# Queueing models with matrix-exponential distributions and rational arrival processes 

## Analytic Methods in Queueing Systems, Eurandom 2/11 2012

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## Extension to QBD with PAP components

- ME/RAP expressions are analytically identical to PH/MAP expressions.
- Proof of the matrix geometric formula and other matrix analytic formulas rely on path wise arguments for countable state Markov chains.
A queue with ME (RAP) components can not be formulated as a Markov process with countable state space.
- We have prooved the MGM formula in the QBD case by two different approaches.
$\diamond$ An approach, where we follow a line of proof similar to the one in Ramaswami95.
$\diamond$ An approach where we apply an operator geometric result by Tweedie82 for discrete time Markov chains with general state space.

Phase Type distribution
(Jensen49), Neuts75

- A Phase type distribution is the distribution of the time to absorption in a Markov chain with $p$ transient states.
- Infinitessimal generator

$$
Q=\left(\begin{array}{cc}
S & s \\
\mathbf{0} & 0
\end{array}\right) \quad S \text { is a sub generator. }
$$

- $\tau$ : Time to absorption.
- $J(t)$ : State/phase value at $t(J(t)=p+1 ; t \geq \tau)$.
- $\mathbb{P}(J(0)=i)=\alpha_{i}, \boldsymbol{\alpha}=\left(\alpha_{1}, \ldots, \alpha_{p}\right)$, frequently $\boldsymbol{\alpha} \boldsymbol{e}=1$, where $\boldsymbol{e}$ is a vector of ones of appropriate dimension.
- $f(x)=\boldsymbol{\alpha} e^{S x} \boldsymbol{s} \quad \mathbb{P}(\tau>x)=e^{S x} \boldsymbol{e}$.

Matrix exponential distribution
$f(x)=\boldsymbol{\beta} e^{T x} \boldsymbol{t} \quad \mathbb{P}(\tau>x)=\boldsymbol{\beta} e^{T x}(-T)^{-1} \boldsymbol{t}$
$H(s)=\mathbb{E}\left(e^{-s \tau}\right)=\frac{f_{1} s^{p-1}+f_{2} s^{p-2} \ldots+f_{p}}{s^{p+g_{1}} s^{p-1}+g_{2} s^{p-2} \ldots+g_{p}}$
The span of the residual life operator is finite-dimensional
The representation $(\boldsymbol{\beta}, T, \boldsymbol{t})$ is not unique, but $T=S$ can be chosen such that $\boldsymbol{t}=-S \boldsymbol{e}=\boldsymbol{s}$ and $\left(\boldsymbol{e}_{i}, S,-S \boldsymbol{e}\right)$ is a representation for all $i$.
$f(x)=\frac{2}{3} e^{-x}(1+\cos (x))$ is matrix exponential but not phase type

Markovian Arrival Processes (MAP)

- Neuts79, Lucantoni et al. 90
- Parameterized by two matrices $\left(D_{0}, D_{1}\right), D=D_{0}+D_{1}$ is a generator, $D_{0}$ a sub-generator, and $D_{1}$ non-negative.
- For $D_{1}=\boldsymbol{d} \boldsymbol{\theta}$ we have a phase-type renewal process
- Bivariate state space $X(t)=(N(t), J(t))$ with generator

$$
Q=\left(\begin{array}{cccc}
D_{0} & D_{1} & 0 & \ldots \\
0 & D_{0} & D_{1} & \ldots \\
\vdots & \vdots & \vdots & \vdots
\end{array}\right.
$$

$E\left(z^{N(t)}\right)=\boldsymbol{\theta} e^{\left(D_{0}+D_{1} z\right) t} \boldsymbol{e} \quad\left(e^{(-\lambda+\lambda z) t}\right)$
Joint density
$\boldsymbol{\theta} e^{D_{0} x_{1}} D_{1} e^{D_{0} x_{2}} D_{1} \ldots e^{D_{0} x_{n}} D_{1} \boldsymbol{e} \quad\left(e^{-\lambda x_{1}} \lambda \ldots e^{-\lambda x_{n}} \lambda\right)$

