

Goodness-of-Fit Testing with Empirical Copulas

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Overview of Copulas

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- A bivariate copula C is a bivariate cdf defined on $[0, 1]^2$ with uniform marginal distributions on $[0, 1]$.
- More precisely, a function $C : [0, 1]^2 \rightarrow [0, 1]$ is called a bivariate copula if
 - $C(x, 0) = C(0, y) = 0$ for any $x, y \in [0, 1]$
 - $C(x, 1) = x, C(1, y) = y$ for any $x, y \in [0, 1]$
 - $C(x_2, y_2) - C(x_1, y_2) - C(x_2, y_1) + C(x_1, y_1) \geq 0$
 for any $x_1, x_2, y_1, y_2 \in [0, 1]$ with $x_1 \leq x_2$ and $y_1 \leq y_2$

- Sklar's Theorem:

Let H be a bivariate cdf with continuous marginal cdf's

$$H(x, \infty) = F(x), \quad H(\infty, y) = G(y).$$

Then there exists a unique copula C such that

$$H(x, y) = C(F(x), G(y)). \quad (1)$$

Conversely, for any univariate cdf's F and G and any copula C , (1) defines a bivariate cdf H with marginals F and G .

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- C captures the dependence structure of two random variables. It is used for dependence modeling in finance and actuarial science.

Goodness-of-Fit Testing

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- In other words, we would like to perform a hypothesis test about C , with a null hypothesis of the form $C = C_0$ or $C \in \{C_\theta, \theta \in \Theta\}$. For now, we consider the simple hypothesis ($C = C_0$) only.

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- In other words, we would like to perform a hypothesis test about C , with a null hypothesis of the form $C = C_0$ or $C \in \{C_\theta, \theta \in \Theta\}$. For now, we consider the simple hypothesis ($C = C_0$) only.
- We would like to use the *empirical copula* to construct a goodness-of-fit test.

- Note that we can write

$$C(x, y) = H(F^{-1}(x), G^{-1}(y)), \quad (x, y) \in [0, 1]^2,$$

with $F^{-1}(x) = \inf\{t \in \mathbb{R} : F(t) \geq x\}$, and similarly for G^{-1} .

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- So a natural way of estimating the copula C is using the *empirical copula*

$$C_n(x, y) = H_n(F_n^{-1}(x), G_n^{-1}(y)), \quad (x, y) \in [0, 1]^2,$$

with

$$H_n(x, y) = \frac{1}{n} \sum_{i=1}^n \mathbf{1}\{X_i \leq x, Y_i \leq y\},$$

$$F_n(x) = \frac{1}{n} \sum_{i=1}^n \mathbf{1}\{X_i \leq x\}, \quad G_n(y) = \frac{1}{n} \sum_{i=1}^n \mathbf{1}\{Y_i \leq y\}$$

- It is known that the *empirical copula process*

$$D_n(x, y) = \sqrt{n}(C_n(x, y) - C(x, y)), \quad (x, y) \in [0, 1]^2$$

converges weakly in $\ell^\infty([0, 1]^2)$ to a C -Brownian pillow, under the assumption that

$C^x(x, y)$ is continuous on $\{(x, y) \in [0, 1]^2 : 0 < x < 1\}$,

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- A C -Brownian sheet $W(x, y)$ is a mean zero Gaussian process with covariance function

$$\text{Cov}[W(x, y), W(x', y')] = C(x \wedge x', y \wedge y'), \quad x, x', y, y' \in [0, 1].$$

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- A C -Brownian pillow $D(x, y)$ is a mean zero Gaussian process that is equal in distribution to the C -Brownian sheet W , conditioned on $W(x, y) = 0$ for any $(x, y) \in [0, 1]^2 \setminus (0, 1)^2$.

- We have

$$D(x, y) = W(x, y) - C^x(x, y)W(x, 1) - C^y(x, y)W(1, y) - (C(x, y) - xC^x(x, y) - yC^y(x, y))W(1, 1).$$

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- So we know the asymptotic distribution of the empirical copula process

$$D_n(x, y) = \sqrt{n}(C_n(x, y) - C(x, y)),$$

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and we can take a functional of D_n (such as the sup over $[0, 1]^2$ or an appropriate integral) as a test statistic for a goodness-of-fit test.

- Problem: The asymptotic distribution of D_n , and that of the test statistic, depends on C . We would like to have a *distribution-free* goodness-of-fit test.

