On the damped geometric telegrapher's process

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Abstract The geometric telegrapher's process has been proposed in 2002 as a model to describe the dynamics of the price of risky assets. In this contribution we consider a similar stochastic process, for which the random times between consecutive slope changes have exponential distribution with linearly increasing parameters. This leads to a process characterized by a damped behavior. We study the main features of the transient probability law of the process, and of its stationary limit.

2 The stochastic model and probability laws

Let us assume that the price of risky assets is described by the following stochastic process, that we shall call damped geometric telegrapher's process:

$$S_t = s_0 \exp[at + X_t], \quad \text{with} \quad X_t = c \int_0^t (-1)^{N_{\text{T}}} d\tau, \quad t \ge 0,$$
 (2)

where $s_0 > 0$, $a \in \mathbb{R}$, c > 0, and where N_t is an alternating counting process characterized by independent random times $U_k, D_k, k \ge 1$. Hence,

$$N_0 = 0,$$
 $N_t = \sum_{n=1}^{\infty} \mathbf{1}_{\{T_n \le t\}},$ $t > 0,$

where $T_{2k} = U^{(k)} + D^{(k)}$ and $T_{2k+1} = T_{2k} + U_{k+1}$ for k = 0, 1, ..., with $U^{(0)} = D^{(0)} = 0$ and

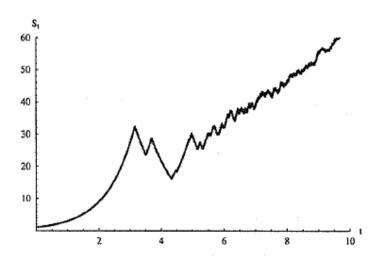
$$U^{(k)} = U_1 + U_2 + \dots + U_k, \qquad D^{(k)} = D_1 + D_2 + \dots + D_k, \qquad k = 1, 2, \dots$$
 (3)

We assume that $\{U_k\}$ and $\{D_k\}$ are mutually independent sequences of independent random variables characterized by exponential distribution with parameters

$$\lambda_k = \lambda k, \qquad \mu_k = \mu k, \qquad (\lambda, \mu > 0, k = 1, 2, \dots), \tag{4}$$

respectively. We remark that process S_t has bounded variations and its sample-paths are constituted by connected lines having exponential behavior, characterized alter-

nately by growth rates a + c and a - c, where a is the growth rate of risky assets' price in the absence of randomness, and c is the intensity of the random factor of alternating type. Assumption (4) implies that the reversal rates λ_k and μ_k linearly increase with the number of reversals, so that the sample paths of S_t are subject to an increasing number of slope changes when t increases, this giving a damped behavior. (An example is shown in Fig. 1).



Proposition 1. For k = 1, 2, ... we have

$$F^{(k)}(u) := P(U^{(k)} \le u) = (1 - e^{-\lambda u})^k, \qquad u \ge 0.$$
(5)

Proof. The proof proceeds by induction on k. For k = 1, the result is obvious. Let us now assume (5) holding for all m = 1, ..., k-1. Denoting by \widetilde{U}_j ($j \ge 1$) independent and exponentially distributed random variables with parameters λ , due to (4) and recalling (3) we have

$$U^{(k)} \stackrel{\mathrm{D}}{=} \sum_{j=1}^{k} \frac{\widetilde{U}_{j}}{j}, \qquad k = 1, 2, \dots$$

Then, due to independence, we have

$$\begin{split} F^{(k)}(u) &= \lambda k \int_0^u \mathrm{e}^{-\lambda k y} (1 - \mathrm{e}^{-\lambda (u - y)})^{k - 1} \mathrm{d}y \\ &= \lambda k \sum_{j = 0}^{k - 1} (-1)^j \binom{k - 1}{j} \mathrm{e}^{-\lambda j u} \int_0^u \mathrm{e}^{-\lambda (k - j) y} \mathrm{d}y \\ &= \frac{k}{k - j} \sum_{j = 0}^{k - 1} (-1)^j \binom{k - 1}{j} \mathrm{e}^{-\lambda j u} [1 - \mathrm{e}^{-\lambda (k - j) u}] \\ &= \sum_{j = 0}^{k - 1} (-1)^j \binom{k}{j} [\mathrm{e}^{-\lambda j u} - \mathrm{e}^{-\lambda k u}] = \sum_{j = 0}^{k - 1} (-1)^j \binom{k}{j} \mathrm{e}^{-\lambda j u} + (-1)^k \mathrm{e}^{-\lambda k u} \\ &= (1 - \mathrm{e}^{-\lambda u})^k, \end{split}$$

this giving (5).