## Rational arrival processes

Asmussen and Bladt99, (Mitchell01)

- A process is RAP if the measure of the prediction process varies in a finite dimensional space.
- There exist matrices $D_{0}, D_{1}$, a row vector $\boldsymbol{\alpha}$ and a column vector $d$, such that

$$
f\left(x_{1}, \ldots, x_{n}\right)=\boldsymbol{\alpha} e^{D_{0} x_{1}} D_{1} e^{D_{0} x_{2}} D_{1} \ldots e^{D_{0} x_{n}} \boldsymbol{d}
$$

- the parameters can be chosen such that $\boldsymbol{d}=D_{1} \boldsymbol{e}$, $\left(D_{0}+D_{1}\right) \boldsymbol{e}=\mathbf{0}$, the maximum eigenvalue of $D_{0}$ is negative, and the maximum eigenvalue of $D_{0}+D_{1}$ is 0 .

MAP/PH/1 queue

$$
Q=\left(\begin{array}{ccccc}
D_{0} & D_{1} \otimes \boldsymbol{\alpha} & 0 & 0 & \ldots \\
I \otimes \boldsymbol{s} & D_{0} \oplus S & D_{1} \otimes I & 0 & \ldots \\
0 & I \otimes \boldsymbol{s} \boldsymbol{\alpha} & D_{0} \oplus S & A_{1} \otimes I & \ldots \\
0 & 0 & I \otimes \boldsymbol{s} \boldsymbol{\alpha} & D_{0} \oplus S & \ldots \\
\vdots & \vdots & \vdots & \vdots & \vdots
\end{array}\right.
$$

The RAP/ME/1 queue can not be formulated as a Markov process with countable state space

## Why ME? (RAP)

- Minimal dimension representation
- Potential for unique representation
- ME includes all distributions with rational transform

Then why not ME? (RAP)

- The question of whether a pair represents a distribution (process) is not resolved


## $A$ Q BD with PAP components

- Define the bivariate process $N(t), \boldsymbol{A}(t)$ such that the elementary probability of an upward jump at time $t$ is $\boldsymbol{A}(t) A_{0} \boldsymbol{e} \mathrm{~d} t$ and the elementary probability of a downward jump is $\boldsymbol{A}(t) A_{2} \boldsymbol{e d} t$.
- or equivalently $e^{A_{1} t} A_{2} \boldsymbol{e}$ and $e^{A_{1} t} A_{0} \boldsymbol{e}$ are degenerate competing ME densities.

The value of $\boldsymbol{A}(t)$ after an upward jump is $\boldsymbol{A}(t-) A_{0} / \boldsymbol{A}(t) A_{0} \boldsymbol{e}$, the value of $\boldsymbol{A}(t)$ after a downward jump is $\boldsymbol{A}(t-) A_{2} / \boldsymbol{A}(t) A_{2} \boldsymbol{e}$.

- Between jumps $\boldsymbol{A}(t)$ evolves deterministically due to the equation

$$
\boldsymbol{a}^{\prime}(t)=\boldsymbol{a}(t) A_{1}(I-\boldsymbol{e} \boldsymbol{a}(t))
$$

The matrix $Q$ represents the process

$$
Q=\left(\begin{array}{ccccc}
B_{0} & A_{0} & 0 & 0 & \ldots \\
B_{1} & A_{1} & A_{0} & 0 & \ldots \\
0 & A_{2} & A_{1} & A_{0} & \ldots \\
0 & 0 & A_{2} & A_{1} & \ldots \\
\vdots & \vdots & \vdots & \vdots & \vdots
\end{array}\right.
$$

Such a process obviously exists
$\diamond$ The MAP/PH/1 queue is a trivial example
$\diamond$ The same matrix form would apply for a RAP/ME/1 queue

The problem of determining when a given $Q$ is a QBD-RAP matrix is harder.

