

TCP AND ISO-STATIONARY TRANSFORMATIONS

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ABSTRACT. We consider piecewise deterministic Markov processes that occur as scaling limits of discrete-time Markov chains that describe the Transmission Control Protocol (TCP). The class of processes allows for general increase and decrease profiles. Our key observation is that stationary results for the general class follow directly from the stationary results for the idealized TCP process in [22]. The latter is a Markov process that increases linearly and experiences downward jumps at times governed by a Poisson process. To establish this connection, we apply space-time transformations that preserve the properties of the class of Markov processes.

1. INTRODUCTION

This paper deals with a class $\mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$ of piecewise deterministic Markov processes (PDMPs) related to the Transmission Control Protocol (TCP) for data transmission over the internet. The process $X(t) \in \mathcal{M}$ has the following behavior. Between random collapse times $(\tau_k)_{k \in \mathbb{N}}$ the evolution is deterministic according to

$$\frac{d}{dt}X(t) = rX(t)^\alpha, \quad r > 0.$$

The random times τ_k are governed by a state-dependent Poisson process with rate $\lambda X(t)^\beta$ and $\beta > \alpha - 1$, and at τ_k the process is multiplied by a random variable Q_k in $[0, 1)$ with distribution $F_Q(x)$, i.e., $X(\tau_k) = Q_k X(\tau_k^-)$, where the random variables $(Q_k)_{k \in \mathbb{N}}$ are independent of $\{X(t) : t \leq \tau_k\}$. If $\alpha = 0$ and $\beta = 0$ the PDMP $X(t)$ has a linear increase profile and independent losses and is known as the additive-increase multiplicative-decrease (AIMD) algorithm, also referred to as idealized TCP [22]. In general, if $X(0) = x$, the process $X(t)$ increases deterministically as

$$\phi(x, t) = \begin{cases} (x^{1-\alpha} + (1-\alpha)rt)^{\frac{1}{1-\alpha}}, & \alpha \neq 1, \\ xe^{rt}, & \alpha = 1. \end{cases} \quad (1)$$

until the time τ_1 of the first jump. It is clear that a larger α leads to a more aggressive increase profile. Also, a larger β will lead to a more aggressive decrease profile, because jumps will occur more frequently. All possible combinations of α and β together present a diverse palet of increase-decrease profiles for the dynamics of TCP.

The dynamical behavior of TCP was originally modeled as a discrete-time Markov chain [22]. Several research papers deal with establishing scaling limits for such Markov

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chains [22, 12, 20, 21, 24, 18]. A key challenge in this area was to prove that the invariant distributions of the scaled Markov chains converge to the invariant distributions of the limiting PDMPs. This issue was first settled in [24], and later extended to a larger class of models in [18]. The class \mathcal{M} at hand is the class of scaling limits that occurs in [18].

In this paper we deal solely with the stationary behavior for processes in \mathcal{M} . Recent time-dependent results can be found in [17, 23]. The key parameters in the description of \mathcal{M} are α and β for which we assume that $\beta > \alpha - 1$. This is necessary to guarantee that the process is stable and admits a stationary distribution, cf. Theorem 1. Our contribution lies in the application of space-time transformations. We show that by a state-space transformation of the type $Y(t) = \gamma X(t)^\delta$ and a time transformation $Z(u) = X(t(u))$ with $t(u) = \int_0^u X(t(s))^{-\nu} ds$, stationary results for the process $X(t) \in \mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$ follow directly from the stationary results for the idealized TCP process $Y(t) \in \mathcal{M}(1, 0; \lambda, 0; F_Q)$. In this way, all known stationary results for idealized TCP can be transferred directly to results for any process in \mathcal{M} .

For a discussion of general PDMPs we refer to [8]. The processes in this paper belong to the special class of growth-collapse processes for which we refer to [7]. Growth-collapse processes are also referred to as stress release models [25, 19, 26]. There is also a close relationship to so-called repairable systems (see [15] and [14]). In [5] and [6] a different class of stress release models with additive jumps is discussed. The papers [27] and [14] give conditions for ergodicity of a very general class of growth collapse models, including our setting (model 4 in [27] and model 1 in [14]); the latter paper also gives conditions for non-explosiveness.

Space-time changes for general Markov processes can be found in the classical literature [10, 11]. For the special class of PDMPs in this paper, such space-time changes allow us to unify several earlier results. Ott et al. [22] solve the idealized TCP case and derive solutions for $\alpha < 1, \beta = 0$ using a state-space transformation. Dumas et al. [9] consider the case $\alpha = 0, \beta > -1$, and present the stationary distribution for $\beta = 1$. Altman et al. [2] consider the case $\alpha < 1$ and give an explicit analysis of $\beta = 0, 1$ using rather general mappings that involve both space and time transformations. Baccelli et al. [3] consider a more general class of models for non-persistent flows, that includes our model for the case $\alpha = 0, \beta \leq 0$. They show that results for the throughput for $\beta > 0$ follow from the case $\beta = 0$ by applying appropriate substitutions to a differential equation. Maulik and Zwart [18] treat the case $\alpha \in [0, 1]$ and $\beta > 0$ and obtain explicit expressions for the stationary moments (cf. Section 3).

The remainder of this paper is structured as follows. In Section 2 we present our main results concerning the transformations that can be used to relate any two processes in \mathcal{M} . In Section 3 we combine a state-space transformation with a time transformation to express the stationary moments of $X(t)$ in terms of the stationary moments of idealized TCP. We also discuss the special cases of deterministic Q and (generalized) uniform Q . As for the moments, the stationary distribution for idealized TCP then yields immediately the stationary distribution for any process $X(t) \in \mathcal{M}$. Some of the more formal proofs in this paper are provided in Appendix B and C, following a short introduction to the generator of the Markov process in Appendix A.

2. MAIN RESULTS

We denote by $q_Q(s) = \mathbb{E}Q^s = \mathbb{E}e^{-s \log(1/Q)}$ the Laplace-Stieltjes transform of the non-negative random variable $\log(1/Q)$. Hence, there exists an $s_0 \leq 0$ such that $q_Q(s)$ is infinite for $s < s_0$ and finite for $s > s_0$. We also define the auxiliary function

$$\psi_Q(s) = \frac{1 - q_Q(s)}{s}, \quad s > 0,$$

and let $\psi_Q(0) = \lim_{s \rightarrow 0} \psi_Q(s) = \mathbb{E}(\log(1/Q))$.

Next we properly define the class \mathcal{M} of PDMPs.

