - atom at zero
  - positive density along  $[t', T]$
3. if  $T > T_e$ 
  - atom at zero
  - positive density along  $[t', T_e]$

# Equilibrium

$N_p$  Poisson( $\lambda p$ )

$X_i$ , iid,  $\exp(\mu)$

$$g(t) = (\alpha + \beta) \mathbf{E} \left( \sum_{i=0}^{N_1} X_i - t \right)^+ + \beta t, \quad t \geq 0$$

$$t^* = \arg \min_{t \geq 0} g(t)$$

$$\text{If } g(t^*) \geq \lambda(\alpha + \beta)/2\mu$$



Pure equilibrium: arrive at  $t = 0$  (for any  $T$ )

# Equilibrium, $T < \infty$

Assume  $g(t^*) \leq \lambda(\alpha + \beta)/2\mu$

$T_1$  = the smallest (among two)  $t$  such that

$$g(t) = (\alpha + \beta) \frac{\lambda}{2\mu}$$

$T_e$  = the latest time to arrive in equilibrium when  $T = \infty$   
(needs to be determined).

The shape of the equilibrium depends if

- $T \leq T_1$  (pure), or
- $T_1 \leq T \leq T_e$  (mixed), or
- $T \geq T_e$  as for  $T_e$



# Equilibrium

If  $T \leq T_1 \Rightarrow$  pure strategy: arrive at  $t = 0$

If  $T_1 \leq T \leq T_e$ ,

- atom of size  $p_0$  at zero
- zero density along  $(0, t')$
- positive density along  $[t', T]$

$$(\alpha + \beta) \frac{\lambda p_0}{2\mu} = (\alpha + \beta) \mathbf{E} \left( \sum_{i=0}^{N_{p_0}} X_i - t' \right)^+ + \beta t',$$

One dimensional search for  $p_0$  based on

$$\int_{t'}^T f(t) dt = 1 - p_0$$

# Equilibrium

If  $T \geq T_e$  as in  $T_e$ .

Finding  $T_2$ :

- For any  $T \in (T_1, T_e)$ ,  $\alpha(1 - P_0(T)) > \beta P_0(T)$  : A bit after  $T$  is a better response (yet, not feasible)
- $T_e$  is the smallest  $T$  with  $\alpha(1 - P_0(T)) = \beta P_0(T)$

The social cost= $\lambda\alpha w$ .

# Fluid approximation

- A mass of water of size  $\Lambda$
- rate of service  $\mu$  units of water per unit of time
- each drop needs to decide when to arrive

# With early birds, fluid, $T \geq \Lambda/\mu$

Jain, Juneja and Shimkin, 2010

## Equilibrium:

- uniform arrival along  $[-\Lambda\beta/(\mu\alpha), \Lambda/\mu]$ , rate  $\mu \frac{\alpha}{\alpha+\beta}$
- social cost  $= \Lambda^2\beta/\mu$  (no  $\alpha$ )

## Social optimization:

- uniform arrival along  $[0, \Lambda/\mu]$ , rate  $\mu$
- social cost  $\Lambda^2\beta/2\mu$  (no waiting)

PoA=2, constantly

# With early birds, fluid, $T < \Lambda/\mu$

## Equilibrium:

- Shift all to the left, make  $T$  the upper end of the arrival interval
- Social cost:  $\Lambda(\Lambda(\alpha + \beta) - \alpha\mu T)/\mu$

## Social optimization:

- Arrive with rate  $\mu$  along  $[0, T]$ . The rest at  $T$ .

# Without early birds, fluid

**Social optimization:** As with early birds

**Equilibrium:**

- **If  $\beta > \alpha$ ,**
  - pure strategy: arrive at 0
  - social cost:  $\Lambda^2(\alpha + \beta)/2\mu$  (an improvement)
  - $\text{PoA} = (\alpha + \beta)/\beta$
- **If  $\beta \leq \alpha$ ,**
  - an atom of  $2\beta/(\alpha + \beta)$  at zero
  - a gap along  $(0, \Lambda\beta/(\alpha\mu))$  (length as early birds horizon)
  - constant rate of  $\mu\alpha/(\alpha + \beta)$  along  $[\Lambda\beta/(\alpha\mu), \Lambda/\mu]$
  - social cost  $\Lambda^2\beta/\mu$  (as with early birds)
  - $\text{PoA} = 2$

# Why Poisson?

- huge potential arrivals  $n$
- each comes with a tiny probability  $p$
- number of arrivals is Poisson with mean  $np$

An external inspector believes that the number of arrivals is Poisson

Each arrival believe the same (and same parameter) regarding the number of other arrivals

# Common prior

Suppose

$$p_k = \frac{(k+1)q_{k+1}}{m}, \quad k \geq 0.$$

$$q_k \geq 0, \quad k \geq 0, \quad \sum_{k=0}^{\infty} q_k = 1, \quad m = \sum_{k=0}^{\infty} k q_k$$

The  $p_k$ 's are the (common) posterior of the (common) prior  $q_k$ 's



# Common prior

Up to the choice of  $q_0$ , any nonnegative distribution is a posterior of a unique prior

Families closed under posterior operation:

- $\text{Poisson}(\lambda) \Rightarrow \text{Poisson}(\lambda)$ . Unique!
- $\text{Binomial}(n, p) \Rightarrow \text{Binomial}(n - 1, p)$
- $\text{Negative binomial}(n, p) \Rightarrow \text{Negative binomial}(n + 1, p)$



Thank You