## Censored process

Let $\{\boldsymbol{B}(t)\}_{t \geq 0}$ be the phase vector of the censored process consisting of level $m$ only, measured in the local time of level $m$, and with level $m-1$ taboo.

Theorem 1 The total lifetime $\ell_{m}(\infty)$ of $\{\boldsymbol{B}(t)\}_{t \geq 0}$ is ME distributed, that is

$$
\mathbb{P}\left(\ell_{m}(\infty)>t \mid \boldsymbol{B}(0)=\boldsymbol{a}\right)=\boldsymbol{a} e^{U t} \boldsymbol{e}
$$

for some matrix $U$.

## Expected value of phase at return to lower

## levels

The distribution of the return state

$$
\psi(\mathcal{B} ; \boldsymbol{a})=\mathbb{P}\left(\boldsymbol{A}\left(\tau_{n-1}\right) \in \mathcal{B}, \tau_{n-1}<\infty \mid X(0)=(n, \boldsymbol{a})\right), \quad \mathcal{B} \subset \mathcal{A}
$$

The expected return state

$$
\Psi(\boldsymbol{a})=\mathbb{E}\left[\boldsymbol{A}\left(\tau_{n-1}\right) I\left(\tau_{n-1}<\infty\right) \mid X(0)=(n, \boldsymbol{a})\right]=\int_{\mathcal{A}} \boldsymbol{b} \psi(\mathrm{d} \boldsymbol{b} ; \boldsymbol{a})
$$

Expected return from restricted path

$$
\Psi_{k}(\boldsymbol{a})=\mathbb{E}\left[\boldsymbol{A}\left(\tau_{n-1}\right) I\left(\tau_{n-1}<\tau_{n+k}\right) \mid X(0)=(n, \boldsymbol{a})\right] .
$$

## The matrix: $G$

Lemma 2 For $k \geq 1$, the vector valued functions $\Psi_{k}(\boldsymbol{a})$ are linear, that is, for all $\boldsymbol{a} \in \mathcal{A}, \Psi_{k}(\boldsymbol{a})=\boldsymbol{a} G_{k}$, for a unique matrix $G_{k}$. Further, $\boldsymbol{a} G_{k} \in \mathcal{A}$, for all $\boldsymbol{a} \in \mathcal{A}$.

Theorem 3 For all $a \in \mathcal{A}$, we have

$$
\Psi(\boldsymbol{a})=\mathbb{E}\left[\boldsymbol{A}\left(\tau_{n-1}\right) I\left(\tau_{n-1}<\infty\right) \mid X(0)=(n, \boldsymbol{a})\right]=\boldsymbol{a} G,
$$

for a unique matrix $G$. Further, $\boldsymbol{a} G \in \mathcal{A}$, for all $\boldsymbol{a} \in \mathcal{A}$.
With this we can prove Theorem 1 on the total lifetime $\ell_{m}(\infty)$.

The matrix geometric solution - $R$
Theorem 4 Assume that $X(\cdot)$ is an ergodic Markov process.

1. Let the vectors $\boldsymbol{\pi}_{n}, n \geq 0$, denote

$$
\begin{gathered}
\lim _{t \rightarrow \infty} \mathbb{E}[\boldsymbol{A}(t) I(L(t)=n) \mid X(0)=(j, \boldsymbol{a})] \text {, then } \\
\boldsymbol{\pi}_{n+1}=\boldsymbol{\pi}_{n} R \quad \text { for all } n \geq 1,
\end{gathered}
$$

with

$$
R=A_{0}(-U)^{-1}
$$

2. The vectors $\boldsymbol{\pi}_{0}$ and $\boldsymbol{\pi}_{1}$ satisfy
$\boldsymbol{\pi}_{1}\left(A_{1}+B_{2}\left(-B_{1}\right)^{-1} B_{0}+R A_{2}\right)=0, \quad \boldsymbol{\pi}_{0}=\boldsymbol{\pi}_{1} B_{2}\left(-B_{1}\right)^{-1}$, subject to

$$
\boldsymbol{\pi}_{1}\left(B_{2}\left(-B_{1}\right)^{-1} \boldsymbol{e}+(I-R)^{-1} \boldsymbol{e}\right)=1
$$