Definition 1. Let $\mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$ denote the class of PDMPs described in Section 1. Here $r, \lambda > 0$ are real constants, F_Q is a distribution function on $[0, 1]$, and α and β are real constants with

$$\beta > \alpha - 1. \quad (2)$$

Moreover, F_Q and α are chosen in a way such that $s_0 < 0$ and

$$s_0 < 1 - \alpha. \quad (3)$$

The condition $s_0 < 0$ ensures that the distribution of Q is reasonably well behaved in the sense that all moments $\mathbb{E}(\log(1/Q)^n)$ are finite. Condition (3), which is redundant if $\alpha \leq 1$, is required to have a stable process.

We now first present some results for the stationary distribution that follow from the general theories of PDMPs and regenerative processes. Let \Rightarrow denote convergence in distribution. For a random variable A we denote throughout the paper by $\Pi_A(u)$ its distribution function, and by $\pi_A(u)$ its density.

Theorem 1. Let $X(t) \in \mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$. Then $X(t)$ stays finite for finite t and $X(t) \Rightarrow X$ as $t \rightarrow \infty$, where X is a random variable with $X > 0$ a.s. The distribution function $\Pi_X(u)$ of X admits a density $\pi_X(u)$ satisfying

$$\pi_X(u) = \frac{\lambda}{r} u^{-\alpha} \int_u^\infty y^\beta F_Q(u/y) \pi_X(y) dy. \quad (4)$$

We give a proof of this Theorem in Appendix B (see also [18]). The idea of the proof is to show that under Conditions (2) and (3), the process has negative drift outside $[0, y]$ for some y and positive drift inside $[0, w]$ for some $w \leq y$. It then follows that the regenerative process $X(t)$ has a finite cycle mean, which in turn proves the convergence result.

We next present a recursion relation for the moments of the random variable X that follows the stationary distribution of $X(t)$.

Theorem 2. Let $X(t) \in \mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$. If $q_Q(s - \alpha + 1) < \infty$ for some $s \in \mathbb{R}$, then $\mathbb{E}(X^s) < \infty$ and

$$\mathbb{E}(X^s) = \frac{\lambda}{r} \psi_Q(s - \alpha + 1) \mathbb{E}(X^{s+\beta-\alpha+1}). \quad (5)$$

Theorem 2 is proved in Appendix C. Replacing the variable s in (5) by $s + (k-1)(\beta - \alpha + 1)$, with $k \geq 1$ an integer, gives upon iteration

$$\mathbb{E}(X^{s+n(\beta-\alpha+1)}) = \mathbb{E}(X^s) \frac{r^n}{\lambda^n} \prod_{k=1}^n \frac{1}{\psi_Q(s + (k-1)\beta - k(\alpha-1))}, \quad (6)$$

which is particularly helpful for $s = 0$.

We now introduce a state-space transformation $\mathcal{M} \rightarrow \mathcal{M}$ that preserves stationarity.

Proposition 1. (state-space transformation) *Let $X(t) \in \mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$ and define a new process $Y(t) = \gamma X(t)^\delta$ with $\delta > 0$. Then*

$$Y(t) \in \mathcal{M}(r\delta\gamma^{\frac{1-\alpha}{\delta}}, 1 + \frac{\alpha-1}{\delta}; \lambda\gamma^{-\beta/\delta}, \frac{\beta}{\delta}; F_{Q^\delta}).$$

Proof. We have on the deterministic part of the trajectory

$$\frac{d}{dt}Y(t) = r\gamma\delta X(t)^{\alpha+\delta-1} = r\delta\gamma^{\frac{1-\alpha}{\delta}}Y(t)^{1+\frac{\alpha-1}{\delta}}.$$

Let $\varphi(x) = \gamma x^\delta$, so that $Y(t) = \varphi(X(t))$. The probability of a jump during $[t, t+h]$ is given by

$$\lambda\varphi^{-1}(Y(t))^\beta h + o(h),$$

where φ^{-1} denotes the inverse function of φ , and hence the jump intensity is given by $\lambda\gamma^{-\beta/\delta}Y(t)^{\beta/\delta}$. Further notice that

$$\frac{Y(\tau_k^-)}{Y(\tau_k)} = \frac{\gamma X(\tau_k^-)^\delta}{\gamma X(\tau_k)^\delta} = Q^\delta.$$

Condition (2) is invariant under the state-space transformation, since $\beta \geq \alpha - 1$ iff $\beta/\delta \geq (\alpha - 1)/\delta$. Letting $Q_Y := Q^\delta$ and $q_{Q_Y}(s) := \mathbb{E}(Q_Y^s) = q_Q(\delta s)$, it is obvious that Condition (3) is fulfilled for $Y(t)$. \square

The following result characterizes the relation between the stationary distributions of the processes $X(t)$ and $Y(t)$ that are both in \mathcal{M} .

Theorem 3. *Let $X(t) \in \mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$ and let $Y(t) = \gamma X(t)^\delta$, $\delta > 0$ as in Proposition 1. Let X be the limit of $X(t)$ as given in Theorem 1. Then $Y(t) \Rightarrow Y \stackrel{d}{=} \gamma X^\delta$, where Y has a distribution with density*

$$\pi_Y(x) = \pi_X((x/\gamma)^{1/\delta}) \cdot \frac{x^{1/\delta-1}}{\delta\gamma^{1/\delta}} \quad (7)$$

and $\mathbb{E}(Y^s) = \gamma^s \mathbb{E}(X^{s\delta})$ for all s with $q_Q(s\delta - \alpha + 1) = q_{Q_Y}(s - (\frac{\alpha-1}{\delta} + 1) + 1) < \infty$.

Proof. The first part follows from the continuous mapping theorem, and the moment relation is trivial. \square

Next we define the random process

$$\vartheta(t) = \int_0^t X(s)^\nu ds, \quad (8)$$

for $t \geq 0$, and ν some real constant, and with inverse function

$$\vartheta^{-1}(u) = \int_0^u \frac{1}{X(\vartheta^{-1}(s))^\nu} ds.$$

Throughout this section we use $u = \vartheta(t)$ and $t = \vartheta^{-1}(u)$. The time transformation of $X(t)$ is defined as follows.