In order to obtain the distribution function of process X_t , let us now introduce the compound process

$$Y_t = \sum_{n=0}^{M_t} D_n$$
, where $M_t := \max\{n \ge 0 : \sum_{j=1}^n U_j \le t\}$, $t > 0$.

Notice that $P(Y_t = 0) = e^{-\lambda t}$. Hereafter we obtain the distribution function of Y_t .

Proposition 2. For any fixed t > 0 and $y \in (0, +\infty)$, we have

$$H(y,t) := P(Y_t \le y) = \frac{e^{-\lambda t}}{e^{-\lambda t} + e^{-\mu y} (1 - e^{-\lambda t})}.$$
 (7)

Proof. For t > 0 the distribution function of Y_t can be expressed as

$$H(y,t) = \sum_{n=0}^{+\infty} P(M_t = n) G^{(n)}(y),$$

where, due to (5),

$$P(M_t = n) = F^{(n)}(t) - F^{(n+1)}(t) = e^{-\lambda t} (1 - e^{-\lambda t})^n, \quad n = 0, 1, \dots$$

Hence, recalling (6), we obtain

$$H(y,t) = e^{-\lambda t} \sum_{n=0}^{+\infty} (1 - e^{-\lambda t})^n (1 - e^{-\mu y})^n,$$

$$W_t = \int_0^t \mathbf{1}_{\{N_s \text{ even}\}} \mathrm{d}s, \qquad t > 0,$$

so that

$$X_t = c(2W_t - t), t > 0.$$
 (8)

Proposition 3. For all $0 < \tau < t$, the distribution function of W_t is:

$$P(W_t \le \tau) = \frac{e^{-\mu(t-\tau)}(1 - e^{-\lambda \tau})}{e^{-\lambda \tau} + e^{-\mu(t-\tau)}(1 - e^{-\lambda \tau})}.$$
 (9)

Moreover,

$$P(W_t < t) = 1 - e^{-\lambda t}, \qquad P(W_t \le t) = 1.$$

Proof. Note that, for a fixed value $t_0 > 0$,

$$W_{t_0} = \inf\{t > 0 : Y(t) \ge t_0 - t\}. \tag{10}$$

Moreover, if $W_{t_0} = \tau$, $\tau \le t_0$, and $Y_{\tau} = t_0 - \tau$ ($Y_{\tau} > t_0 - \tau$), then the motion is moving upward (downward) at time t_0 . Finally, since Y_t is an increasing process, due to (10), the survival distribution of W_t is given by

$$P(W_t > \tau) = H(t - \tau, \tau), \qquad 0 < \tau \le t,$$

so that Eq. (9) follows.

Proposition 4. Let $\tau_* = \tau_*(x,t) = (x+ct)/(2c)$. For all t > 0 and x < ct we have

$$P(X_{l} \leq x) = \frac{e^{-\mu(l-\tau_{*})}(1 - e^{-\lambda \tau_{*}})}{e^{-\lambda \tau_{*}} + e^{-\mu(l-\tau_{*})}(1 - e^{-\lambda \tau_{*}})}.$$

Moreover, $P(X_t < ct) = 1 - e^{-\lambda t}$ and $P(X_t \le ct) = 1$.

In the following proposition we finally obtain the distribution function of S_t .