A discrete time Markov chain on a general state space - 「weedie82

A discrete time Markov $X_{n}=\left(N_{n}, A_{n}\right)$ chain on the state space $\mathbb{N} \times \mathcal{E}$

Consider the kernel (here in QBD-version)
$\tilde{P}(x, \mathcal{B})=\left(\begin{array}{ccccc}\tilde{B}_{0}(x, \mathcal{B}) & \tilde{A}_{0}(x, \mathcal{B}) & 0 & 0 & \ldots \\ \tilde{B}_{1}(x, \mathcal{B}) & \tilde{A}_{1}(x, \mathcal{B}) & \tilde{A}_{0}(x, \mathcal{B}) & 0 & \ldots \\ 0 & \tilde{A}_{2}(x, \mathcal{B}) & \tilde{A}_{1}(x, \mathcal{B}) & \tilde{A}_{0}(x, \mathcal{B}) & \ldots \\ 0 & 0 & \tilde{A}_{2}(x, \mathcal{B}) & \tilde{A}_{1}(x, \mathcal{B}) & \ldots \\ \vdots & \vdots & \vdots & \vdots & \vdots\end{array}\right.$
where $\tilde{A}_{0}(x, \mathcal{B})=\mathbb{P}\left(N_{n}=N_{n-1}+1, A_{n} \in \mathcal{B} \mid A_{n-1}=x\right)$
the stationary measure is given by

$$
\nu_{k}(\mathcal{B})=\int_{\mathcal{E}} \nu_{k-1}(\mathrm{~d} x) \tilde{S}(x, \mathcal{B})
$$

where $\tilde{S}$ is the minimal nonnegative solution of

$$
\begin{aligned}
\tilde{S}(x, \mathcal{B}) & =\tilde{A}_{0}(x, \mathcal{B}) \\
& +\int_{\mathcal{E}} \tilde{S}(x, \mathrm{~d} y) \tilde{A}_{1}(y, \mathcal{B})+\int_{\mathcal{E}} \int_{\mathcal{E}} \tilde{S}(x, \mathrm{~d} u) \tilde{S}(u, \mathrm{~d} y) \tilde{A}_{2}(y, \mathcal{B})
\end{aligned}
$$

That is the operator version of

$$
R=A_{0}+R A_{1}+R^{2} A_{2}
$$

## The QBD-PAP at level transitions $X_{n}=\left(N_{n}, A_{n}\right)$

- For the QBD-RAP the kernels are given by

$$
\tilde{A}_{i}(\boldsymbol{a}, \mathcal{B})=\int_{0}^{\infty} \boldsymbol{a} e^{A_{1} t} A_{i} \boldsymbol{e} I_{\mathcal{B}}\left(\frac{\boldsymbol{a} e^{A_{1} t} A_{i}}{\boldsymbol{a} e^{A_{1} t} A_{i} \boldsymbol{e}}\right) \mathrm{d} t \quad i=0,2
$$