Proposition 2. (time transformation) *Let ν be a real number, such that*

$$q_Q(\nu - \alpha + 1) < \infty. \quad (9)$$

Suppose that $X(t) \in \mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$. Define the time-changed process by $Z(u) = X(\vartheta^{-1}(u))$. Then

$$Z(u) \in \mathcal{M}(r, \alpha - \nu; \lambda, \beta - \nu; F_Q). \quad (10)$$

Proof. We have, in between two jumps,

$$\frac{d}{du} Z(u) = \frac{d}{du} X(\vartheta^{-1}(u)) = rZ(u)^{\alpha-\nu}.$$

Then

$$\mathbb{P}(Z(u) \text{ jumps during } [u, u+h]) = \mathbb{P}(X(\vartheta^{-1}(u)) \text{ jumps during } [u, u+h]).$$

Since $\vartheta^{-1}(u) = t$ and

$$\vartheta^{-1}(u+h) = \vartheta^{-1}(u) + \int_u^{u+h} X(\vartheta^{-1}(s))^{-\nu} ds = t + hZ(t)^{-\nu} + o(h),$$

as $h \rightarrow 0$, it follows that

$$\begin{aligned} \mathbb{P}(Z(u) \text{ jumps during } [u, u+h]) &= \mathbb{P}(X(t) \text{ jumps during } [t, \vartheta^{-1}(u+h)]) \\ &= \lambda Z(t)^{\beta-\nu} h + o(h). \end{aligned}$$

The intensity of the time-changed process is therefore given by $\lambda Z(t)^{\beta-\nu}$.

The jump sizes are not affected by a time change, hence leaving the jump distribution F_Q unchanged, so Condition 3 is still fulfilled for the time-changed process. Condition (2) remains unaltered under the time change, and (10) follows. \square

Note that $X(t)$ and $Z(t)$ are both regenerative processes, with cycles $C_X(y)$ and $C_Z(y)$ defined by successive visits to some fixed state y . We point out the following interesting property of the time-changed process, which links the cycle means to the stationary moments.

Proposition 3. *Let $X(t) \in \mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$ and let $Z(t)$ be the time-changed process as in Proposition 2, and assume that (9) holds. Then $\mathbb{E}(X^\nu) < \infty$, $\mathbb{E}(C_X(y)) < \infty$, $\mathbb{E}(C_Z(y)) < \infty$ and, for all $y > 0$,*

$$\mathbb{E}(C_Z(y)) = \mathbb{E}(C_X(y))\mathbb{E}(X^\nu).$$

Proof. It is shown in Appendix B that under Condition (2) the cycle $C_X(y)$ has a finite mean for some $y \in (0, \infty)$. We have $C_Z(y) = \vartheta(C_X(y)) = \int_0^{C_X(y)} d\vartheta(s)$ and hence

$$\mathbb{E}(C_Z(y)) = \mathbb{E}(\vartheta(C_X(y))) = \mathbb{E}\left(\int_0^{C_X(y)} X(s)^\nu ds\right) = \mathbb{E}(C_X(y))\mathbb{E}(X^\nu), \quad (11)$$

where the last equality follows from regeneration theory. Since $\mathbb{E}(X^\nu) < \infty$ by Theorem 2, we have that $\mathbb{E}(C_Z(y)) < \infty$. \square

Theorem 4. *Let $X(t) \in \mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$ and let $Z(\vartheta) = X(t)$ be the time-changed process as in Proposition 2. Let X denote the limit of $X(t)$ given in Theorem 1. Then $Z(t) \Rightarrow Z$, and the distribution of Z admits a density that satisfies*

$$\pi_Z(x) = \frac{x^\nu}{\mathbb{E}(X^\nu)} \pi_X(x). \quad (12)$$

For all s such that $q_Q(\nu + s - \alpha + 1) < \infty$ the s -th moment of Z is given by

$$\mathbb{E}(Z^s) = \frac{\mathbb{E}(X^{s+\nu})}{\mathbb{E}(X^\nu)}. \quad (13)$$

Proof. In Appendix B it is shown that $\mathbb{E}(C_X(y)) < \infty$, so that by Proposition 3 we have that $\mathbb{E}(C_Z(y)) < \infty$. It follows that $Z(t)$ has a unique stationary distribution. Moreover, again by the well known limit theorems for regenerative processes and Proposition 3,

$$\begin{aligned} \mathbb{P}(Z \leq y) &= \frac{\mathbb{E}(\int_0^{C_Z(y)} \mathbb{1}_{\{Z(s) \leq y\}} ds)}{\mathbb{E}(C_Z(y))} = \frac{\mathbb{E}(\int_0^{\vartheta(C_X(y))} \mathbb{1}_{\{Z(s) \leq y\}} ds)}{\mathbb{E}(C_X(y))\mathbb{E}(X^\nu)} \\ &= \frac{\mathbb{E}(\int_0^{C_X(y)} \mathbb{1}_{\{X(t) \leq y\}} X(t)^\nu dt)}{\mathbb{E}(C_X(y))\mathbb{E}(X^\nu)} = \frac{\mathbb{E}(\mathbb{1}_{\{X \leq y\}} X^\nu)}{\mathbb{E}(X^\nu)}, \end{aligned}$$

so that (12) follows. If $q_Q(\nu + s - \alpha + 1) < \infty$ then $\mathbb{E}(X^{s+\nu}) < \infty$ and formula (13) follows immediately from (12). \square

We conclude this section by a discussion on how to combine the introduced transformations. Denote by V the admissible region in the (α, β) -plane for which Condition (2) is fulfilled. The transformations can be represented as mappings from $V \rightarrow V$, see Figure 1, where the grey areas represent V . A state-space transformation is represented as a movement on radial lines, meeting in the point $(1, 0)$, while a time change is represented as a movement parallel to the $\beta = \alpha - 1$ line.

By using the transformations, each point (α, β) in V can be mapped to any other point (α', β') in V by first applying a state-space transformation as in Proposition 1 with

$$\delta = \frac{\beta - \alpha + 1}{\beta' - \alpha' + 1} \quad (14)$$

and then the time change in (10) with

$$\nu = \beta \frac{\beta' - \alpha' + 1}{\beta - \alpha + 1} - \beta'. \quad (15)$$

Note that the order of the two transformations is important here. If first a time change is performed, the appropriate parameter is

$$\nu = \beta - \frac{\beta - \alpha + 1}{\beta' - \alpha' + 1} \beta',$$

followed by the δ from (14) for the subsequent state-space transformation.