Proposition 5. For all t > 0 and $x < s_0 e^{(a+c)t}$, we have

$$P(S_t \le x) = \frac{A_{\mu}(t) [x/s_0]^{(\lambda+\mu)/(2c)} - A_{\mu}(t) A_{\lambda}(t) [x/s_0]^{\mu/(2c)}}{A_{\lambda}(t) + A_{\mu}(t) [x/s_0]^{(\lambda+\mu)/(2c)} - A_{\mu}(t) A_{\lambda}(t) [x/s_0]^{\mu/(2c)}},$$

where $A_{\lambda}(t) = \exp\left\{-\frac{\lambda}{2}\left(1 - \frac{a}{c}\right)t\right\}$ and $A_{\mu}(t) = \exp\left\{-\frac{\mu}{2}\left(1 + \frac{a}{c}\right)t\right\}$. Moreover,

$$P(S_t < s_0 e^{(a+c)t}) = 1 - e^{-\lambda t}, \qquad P(S_t \le s_0 e^{(a+c)t}) = 1.$$

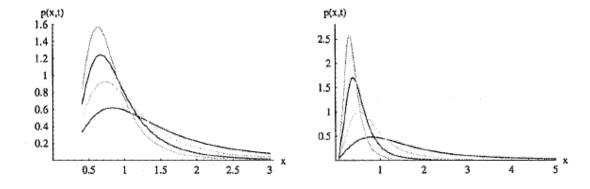


Fig. 2 Piot of p(x,t) for $s_0 = 1$, a = 0.1, c = 1, $\mu = 2$, and $\lambda = 2,3,4,5$, from bottom to top near the origin, with t = 1 (left-hand-side) and t = 3 (right-hand-side).

Proposition 6. For all t > 0 we have

$$P\{S_t = s_0 e^{(a+c)t}\} = e^{-\lambda t},$$

and, for $x \in (s_0 e^{(a-c)t}, s_0 e^{(a+c)t})$

$$p(x,t) := \frac{d}{dx} P(S_t \le x) = \frac{\lambda + \mu - \mu \left(\frac{x}{s_0}\right)^{-\frac{\lambda}{2c}} \exp\left\{-\frac{\lambda t(c-a)}{2c}\right\}}{2cx} \times \frac{1}{\left\{2\cosh\left\{\frac{\lambda + \mu}{4c}\log\left(\frac{x}{s_0}\right) + \frac{\lambda t(c-a) - \mu t(c+a)}{4c}\right\} - \left(\frac{x}{s_0}\right)^{-\frac{\lambda - \mu}{4c}} \exp\left\{-\frac{\lambda t(c-a) + \mu t(c+a)}{4c}\right\}\right\}^2}$$

Some plots of density p(x,t) are shown in Fig. 2 for various choices of λ and t. Let us now analyze the behavior of p(x,t) in the limit as t tends to $+\infty$.

Corollary 1. If $\lambda(c-a) = \mu(c+a)$ then

$$\lim_{t \to +\infty} p(x,t) = \frac{\beta}{s_0} \frac{(x/s_0)^{\beta - 1}}{[1 + (x/s_0)^{\beta}]^2}, \qquad x \in (0, +\infty),$$

where $\beta = \lambda/(c+a)$; whereas, if $\lambda(c-a) \neq \mu(c+a)$ then

$$\lim_{t\to +\infty} p(x,t) = 0.$$

Corollary 2. Let $\alpha_i = s_0 \exp\{at\}$. If $\lambda = \mu \to +\infty$, $c \to +\infty$, with $\lambda/c \to \theta$, then

$$p(x,t) \to \frac{\theta}{\alpha_t} \frac{(x/\alpha_t)^{\theta-1}}{[1+(x/\alpha_t)^{\theta}]^2}, \qquad x \in (0,+\infty).$$

We now analyse the behavior of p(x,t) when x approaches the endpoints of its support, i.e. the interval $[s_1, s_2] := [s_0 e^{(a-c)t}, s_0 e^{(a+c)t}].$

Corollary 3. For any fixed t > 0, we have

$$\lim_{x \downarrow s_1} p(x,t) = \frac{\lambda}{2cs_0} e^{(c-a-\mu)t}, \qquad \lim_{x \uparrow s_2} p(x,t) = \frac{[\lambda + \mu(1-e^{-\lambda t})] e^{-(c+a+\lambda)t}}{2cs_0}.$$

Hereafter we express the *m*-th moment of S_t in terms of the Gauss hypergeometric function ${}_2F_1$.