We now introduce the expected value of the phase process at level changes into level $n$

$$
\begin{gathered}
\boldsymbol{\nu}_{i+1}=\int_{\mathcal{A}_{2}} \boldsymbol{x} \nu_{i+1}(\mathrm{~d} \boldsymbol{x})=\int_{\mathcal{A}_{2}} \boldsymbol{x} \int_{\mathcal{A}_{2}} \nu_{i}(\mathrm{~d} \boldsymbol{y}) \tilde{S}(\boldsymbol{y}, \mathrm{~d} \boldsymbol{x}) \\
=\int_{\mathcal{A}_{2}} \int_{\mathcal{A}_{2}} \nu_{i}(\mathrm{~d} \boldsymbol{y}) \boldsymbol{x} \tilde{S}(\boldsymbol{y}, \mathrm{~d} \boldsymbol{x})=\int_{\mathcal{A}_{2}} \nu_{i}(\mathrm{~d} \boldsymbol{y}) \bar{S}(\boldsymbol{y})
\end{gathered}
$$

where $\bar{S}(\boldsymbol{y})=\int_{\mathcal{A}_{2}} \boldsymbol{x} \tilde{S}(\boldsymbol{y}, \mathrm{~d} \boldsymbol{x})$.

We define the successive iterates $\tilde{S}_{k}(\boldsymbol{a}, B) \quad \tilde{S}_{k}(\boldsymbol{a}, B)=0$
$\tilde{S}_{k+1}(\boldsymbol{a}, B)=\tilde{A}_{0}(\boldsymbol{a}, B)+\int_{\mathcal{A}_{2}} \int_{\mathcal{A}_{2}} \tilde{S}_{k}(\boldsymbol{a}, \mathrm{~d} \boldsymbol{x}) \tilde{S}_{k}(\boldsymbol{x}, \mathrm{~d} \boldsymbol{y}) \tilde{A}_{2}(\boldsymbol{y}, B)$
Corresponding to $\bar{S}(\boldsymbol{a})$ we introduce

$$
\bar{S}_{k}(\boldsymbol{a})=\int_{\mathcal{A}_{2}} \boldsymbol{x} \tilde{S}_{k}(\boldsymbol{a}, \mathrm{~d} \boldsymbol{x})
$$

with $\bar{S}_{0}(\boldsymbol{a})=\mathbf{0}$
Lemma 5 The mean operator $\bar{S}_{k}(\boldsymbol{a})$ of the $k$ 'th iterate $\tilde{S}_{k}(\boldsymbol{a})$ is linear for all $k$ i.e.

$$
\bar{S}_{k}(\boldsymbol{a})=\boldsymbol{a} S_{k}
$$

where

$$
S_{k+1}=\left(-A_{1}\right)^{-1} A_{0}+S_{k}^{2}\left(-A_{1}\right)^{-1} A_{2}
$$

Operator equation for $\bar{S}$
Theorem 6 The mean operator $\bar{S}(\boldsymbol{a})$ of $\tilde{S}(\boldsymbol{a})$ is linear i.e.

$$
\bar{S}(\boldsymbol{a})=\boldsymbol{a} S
$$

$$
\begin{aligned}
\bar{S}(\boldsymbol{a})= & \int_{\mathcal{E}} \boldsymbol{x} \tilde{S}(\boldsymbol{a}, \mathrm{~d} \boldsymbol{x}) \\
= & \int_{\mathcal{E}} \boldsymbol{x} \tilde{A}_{0}(\boldsymbol{a}, \mathrm{~d} \boldsymbol{x})+\int_{\mathcal{E}} \boldsymbol{x} \int_{\mathcal{E}} \tilde{S}^{(2)}(\boldsymbol{a}, \mathrm{d} \boldsymbol{y}) \tilde{A}_{2}(\boldsymbol{y}, \mathrm{~d} \boldsymbol{x}) \\
= & \boldsymbol{a}\left(-A_{1}\right)^{-1} A_{0}+\int_{\mathcal{E}} \boldsymbol{x} \int_{\mathcal{E}} \tilde{S}^{(2)}(\boldsymbol{a}, \mathrm{d} \boldsymbol{y}) \tilde{A}_{2}(\boldsymbol{y}, \mathrm{~d} \boldsymbol{x}) \\
= & \boldsymbol{a}\left(-A_{1}\right)^{-1} A_{0} \\
& +\int_{\mathcal{E}} \boldsymbol{x} \int_{\mathcal{E}} \int_{\mathcal{E}} \tilde{S}(\boldsymbol{a}, \mathrm{~d} \boldsymbol{u}) \tilde{S}(\boldsymbol{u}, \mathrm{~d} \boldsymbol{y}) \int_{0}^{\infty} \boldsymbol{y} e^{A_{1} t} A_{2} \boldsymbol{e} I_{\mathrm{d} \boldsymbol{x}}\left(\frac{\boldsymbol{y} e^{A_{1} t} A_{2}}{\boldsymbol{y} e^{A_{1} t} A_{2} \boldsymbol{e}}\right) \mathrm{d} t
\end{aligned}
$$