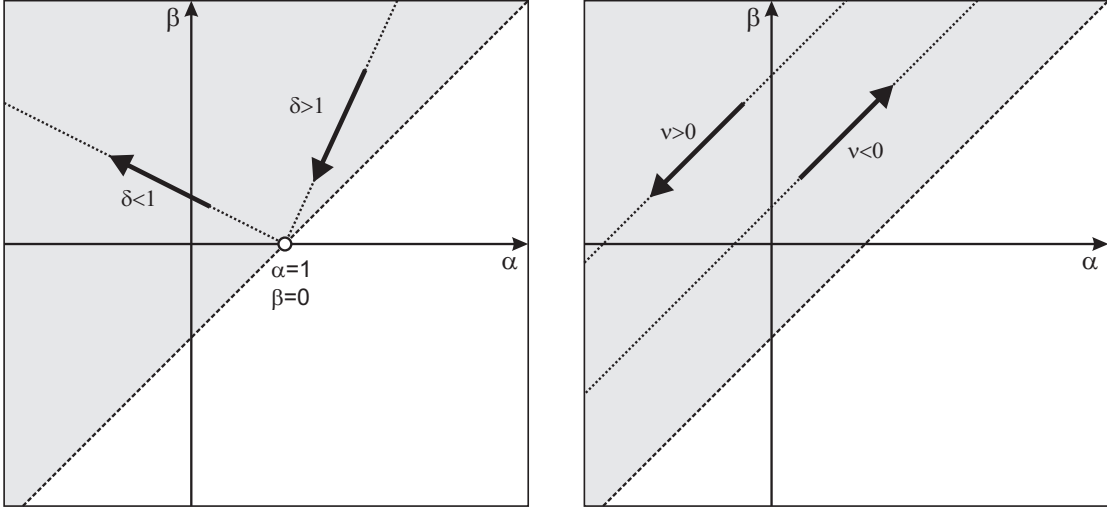


FIGURE 1. Iso-stationary transformations: state-space transformation (left) and time change (right) in the (α, β) -plane.

3. STATIONARY MOMENTS AND DISTRIBUTION

With the transformations of Section 2 in our hands, the stationary results for one process in $\mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$ yield stationary results for all other combinations of α, β in the admissible region V . In this section we shall use known results for the idealized TCP class $\mathcal{M}(1, 0; \lambda, 0; F_Q)$ to derive expressions for the stationary moments and the stationary distribution of any process in the class \mathcal{M} .

Theorem 5. *Let $X(t) \in \mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$. Then*

$$\begin{aligned} \mathbb{E}(X^s) &= \left(\frac{r(1-\alpha+\beta)}{\lambda} \right)^{\frac{s}{1-\alpha+\beta}} \frac{\Gamma(\frac{1-\alpha+s}{1-\alpha+\beta})}{\Gamma(\frac{1-\alpha}{1-\alpha+\beta})} \\ &\quad \times \prod_{k=1}^{\infty} \frac{1 - q_Q(s - \beta + k(1 - \alpha + \beta))}{1 - q_Q(-\beta + k(1 - \alpha + \beta))}, \end{aligned} \quad (16)$$

for all s for which $q_Q(s - \alpha + 1) < \infty$.

Proof. Let $Z(t) \in \mathcal{M}(1, 0; \lambda, 0; F_{Q^*})$ with $Q^* = Q^{1-\alpha+\beta}$. It is known that

$$\mathbb{E}(Z^s) = \frac{\Gamma(s+1)}{\lambda^s} \prod_{k=1}^{\infty} \frac{1 - q_{Q^*}(s+k)}{1 - q_{Q^*}(k)}, \quad (17)$$

see [12, 17, 22]. Following Proposition 1 and (14), a state-space transformation with $\delta = 1/(1 - \alpha + \beta)$ and $\gamma = (r(1 - \alpha + \beta))^{1/(1-\alpha+\beta)}$ maps $Z(t)$ to

$$Y(t) = (r(1 - \alpha + \beta)Z(t))^{1/(1-\alpha+\beta)},$$

with $Y(t) \in \mathcal{M}(r, \alpha - \beta; \lambda, 0; F_Q)$. Since $q_{Q^*}(s) = q_Q(s(1 - \alpha + \beta))$, we obtain

$$\begin{aligned} \mathbb{E}(Y^s) &= \left(r(1 - \alpha + \beta) \right)^{s/(1-\alpha+\beta)} \mathbb{E}(Z^{s/(1-\alpha+\beta)}) \\ &= \left(\frac{r(1 - \alpha + \beta)}{\lambda} \right)^{s/(1-\alpha+\beta)} \Gamma\left(\frac{s}{1-\alpha+\beta} + 1\right) \\ &\quad \times \prod_{k=1}^{\infty} \frac{1 - q_Q(s + k(1 - \alpha + \beta))}{1 - q_Q(k(1 - \alpha + \beta))}. \end{aligned} \quad (18)$$

Applying a time transformation according to Proposition 2 and (15) with $\nu = -\beta$ then yields the process $X(t)$. Formula (16) follows immediately from (13). \square

Theorem 5 is due to Maulik & Zwart [18], who derive (16) by solving a difference equation for the Mellin transform. With $l = 1 - \alpha + \beta$ and γ_* the Euler constant, they find that, for $s > \alpha - 1$, $\alpha < 1$ and $r = \lambda = 1$,

$$\begin{aligned} \mathbb{E}(X^s) &= (le^{-\gamma_*})^{\frac{s}{l}} \frac{1 - \alpha}{s + 1 - \alpha} \frac{1 - q_Q(s + 1 - \alpha)}{1 - q_Q(1 - \alpha)} \\ &\quad \times \prod_{k=1}^{\infty} e^{\frac{s}{kl}} \frac{1 - q_Q(kl + s + 1 - \alpha)}{1 - q_Q(kl + 1 - \alpha)} \frac{kl + 1 - \alpha}{kl + s + 1 - \alpha}. \end{aligned} \quad (19)$$

It can be easily verified that (16) and (19) correspond.

Let us next turn to the stationary distribution. In [12], Proposition 5, it is shown that for $Z(t) \in \mathcal{M}(1, 0; \lambda, 0; F_Q)$ the Laplace-Stieltjes transform of the stationary distribution is given by

$$\mathbb{E}(e^{-sZ}) = \mathbb{E}\left(\prod_{n=0}^{\infty} \frac{1}{1 + (s/\lambda) \prod_{k=1}^n Q_k} \right), \quad (20)$$

and alternatively (see [17])

$$\mathbb{E}(e^{-sZ}) = \sum_{n=0}^{\infty} \frac{(-s/\lambda)^n}{\prod_{k=1}^n (1 - q_Q(k))}, \quad 0 \leq s \leq \lambda. \quad (21)$$

However, both expressions do not lead (in general) to tractable explicit representations for the stationary density. Fortunately, the classical case of deterministic jumps does give an explicit form for the density.