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- We use E. Khmaladze's "scanning" idea to transform D into a standard two-parameter Wiener process Z defined on $[0, 1]^2$. The same transformation applied to D_n will then produce a process Z_n that will, hopefully, converge to Z .
- Assumptions on C : Continuous first-order partial derivatives on $[0, 1]^2 \setminus \{(0, 0), (0, 1), (1, 0), (1, 1)\}$, continuous second-order partial derivatives on $(0, 1)^2$, strictly positive mixed partial C^{xy} on $(0, 1)^2$, and more (to be determined).

- Define a grid $\{(x_i, y_j) : 0 \leq i, j \leq N\}$ on $[0, 1]^2$ such that

$$0 = x_0 < x_1 < \dots < x_N = 1$$

$$0 = y_0 < y_1 < \dots < y_N = 1$$

and define filtrations

$$\mathcal{F}_x(x_i) = \sigma\{D(x_h, y_k) : 0 \leq h \leq i, 0 \leq k \leq N\}, \quad 0 \leq i \leq N$$

$$\mathcal{F}_y(y_j) = \sigma\{D(x_h, y_k) : 0 \leq h \leq N, 0 \leq k \leq j\}, \quad 0 \leq j \leq N$$

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- “Scan” the process D with respect to the filtration $\{\mathcal{F}_x\}$:

$$K_1^{(N)}(x_i, y_j) = \sum_{h=0}^{i-1} \left(D(x_{h+1}, y_j) - D(x_h, y_j) \right. \\ \left. - E[D(x_{h+1}, y_j) - D(x_h, y_j) | \mathcal{F}_x(x_h)] \right)$$

- We compute

$$\begin{aligned}
 & K_1^{(N)}(x_i, y_j) \\
 &= D(x_i, y_j) - \sum_{h=0}^{i-1} D(x_h, y_j) \left(\frac{E[D(x_h, y_j)D(x_{h+1}, y_j)]}{E[D(x_h, y_j)^2]} - 1 \right)
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 \end{aligned}$$

- Making the x -partitioning finer and finer, we obtain

$$K_1(x, y_j) = D(x, y_j) - \int_0^x D(s, y_j) \xi_1(ds, y_j)$$

as a limit in probability, where ξ_1 is an absolutely continuous measure whose density is determined by C and its first- and second-order derivatives.

- Next, we scan K_1 with respect to the filtration $\{\mathcal{F}_y\}$ and take the limit as the y -partition gets finer and finer:

$$K(x, y) = D(x, y) - \int_0^x D(s, y) \xi_1(ds, y) - \int_0^y D(x, t) \xi_2(x, dt) \\ + \int_0^x \int_0^y D(s, t) \xi_1(ds, t) \xi_2(s, dt),$$

where ξ_2 is another absolutely continuous measure whose density is determined by C and its first- and second-order derivatives.

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 - K has independent (rectangle) increments by construction, so it will suffice to show that $\text{Var}[K(x, y)] = C(x, y)$ for all $(x, y) \in [0, 1]^2$.
 - We need a technical lemma about the densities ψ_1 and ψ_2 of the measures ξ_1 and ξ_2 , namely that

$$\sup_{(x,y) \in [0,1]^2} x(1-x)|\psi_1(x, y)| < \infty,$$

$$\sup_{(x,y) \in [0,1]^2} y(1-y)|\psi_2(x, y)| < \infty.$$

- **Corollary:** The process

$$Z(x, y) = \int_0^x \int_0^y \frac{1}{\sqrt{C^{xy}(s, t)}} dK(s, t), \quad (x, y) \in [0, 1]^2$$

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- We have thus transformed the C -Brownian pillow D into a standard two-parameter Wiener process Z , through a two-step transformation:

$$D \mapsto K \mapsto Z.$$

- We apply the same two-step transformation to D_n , i.e. we define

$$\begin{aligned}
 K_n(x, y) &= D_n(x, y) - \int_0^x D_n(s, y) \xi_1(ds, y) \\
 &\quad - \int_0^y D_n(x, t) \xi_2(x, dt) \\
 &\quad + \int_0^x \int_0^y D_n(s, t) \xi_1(ds, t) \xi_2(s, dt), \\
 Z_n(x, y) &= \int_0^x \int_0^y \frac{1}{\sqrt{C^{xy}(s, t)}} dK_n(s, t)
 \end{aligned}$$

for $(x, y) \in [0, 1]^2$.

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- Future: Construct actual test statistics and procedures for goodness-of-fit tests. Consider composite null hypotheses of the form $C \in \{C_\theta : \theta \in \Theta\}$ and consider m -dimensional copulas with $m > 2$.

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- Future: Construct actual test statistics and procedures for goodness-of-fit tests. Consider composite null hypotheses of the form $C \in \{C_\theta : \theta \in \Theta\}$ and consider m -dimensional copulas with $m > 2$.
- Thank you for listening!