$$
\begin{aligned}
\bar{S}(\boldsymbol{a}) & =\boldsymbol{a}\left(-A_{1}\right)^{-1} A_{0} \\
& +\int_{\mathcal{E}} \int_{\mathcal{E}} \tilde{S}(\boldsymbol{a}, \mathrm{~d} \boldsymbol{u}) \int_{0}^{\infty} \tilde{S}(\boldsymbol{u}, \mathrm{~d} \boldsymbol{y}) \boldsymbol{y} e^{A_{1} t} A_{2} \boldsymbol{e} \int_{\mathcal{E}} \boldsymbol{x} I_{\mathrm{d} \boldsymbol{x}}\left(\frac{\boldsymbol{y} e^{A_{1} t} A_{2}}{\boldsymbol{y} e^{A_{1} t} A_{2} e}\right) \mathrm{d} t \\
& =\boldsymbol{a}\left(-A_{1}\right)^{-1} A_{0}+\int_{\mathcal{E}} \int_{\mathcal{E}} \tilde{S}(\boldsymbol{a}, \mathrm{~d} \boldsymbol{u}) \boldsymbol{y} \tilde{S}(\boldsymbol{u}, \mathrm{~d} \boldsymbol{y})\left(-A_{1}\right)^{-1} A_{2} \\
& =\boldsymbol{a}\left(-A_{1}\right)^{-1} A_{0}+\int_{\mathcal{E}} \tilde{S}(\boldsymbol{a}, \mathrm{~d} \boldsymbol{u}) \bar{S}(\boldsymbol{u})\left(-A_{1}\right)^{-1} A_{2} \\
& =\boldsymbol{a}\left(-A_{1}\right)^{-1} A_{0}+\boldsymbol{a} S^{2}\left(-A_{1}\right)^{-1} A_{2}
\end{aligned}
$$

where we use, that $\bar{S}(\boldsymbol{u})=\boldsymbol{u} S$

$$
\boldsymbol{\nu}_{i+1}=\boldsymbol{\nu}_{i} S
$$

## The time stationary solution

$$
\boldsymbol{\pi}_{k}=c \boldsymbol{\nu}_{k}\left(-A_{1}\right)^{-1}
$$

inserting we get

$$
\boldsymbol{\pi}_{k}\left(-A_{1}\right)=\boldsymbol{\pi}_{k-1}\left(-A_{1}\right) S \quad \boldsymbol{\pi}_{k}=\boldsymbol{\pi}_{k-1} R
$$

where

$$
R=\left(-A_{1}\right) S\left(-A_{1}\right)^{-1} \quad S=\left(-A_{1}\right)^{-1} R\left(-A_{1}\right)
$$