3.1. Deterministic jumps. In the case where Q is always a constant c , we have that $q_Q(u) = c^s$ and hence $s_0 = -\infty$. Consequently all moments of X exist, and from (4) we conclude that the stationary density fulfills

$$\pi_X(u) = \frac{\lambda}{ru^\alpha} \int_u^{u/c} y^\beta \pi_X(y) dy.$$

If $Z(t) \in \mathcal{M}(1, 0; \lambda, 0; F_{Q^*})$ then

$$\pi_Z(u) = \lambda(\Pi_Z(u/c) - \Pi_Z(u)). \quad (22)$$

In [22] the solution to (22) was found to be

$$\pi_Z(u) = \frac{\lambda}{(c; c)_\infty} \sum_{k=0}^{\infty} \frac{(-1)^k c^{\frac{1}{2}k(k-1)}}{(c; c)_k} e^{-\lambda c^{-k}u}, \quad (23)$$

with $(q; q)_k = \prod_{j=1}^k (1-q^j)$ the q -Pochhammer symbol. Alternatively, (23) can be written as

$$\pi_Z(u) = \frac{\lambda}{(c; c)_\infty} \sum_{k=0}^{\infty} \frac{c^{-k}}{\prod_{j=1}^k (c^{-j} - 1)} e^{-\lambda c^{-k}u}, \quad (24)$$

in which form the solution to (22) was derived in [9, 12].

We now employ our transformations to obtain the stationary distribution for the general (α, β) -case.

Theorem 6. *If $Q = c$ a.s. then*

$$\pi_X(x) = \frac{1}{rx^\alpha} \left(\frac{r(1-\alpha+\beta)}{\lambda} \right)^{\frac{\beta}{1-\alpha+\beta}} \frac{\pi_Z\left(\frac{x^{1-\alpha+\beta}}{r(1-\alpha+\beta)}\right)}{\Gamma\left(\frac{1-\alpha}{1-\alpha+\beta}\right)} \prod_{k=1}^{\infty} \frac{1 - c^{k(1-\alpha+\beta)}}{1 - c^{k(1-\alpha+\beta)-\beta}}. \quad (25)$$

Proof. Let $Z(t) \in \mathcal{M}(1, 0; \lambda, 0; F_{Q^*})$ with $Q^* = c^{1-\alpha+\beta}$. As in the proof of Theorem 5 we apply a state-space transformation with $\delta = 1/(1-\alpha+\beta)$ and $\gamma = (r(1-\alpha+\beta))^{1/(1-\alpha+\beta)}$ to $Z(t)$, yielding a process $Y(t) \in \mathcal{M}(r, \alpha - \beta; \lambda, 0; F_Q)$, where F_Q in this case is the distribution function having mass 1 at c . According to (7)

$$\pi_Y(x) = \pi_Z\left(\frac{x^{1-\alpha+\beta}}{r(1-\alpha+\beta)}\right) \frac{x^{\beta-\alpha}}{r}.$$

Next we perform a time change with $\nu = -\beta$, yielding a process $X(t)$ in the class $\mathcal{M}(r, \alpha; \lambda, \beta; F_Q)$. From (12) it follows that

$$\pi_X(x) = \frac{x^{-\beta}}{\mathbb{E}(Y^{-\beta})} \pi_Y(x).$$

According to (18) we have

$$\mathbb{E}(Y^{-\beta}) = \left(\frac{\lambda}{r(1-\alpha+\beta)} \right)^{\frac{\beta}{1-\alpha+\beta}} \Gamma\left(\frac{1-\alpha}{1-\alpha+\beta}\right) \prod_{k=1}^{\infty} \frac{1 - c^{k(1-\alpha+\beta)-\beta}}{1 - c^{k(1-\alpha+\beta)}}, \quad (26)$$

and (25) follows. \square

3.2. Another special case. If $Q = U^{1/\kappa}$, with some $\kappa > 0$ and U having a uniform distribution on $[0, 1]$, then

$$q_Q(s) = \mathbb{E}(Q^s) = \int_0^1 x^{s/\kappa} dx = \frac{\kappa}{\kappa + s},$$

so that $s_0 = -\kappa$ and $q_Q(s_0) = \infty$. It follows that $\mathbb{E}(X^s) < \infty$ if $s > \alpha - 1 - \kappa$.

Theorem 7. *If $Q = U^{1/\kappa}$ with $\kappa > 0$ then*

$$\pi_X(x) = \frac{\Gamma\left(\frac{\kappa-\alpha+1}{\beta-\alpha+1}\right)}{\beta-\alpha+1} \left(\frac{\lambda}{r(\beta-\alpha+1)} \right)^{-\frac{\kappa-\alpha+1}{\beta-\alpha+1}} x^{\kappa-\alpha} e^{-\frac{\lambda}{r(\beta-\alpha+1)}x^{\beta-\alpha+1}}. \quad (27)$$

Proof. We derive the proposition directly for all admissible α and β , without using the transformations. From (4) we obtain

$$\pi_X(u) = \frac{\lambda u^{\kappa-\alpha}}{r} \int_u^\infty y^{\beta-\kappa} \pi_X(y) dy \quad (28)$$

and hence

$$\pi'_X(u) = \left(\frac{\kappa - \alpha}{u} - \frac{\lambda u^{\beta-\alpha}}{r} \right) \pi_X(u). \quad (29)$$

For $y < x$,

$$\pi_X(x) = \frac{\pi_X(y)}{y^{\kappa-\alpha}} x^{\kappa-\alpha} \exp\left(-\frac{\lambda}{r(\beta-\alpha+1)}(x^{\beta-\alpha+1} - y^{\beta-\alpha+1})\right).$$

Letting $y \rightarrow 0$, and assuming that the limit $C = \lim_{y \rightarrow 0} \pi_X(y) y^{\alpha-\kappa}$ exists, we obtain

$$\pi_X(x) = C x^{\kappa-\alpha} \exp\left(-\frac{\lambda}{r(\beta-\alpha+1)} x^{\beta-\alpha+1}\right).$$

The constant follows upon normalization. \square

We note that (27) was derived for the case $\kappa = \alpha = \beta = 0$ in [12], p. 99.