Proposition 7. Let m be a positive integer. Then, for t > 0,

$$E[S_t^m] = s_0^m e^{m(a-c)t} \left\{ 1 + \frac{2mc}{\lambda} \sum_{k=0}^{+\infty} \frac{(1 - e^{-\lambda t})^{k+1}}{k+1} \right.$$

$$\times \sum_{r=0}^k {k \choose r} (-e^{-\mu t})^r {}_2F_1 \left(\frac{2mc}{\lambda} + \frac{\mu}{\lambda} r, k+1; k+2; 1 - e^{-\lambda t} \right) \right\}. \quad (11)$$

Proof. Due to Proposition 4, by setting y = (ct + x)/2c we have

$$M_{X_t}(s) := E\left[e^{sX(t)}\right] = e^{-sct} \left\{ 1 + 2sc \int_0^t \frac{e^{-(\lambda - 2cs)y}}{e^{-\lambda y} + e^{-\mu(t-y)}(1 - e^{-\lambda y})} dy \right\}. \quad (12)$$

After some calculations Eq. (12) becomes

$$M_{X_t}(s) = e^{-sct} \left\{ 1 + \frac{2sc}{\lambda} \sum_{k=0}^{+\infty} \sum_{r=0}^{k} {k \choose r} (-e^{-\mu t})^{k-r} \int_{\mathscr{I}} x^k (1-x)^{-[2cs+\mu(k-r)]/\lambda} dx \right\},\,$$

where $\mathscr{I} = (0, 1 - e^{-\lambda t})$. Hence, recalling Eq. (3.194.1) of [5], and noting that $E[S_t^m] = s_0^m e^{mat} M_{X_t}(m)$ due to (2), the right-hand-side of (11) immediately follows.

Remark 1. If $\lambda = \mu$, then the moment (11) can be expressed as:

$$E[S_t^m] = s_0^m e^{m(a-c)t} \left\{ 1 + \frac{2mc}{\lambda} \sum_{k=0}^{+\infty} \frac{(k!)^2 (1 - e^{-\lambda t})^{2k+1}}{(2k+1)!} \right.$$
$$\times {}_2F_1 \left(\frac{2mc}{\lambda} + k, k+1; 2k+2; 1 - e^{-\lambda t} \right) \left. \right\}.$$

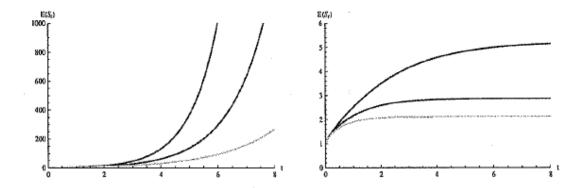


Fig. 3 Plot of $E(S_t)$ for $(\lambda, \mu) = (1.5, 0.9), (1.75, 1.05), (2, 1.2)$ (left-hand-side) and for $(\lambda, \mu) = (3, 1.8), (3.5, 2.1), (4, 2.4)$ (right-hand-side) from top to bottom, with $s_0 = 1$, a = 0.5, c = 2.

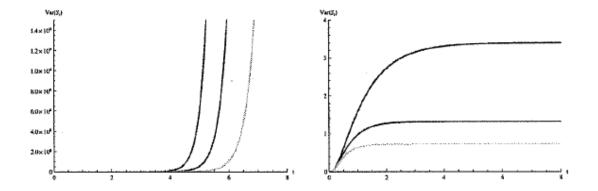


Fig. 4 Plot of $Var(S_t)$ for $(\lambda, \mu) = (1, 0.6), (1.5, 0.9), (2, 1.2)$ (left-hand-side) and for $(\lambda, \mu) = (6, 3.6), (7, 4.2), (8, 4.8)$ (right-hand-side) from top to bottom, with $s_0 = 1$, a = 0.5, c = 2.

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