Rather than solving for $S$ we can solve for $R$
$\left(-A_{1}\right)^{-1} R\left(-A_{1}\right)=\left(-A_{1}\right)^{-1} A_{0}+\left[\left(-A_{1}\right)^{-1} R\left(-A_{1}\right)\right]^{2}\left(-A_{1}\right)^{-1} A_{2}$
which is equivalent to

$$
A_{0}+R A_{1}+R^{2} A_{2}=0
$$

## Ekample: RAP,MAP/PH,ME/I-queue

Service time distribution:

$$
f(x)=\frac{\left.\frac{\lambda}{2}\left((\lambda x-\epsilon)^{2}+a \epsilon^{2}\right)\right)}{1-\epsilon+\frac{1+a}{2} \epsilon^{2}} e^{-\lambda x}
$$

which is an ME distribution of order 3 with $\boldsymbol{\alpha}$ and $S$ given by

$$
\boldsymbol{\alpha}=\frac{1}{1+\frac{1+a}{2} \epsilon^{2}-\epsilon}\left(1,-\epsilon, \frac{1+a}{2} \epsilon^{2}\right), \text { and } S=\left[\begin{array}{ccc}
-\lambda & \lambda & 0 \\
0 & -\lambda & \lambda \\
0 & 0 & -\lambda
\end{array}\right]
$$

which is also in PH whenever $a>0$.

## Example arrival process: Generic ME

 distributionWith

we get $\boldsymbol{g}(t)=e^{C t} \boldsymbol{e}$

$$
\boldsymbol{g}(t)=\left[\begin{array}{c}
\lambda_{1} e^{-\lambda_{1} t} \\
\frac{\lambda_{1}\left(\lambda_{2}^{2}+\omega^{2}\right)}{\lambda_{2}^{2}+\omega^{2}+b \lambda_{1} \lambda_{2}}\left(e^{-\lambda_{1} t}+b e^{-\lambda_{2} t} \sin (\omega t)\right) \\
\frac{\lambda_{1}\left(\lambda_{2}^{2}+\omega^{2}\right)}{\lambda_{2}^{2}+\omega^{2}+b \lambda_{1} \lambda_{2}}\left(e^{-\lambda_{1} t}+b e^{-\lambda_{2} t} \cos (\omega t)\right)
\end{array}\right]
$$

The distribution is not phase type for $\lambda_{1}=\lambda_{2}$ and for $|b|=1$.

## Arrival process

A $\operatorname{RAP}(C, D)$ with

$$
D=\left(\begin{array}{cccc}
\gamma_{1} & 0 & 0 & 0 \\
0 & \gamma_{1} & 0 & 0 \\
0 & 0 & \gamma_{1} & 0 \\
0 & 0 & 0 & \gamma_{2}
\end{array}\right) .
$$

(i.e. an alternating Poisson process), as a MAP

$$
C=\left(\begin{array}{cc}
T-\gamma_{1} I & \lambda(1-p) \boldsymbol{e}_{n} \\
\lambda_{3} \boldsymbol{\alpha}_{2} & -\lambda_{3}-\gamma_{2}
\end{array}\right), \quad D=\left(\begin{array}{cc}
\gamma_{1} I & 0 \\
0 & \gamma_{2}
\end{array}\right),
$$

where $T$ and $\alpha_{2}$ have dimension

$$
n=\frac{2 \pi}{\arcsin \left(\frac{2 \omega\left(\lambda_{2}-\lambda_{1}\right)}{\left(\lambda_{2}-\lambda_{1}\right)^{2}+\omega^{2}}\right)} \in \mathbb{Z}_{+},
$$

where the expression has to give an integer value.

## Experience

The dimension of the MAP/PH/1 increases linearly with $n$
The dimension of the RAP/ME/1 queue is 12
The results agreed to $10^{-10}$


A generalization

- A descriptor of measures corresponding to $\bar{S}(\boldsymbol{a})$
- Operators on measures which has linear effect on the descriptors
- Still miss characterization of matrix equation
$\diamond$ Do not invalidate results with respects to algorithms
$\diamond$ But could potentially cause numerical instability


## Conclusion

- The ME/RAP generalization can be analyzed by solving the matrix equations by Neuts
- So far we have not experienced numerical problems
- We are considering other relevant queues that can be included in the general framework (no candidates yet)