APPENDIX A. PDMPs, GENERATOR, DOMAIN

Let $\mathbb{E}_x(\cdot) = \mathbb{E}(\cdot | X(0) = x)$. For bounded measurable functions f the infinitesimal generator refers to the linear operator defined by the limit

$$\mathcal{A}^* f(x) = \lim_{t \rightarrow 0} (\mathbb{E}_x f(X(t)) - f(x)) / t,$$

in the strong sense, that is with respect to the sup norm. It is well known that for functions for which this limit exists, the process

$$f(X(t)) - \int_0^t \mathcal{A}^* f(X(s)) ds$$

is a martingale. The extended generator $\mathcal{A}f$ is a generalization of this operator and is defined as any measurable function g , for which $f(X(t)) - \int_0^t g(X(s)) ds$ becomes a martingale. Following the exposition in [8] we see that our Markov process $X(t)$ is a PDMP and has the extended generator

$$\mathcal{A}f(x) = rx^\alpha f'(x) + \lambda x^\beta \int_0^1 (f(xy) - f(x)) dF_Q(y). \quad (30)$$

For PDMPs Davis [8] has given criteria for a measurable function f to belong to the domain $\mathcal{D}(\mathcal{A})$ of the extended generator. The domain $\mathcal{D}(\mathcal{A})$ contains all measurable functions $f : E \rightarrow \mathbb{R}$ for which f is absolutely continuous and

$$\mathbb{E}_x \sum_{i=1}^{N_t} |f(X(\tau_i-)) - f(X(\tau_i))| < \infty, \quad \forall x \in [0, \infty), t \geq 0.$$

If f is absolutely continuous and locally bounded on $[0, \infty)$ then $f \in \mathcal{D}(\mathcal{A})$ follows immediately from the fact that according to (1) $X(t)$ has a deterministic upper bound. Since the domain of \mathcal{A} is probably much larger, it is desirable to know whether $\mathcal{D}(\mathcal{A})$

comprises function that are not locally bounded. From Lemma 1 in [17] we obtain the following result:

Lemma 1 (domain). *The function $f_s(x) = x^s$ is in the domain $\mathcal{D}(\mathcal{A})$ if $q_Q(s) < \infty$.*

We define the counting process $N_t = \inf\{n \in \mathbb{N} : \tau_n \geq t\}$. In order to use the methodology of PDMPs we have to assure that $\mathbb{E}(N_t) < \infty$.

Lemma 2. $\mathbb{E}_x N_t < \infty$ for all $t > 0$, $x \geq 0$.

Proof. If $\beta = 0$ then the jump intensity is just λ and clearly $\mathbb{E}N_t = \lambda t$ is finite for all $t > 0$. The general case with $\beta > 0$ can be transformed via a time change as described in Proposition 2 and 3 to the $\beta = 0$ case (see Figure 1). \square

Next we show that the process will not escape to infinity in finite time.

Lemma 3. $T_\infty := \sup\{t > 0 : X(t) < \infty\} = \infty$ a.s.

Proof. If $\alpha < 1$ then this follows from the fact that $\phi(x, t)$ is finite for all $t > 0$ and $X(t) \leq \phi(x, t)$. If $\alpha \geq 1$ a time change with $\nu = \beta$ transforms the process into a process $Z(t)$ with increase rate parameter $\alpha - \beta < 1$ and constant jump intensity (see Proposition 2). $T_\infty = \infty$ iff

$$\vartheta^{-1}(u) = \int_0^u \frac{1}{Z(s)^\beta} ds < \infty$$

for all $u > 0$. Since $N_t < \infty$ for all t (see Lemma (2)) we have $\min\{Z(s) : 0 \leq s \leq t\} > 0$ and hence $\vartheta^{-1}(u) \leq u(\min\{Z(s) : 0 \leq s \leq t\})^{-\beta} < \infty$. \square

APPENDIX B. PROOF OF THEOREM 1

Lemma 3 proves that $X(t)$ stays finite for finite t . We now prove Theorem 1 for the case $\beta = 0$. The general form of Theorem 1 then follows by applying a time change as described in Proposition 2 and 3.

Theorem 8. *Let $X(t) \in \mathcal{M}(r, \alpha; \lambda, 0; F_Q)$. Then $X(t) \Rightarrow X$ as $t \rightarrow \infty$, where X is a random variable with $X > 0$ a.s. The distribution function $\Pi_X(u)$ of X admits a density $\pi_X(u)$ satisfying*

$$\pi_X(u) = \frac{\lambda}{r} u^{-\alpha} \int_u^\infty F_Q(u/y) \pi_X(y) dy. \quad (31)$$

We need two lemmas and the following definitions for the proof of Theorem 8. Let $T_y^* = \inf\{t > 0 : X(t) \leq y\}$ and $T_y = \inf\{t > 0 : X(t) = y\}$ denote the first time the process jumps below y and the first hitting time of y . Note that we have $\alpha < 1$ throughout the proof.

Lemma 4. *There is a $y \in (0, \infty)$ such that $\mathbb{E}_x T_y^* < \infty$ for all $x \geq y$.*

Proof. Recall that $f_s(x) = x^s$ and that $f_s \in \mathcal{D}(\mathcal{A})$ if $q_Q(s) < \infty$. We first show that there is a $s > 0$, such that $\limsup_{x \rightarrow \infty} \mathcal{A}f_s(x) < 0$ as $x \rightarrow \infty$. We have

$$\mathcal{A}f_s(x) = sr x^{\alpha+s-1} - \lambda x^s (1 - q_Q(s)) = x^s (sr x^{\alpha-1} - \lambda(1 - q_Q(s))).$$

Let $s > 0$, then $x^s \rightarrow \infty$ and $x^{\alpha-1} \rightarrow 0$ as $x \rightarrow \infty$, hence $\mathcal{A}f_s(x) \rightarrow -\infty$ and thus $\limsup_{x \rightarrow \infty} \mathcal{A}f_s(x) < 0$ (note that $(1 - q_Q(s)) > 0$ since $s > 0$).

We have shown that there is a $y \in (0, \infty)$, such that $\mathcal{A}f_s(x) < -\varepsilon$ for all $x \geq y$ and some $s > 0$. Pick one such x and recall from the definition of the extended generator that the process $M_t = f_s(X(t)) - f_s(x) - \int_0^t \mathcal{A}f_s(X(s)) ds$ is a zero-mean martingale. Optional sampling yields $\mathbb{E}_x M_{t \wedge T_y^*} = 0$, and hence

$$0 = \mathbb{E}_x \left(f_s(X(t \wedge T_y^*)) - f_s(x) - \int_0^{t \wedge T_y^*} \mathcal{A}f_s(X(s)) ds \right) \quad (32)$$

It follows that $\mathbb{E}_x (f_s(X(t \wedge T_y^*)) - f_s(x) + (t \wedge T_y^*)\varepsilon) \leq 0$. Consequently

$$\mathbb{E}_x(t \wedge T_y^*) \leq \frac{1}{\varepsilon} (f_s(x) - \mathbb{E}_x f_s(X(t \wedge T_y^*))) \leq \frac{f_s(x)}{\varepsilon}.$$

Letting $t \rightarrow \infty$ it follows that $\mathbb{E}_x T_y^* < \infty$. \square

Next we show that the mean of the hitting time $\mathbb{E}_x T_w$ is also finite for some w .

Lemma 5. *There is a $w \in (0, \infty)$ such that $\mathbb{E}_x T_w < \infty$ for all $x \leq y$.*

Proof. It is enough to show that Lemma 4 holds for the reciprocal process $R(t) = X^{-1}(t)$, which is the case if we can show that $\limsup_{x \rightarrow \infty} \mathcal{A}_R f_s(x) < 0$, where

$$\mathcal{A}_R f(x) = -rx^{2-\alpha} f'(x) + \lambda \int_0^1 (f(x/y) - f(x)) dF_Q(y).$$

is the generator of the Markov process $R(t)$. By following the proof of Lemma 1 it is clear that f_s is in the domain of \mathcal{A}_R if $q_Q(-s) < \infty$. We have

$$\mathcal{A}_R f_s(x) = -sr x^{1-\alpha+s} + \lambda x^s (q_Q(-s) - 1) = x^{1-\alpha+s} (-sr + \lambda x^{\alpha-1} (q_Q(-s) - 1)).$$

Let $0 < s < -s_0$. Then f_s is in the domain of \mathcal{A}_R . Since $x^{1-\alpha+s} \rightarrow \infty$ and $x^{\alpha-1} \rightarrow 0$ as $x \rightarrow \infty$, it follows that $\limsup_{x \rightarrow \infty} \mathcal{A}_R f_s(x) < 0$. The result then follows along the lines of the second part of the proof of Lemma 4. \square

Proof of Theorem 8. The process $X(t)$ is a regenerative process with regeneration cycles $C_X(y)$ starting at upcrossings of y .

Under the stated conditions the assertions of Lemma 4 and 5 are both fulfilled, i.e. there are $y, w \in (0, \infty)$ such that $\mathbb{E}_x T_y^* < \infty$ for all $x \geq y$ and $\mathbb{E}_x T_w < \infty$ for all $x \leq y$. Without loss of generality we can assume that $w < y$.

We will show that $\mathbb{E}_w T_y < \infty$. Then $\mathbb{E}_x T_y$ for all $x \leq y$ and we have already shown that under the given conditions $\mathbb{E}_y T_y^* < \infty$. It follows then that for the regenerative process $X(t)$, with regeneration cycles starting at upcrossings of y , the cycle length has a finite mean and the assertion of the lemma follows.

To prove $\mathbb{E}_w T_y < \infty$ consider a process $W(t)$ which behaves similar to $X(t)$ if $X(t) \in [0, \infty) \setminus [w, y]$, but has the following behavior if $W(t) \in [w, y]$: it jumps with intensity λ and if there is a jump at time t then $W(t) = w \wedge (W(t-) \cdot Q)$. We can construct $X(t)$ and $W(t)$ on a common probability space, such that $X(t) \geq W(t)$ w.p. one. Consequently, $W(t)$ will reach y later than $X(t)$, so it remains to show that, starting in $W(0) = w$, the expected time to reach y is finite.

Starting in w the process $W(t)$ will reach y without any jump with some positive probability, or it jumps down into the set $[0, w]$. We proved in Lemma 4 that the time until the process returns to w has finite mean. It follows that the time until $W(t)$ reaches y is a geometric sum of random variables with finite mean, and hence has finite mean.

We have seen that the regeneration cycles $C_X(y)$ of $X(t)$ have finite mean, hence $X(t) \Rightarrow X$ as $t \rightarrow \infty$ follows from regeneration theory (see e.g. [1]). But we have actually shown more. Under the given conditions, Lemma 5 shows that also the reciprocal process $R(t) = X^{-1}(t)$ tends to a finite limit, hence X is not zero.

To prove (4) we use the fact that $\int_0^\infty \mathcal{A}f(x)\pi_X(x) dx = 0$ for bounded functions $f \in \mathcal{D}(\mathcal{A})$. Let $g_n(x) = n \int_0^x \mathbb{1}_{\{w \in [u-1/n, u+1/n]\}} dw$, then g_n is bounded and absolutely continuous. We have

$$\int_0^\infty r x^\alpha g'_n(x) \pi_X(x) dx = n \int_{u-1/n}^{u+1/n} r x^\alpha \pi_X(x) dx$$

which tends to $ru^\alpha \pi_X(u)$ as $n \rightarrow \infty$. Since $g_n(x) \rightarrow \mathbb{1}_{\{x \geq u\}}$, we obtain

$$\begin{aligned} & \int_0^\infty \left(\lambda \int_0^1 (g_n(xy) - g_n(x)) dF_Q(y) \right) \pi_X(x) dx \\ \rightarrow & \int_0^\infty \lambda \left(\int_0^1 \mathbb{1}_{\{xy \geq u\}} dF_Q(y) - \mathbb{1}_{\{x \geq u\}} \right) \pi_X(x) dx \\ = & \int_u^\infty \lambda \left(\int_{u/x}^1 dF_Q(y) - 1 \right) \pi_X(x) dx = - \int_u^\infty \lambda F_Q(u/x) \pi_X(x) dx. \end{aligned}$$

Hence (31) follows. \square

Having proved Theorem 8 the final step to establish Theorem 1 is a application of the time change, which transforms the $\beta = 0$ case to the $\beta \neq 0$ case (Proposition 2 with $\nu = \beta$). Note that according to Proposition 3, we have to fulfill (9), which yields the extra condition (3) in Definition 1.

APPENDIX C. PROOF OF THEOREM 2

We shall now prove that $\mathbb{E}_x(X^s) < \infty$ if $q_Q(s - \alpha + 1) < \infty$ and that (5) holds. Let $f_{s,w}(x) = \mathbb{1}_{\{x \leq w\}} x^{s-\alpha+1}$. Since $q_Q(s - \alpha + 1) < \infty$ it follows from Lemma 1 that $f_{s,w} \in \mathcal{D}(\mathcal{A})$. Then

$$\begin{aligned} \mathcal{A}f_{s,w}(x) &= \left(r(s - \alpha + 1)x^s + \lambda(q_Q(s - \alpha + 1) - 1)x^{s+\beta-(\alpha-1)} \right) \mathbb{1}_{\{x \leq w\}} \\ &+ \left(\lambda x^{s+\beta-(\alpha-1)} \int_0^{w/x} y^{s-\alpha+1} dF_Q(y) \right) \mathbb{1}_{\{x > w\}}. \end{aligned}$$

Since $f_{s,w}$ is bounded it follows from $\int_0^\infty \mathcal{A}f(x)\pi_X(x) dx = 0$ that

$$\int_0^w x^s \pi_X(x) dx = \frac{\lambda}{r} \psi_Q(s - \alpha + 1) \int_0^w x^{s+\beta-(\alpha-1)} \pi_X(x) dx - C(w), \quad (33)$$

where $C(w) = \frac{\lambda}{r(s-\alpha+1)} \int_w^\infty x^{s+\beta-(\alpha-1)} \int_0^{w/x} y^{s-\alpha+1} dF_Q(y) \pi_X(x) dx$. We have

$$\int_0^{w/x} y^{s-\alpha+1} dF_Q(y) \leq \int_0^1 y^{s-\alpha+1} dF_Q(y) = q_Q(s - \alpha + 1),$$

and hence

$$0 \leq C(w) \leq \frac{\lambda q_Q(s - \alpha + 1)}{r(s - \alpha + 1)} \int_w^\infty x^{s+\beta-(\alpha-1)} \pi_X(x) dx.$$

We see that $C(w) \rightarrow 0$ when $w \rightarrow \infty$. Consequently, relation (5) follows from (33) by dominated convergence.

To prove that $\mathbb{E}_x(X^s) < \infty$ we first assume that $s = 0$. Then $q_Q(s - \alpha + 1) = q_Q(1 - \alpha) < \infty$ and since $\mathbb{E}_x(X^s) = 1 < \infty$ it follows that $\mathbb{E}_x(X^{\beta-(\alpha-1)}) < \infty$. We thus have $\mathbb{E}_x(X^t) < \infty$ for all $t \in [0, \beta - (\alpha - 1)]$ and by induction $\mathbb{E}_x(X^t) < \infty$ for all $t > 0$.

On the other hand, if $s < 0$ and $q_Q(s - \alpha + 1) < \infty$, then there is a $k \in \mathbb{N}$, such that $s_k = s + k(\beta - (\alpha - 1)) \geq 0$ and hence $\mathbb{E}_x(X^{s_k}) < \infty$. Moreover $s_k - \beta \geq s - \alpha + 1$ and also $q_Q(s_k - \beta) < \infty$. It follows then from $\mathbb{E}_x(X^{s_k}) < \infty$ that $\mathbb{E}_x(X^{s_{k-1}}) < \infty$ and by induction $\mathbb{E}_x(X^s) < \infty$.

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REFERENCES

- [1] Asmussen, S. (2003). *Applied Probability and Queues*. 2nd edition. New York, Springer.
- [2] Altman, A., Avrachenkov, K., Kherani, A.A. and Prabhu, B.J. (2005). Performance Analysis and Stochastic Stability of Congestion Control Protocols. In *Proceedings of IEEE Infocom 2005*.
- [3] Baccelli, F., Kim, K. B. and McDonald, D. (2007). Equilibria of a class of transport equations arising in congestion control. *Queueing Systems* **55** pp. 1-8.
- [4] Borovkov, K. and Last, G. (2008). On level crossings for a general class of piecewise-deterministic Markov processes. *Adv. Appl. Probab.* **40** pp. 815–834.
- [5] Borovkov, K. and Novikov, A. (2001). On a piece-wise deterministic Markov process model. *Stat. Probab. Lett.* **53** pp. 421–428.
- [6] Borovkov, K. and Vere-Jones, D. (2000). Explicit formulae for stationary distributions of stress release processes. *J. Appl. Probab.* **37** pp. 315–321.
- [7] Boxma, O., Perry, D., Stadje, W. and Zacks, S. (2006). A Markovian growth-collapse model. *Adv. Appl. Probab.* **38** pp. 221–243.
- [8] Davis, M.H.A. (1993). *Markov Models and Optimization*. London, Chapman & Hall.
- [9] Dumas, V., Guillemin, F. and Robert, Ph. (2002). A Markovian analysis of Additive-Increase Multiplicative-Decrease (AIMD) algorithms. *Adv. Appl. Probab.* **34** pp. 85–111.
- [10] Dynkin, E.B. (1965). *Markov Processes I*. New York, Springer.
- [11] Ethier, S.N. and Kurtz, T.G. (1986). *Markov Processes. Characterization and Convergence*. John Wiley & Sons.
- [12] Guillemin, F., Robert, Ph. and Zwart, B. (2004). AIMD algorithms and exponential functionals. *Ann. Appl. Probab.* **14** pp. 90–117.
- [13] Kelly, T. (2003). Scalable TCP: Improving performance in highspeed wide area networks. *Computer Communications Review* **32** pp. 83–91.
- [14] Last, G. (2004). Ergodicity properties of stress release, repairable system and workload models. *Adv. Appl. Probab.* **36** pp. 471–498.

- [15] Last, G. and Szekli, R. (1999). Time and Palm stationarity of repairable systems. *Stochastic Processes Appl.* **79** pp. 17–43.
- [16] van Leeuwaarden, J.S.H. and Löpker, A.H. (2008). Hitting times for multiplicative growth-collapse processes. Submitted.
- [17] Löpker, A.H. and van Leeuwaarden, J.S.H. (2008). Transient moments of the TCP window size process. *J. Appl. Probab.* **45** pp. 163–175.
- [18] Maulik, K. and Zwart, B. (2008). An extension of the square root law of TCP. To appear in *Annals of Operations Research*.
- [19] Ogata, J. and Vere-Jones, D. (1984). Inference for earthquake models: a self-correcting model. *Stochastic Processes Appl.* **17** pp. 337–347.
- [20] Ott, T.J. (2005). Transport protocols in the TCP paradigm and their performance. *Telecommunication Systems* **30** pp. 351–385.
- [21] Ott, T.J. (2006). Rate of convergence for the square root formula in the Internet transmission control protocol. *Adv. Appl. Probab.* **38** pp. 1132–1154.
- [22] Ott, T. and Kemperman, J. and Mathis, M. (1996) The stationary behavior of ideal TCP congestion avoidance. Unpublished manuscript.
- [23] Ott, T.J. and Kemperman, J.H.B. (2008). The transient behavior of processes in the TCP paradigm *Probability in the Engineering and Informational Sciences* **22** pp. 431–471.
- [24] Ott, T.J. and Swanson, J. (2007). Asymptotic behavior of a generalized TCP congestion avoidance algorithm. *J. Appl. Probab.* **44** pp. 618–635.
- [25] Vere-Jones, D. (1978). Earthquake prediction: a statistician’s view *J. Phys. Earth* **26** pp. 129–146.
- [26] Vere-Jones, D. (1988). On the variance properties of stress release models. *Austral. J. Stat.* **30** pp. 123–135.
- [27] Zheng, X. (1991). Ergodic theorems for stress release processes. *Stochastic Processes Appl.* **37** pp. 239–258.

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